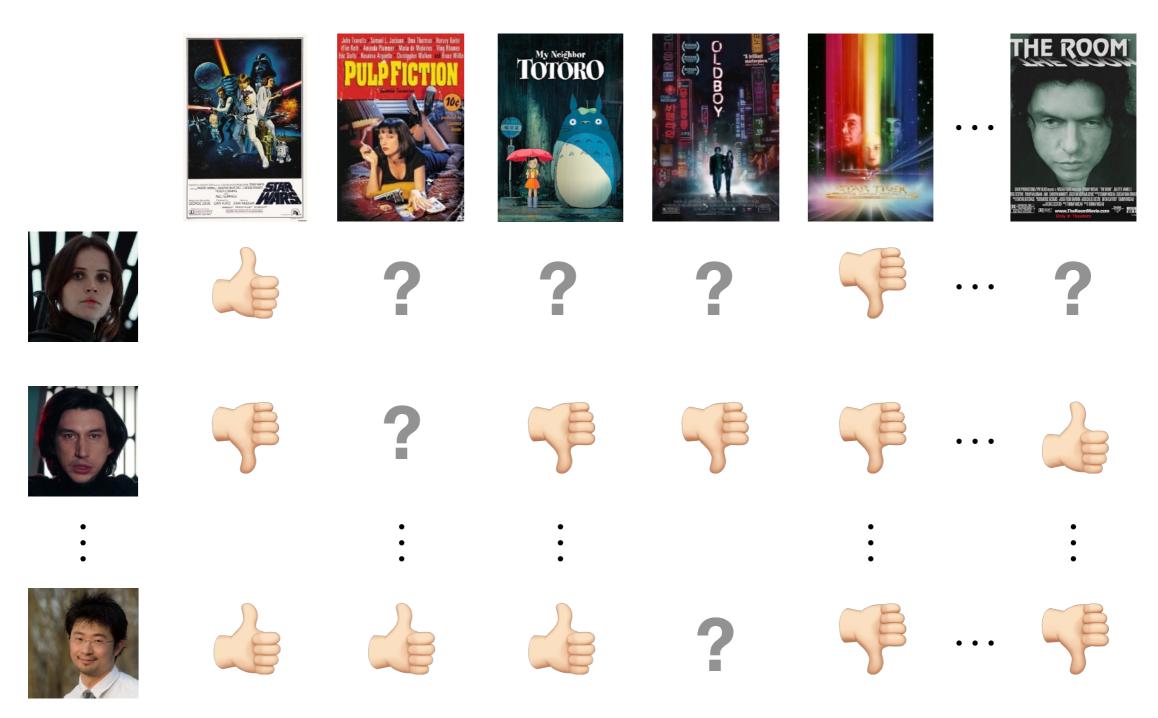
Missing Not at Random in Matrix Completion

The Effectiveness of Estimating Missingness Probabilities Under a Low Nuclear Norm Assumption

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Matrix Completion



Goal: Fill in question marks (subject to constraints)

Largely popularized by the Netflix Prize (Bennett & Lanning 2007)

Application: Prediction with Missing Values

	Feature vectors				Labe	Labels to predict	
	Gluten allergy	lmmuno- suppressant	Low resting heart rate	Irregular heart beat	High BMI	Time of death	
	X	?			?	?	
		?		?		?	
1	X		?		*	?	

Common approach:

- I. Impute missing features with matrix completion
- 2. Use imputed feature vectors to solve prediction task

Missing Not at Random (MNAR) in MC

MNAR: missingness is not uniform at random and can depend on value of entry (if it were forced to be revealed)

- Restaurant ratings: a vegan is unlikely to go to & rate a BBQ restaurant
- Movie ratings: some people refuse to watch horror movies
- Health care: doctor chooses measurements to take for a patient

The vast majority of existing literature on matrix completion assumes entries are missing with equal probability independent of everything else (Candès & Recht 2009, Recht 2009, Cai et al 2010, Keshavan et al 2010, ...)

Many methods rely on this missing-completely-at-random (MCAR)
assumption and produce biased predictions when the data are MNAR

This paper: new approach to MNAR matrix completion with (1) finite sample debiasing guarantees & (2) competitive empirical accuracy

Example of Bias in MC (Steck 2010)

True ratings matrix $S \in \mathbb{R}^{m \times n}$

Goal: Given X, construct estimate \widehat{S} of S

Revealed ratings matrix \boldsymbol{X}

 Ω : set of revealed indices

$$\begin{bmatrix} +1 & +1 & ? & ? & ? \\ ? & +1 & +1 & ? & ? \\ +1 & ? & +1 & ? & ? \\ \hline ? & ? & ? & +1 & +1 \\ ? & ? & ? & +1 & ? \\ \end{bmatrix}$$

Predict all I's (set \widehat{S} to all I's)

Ideally, minimize:
$$L_{\text{ideal-MSE}}(\widehat{S}) = \frac{1}{mn} \sum_{i=1}^m \sum_{j=1}^n (S_{i,j} - \widehat{S}_{i,j})^2 = 1.92$$

In practice, minimize:
$$L_{\text{naive-MSE}}(\widehat{S}) = \frac{1}{|\Omega|} \sum_{(i,j) \in \Omega} (X_{i,j} - \widehat{S}_{i,j})^2 = 0$$

If every entry revealed with equal probability:

$$L_{\text{naive-MSE}}(\widehat{S})$$
 is unbiased estimate of $L_{\text{ideal-MSE}}(\widehat{S})$

Model

True ratings matrix $S \in \mathbb{R}^{m \times n}$

		Horror		Romance		
		movies		movies		
Horror	$\lceil +1 \rceil$	$+1 \\ +1$	+1	-1	-1	
lovers	+1	+1	+1	-1	-1	
107613	+1	+1	+1	-1	-1	
Romance	$\overline{-1}$	-1			$\overline{+1}$	
lovers	-1	-1	-1	+1	+1	

Goal: Given X, construct estimate \widehat{S} of S

Revealed ratings matrix \boldsymbol{X}

 Ω : set of revealed indices

$$\begin{bmatrix} -1 & -1 & ? & ? & ? \\ ? & +1 & -1 & ? & ? \\ +1 & ? & +1 & ? & ? \\ \hline ? & ? & ? & -1 & +1 \\ ? & ? & ? & +1 & ? \end{bmatrix}$$

Probabilities of entries being revealed $P \in [0, 1]^{m \times n}$

0.5		0.5	0.0	0.0
0.5	0.5	0.5	0.0	0.0
0.5			0.0	0.0
0.0	0.0	0.0	0.5	0.5
0.0	0.0	0.0	0.5	0.5

Generative process:

- I. Reveal entries of S based on P
- 2. Add noise to revealed entries

Debiasing MC with Propensity Scores

Goal: Given X, construct estimate \widehat{S} of S

Probabilities of entries being revealed $P \in [0,1]^{m \times n}$

0.5	0.5	0.5	0.0	0.0
0.5	0.5	0.5	0.0	0.0
0.5	0.5	0.5	0.0	0.0
0.0	0.0	0.0	0.5	0.5
0.0	0.0	0.0	0.5	0.5

ightharpoonup Will need probabilities > 0

Think of revealing an entry as a "treatment" (Schnabel et al 20 6)

Use inverse propensity score weighting (Horvitz & Thompson 1952, ...)

Matrix of propensity scores

- I. Construct estimate \widehat{P} of P
- 2. Minimize:

$$L(\widehat{S}|\widehat{P}) = \frac{1}{mn} \sum_{(i,j)\in\Omega} \frac{(X_{i,j} - \widehat{S}_{i,j})^2}{\widehat{P}_{i,j}}$$

Unbiased estimate of $L_{\text{ideal-MSE}}(\widehat{S})$ if $\widehat{P} = P$

(Other weighting schemes are also possible)

Debiasing MC

I. Construct estimate \widehat{P} of P

Typically done via parametric model (logistic regression, naive Bayes) (for MC: Liang et al 2016, Schnabel et al 2016, Wang et al 2018/2019, ...)

- Usually requires auxiliary information (on rows/cols, some MCAR data)
- Unclear what error is for estimating propensity scores
- 2. Solve modified version of standard MC problem:

$$\widehat{S} = \underset{Z \in [-1,1]^{m \times n}}{\operatorname{argmin}} \{ L(Z|\widehat{P}) + \lambda \|Z\|_* \} \qquad \text{Convex program nuclear norm}$$

$$\qquad \qquad \text{(encourages low rank)}$$

where

$$L(Z|\widehat{P}) = \frac{1}{mn} \sum_{(i,j)\in\Omega} \frac{(X_{i,j} - Z_{i,j})^2}{\widehat{P}_{i,j}}$$

Standard approach uses $L_{\mathrm{naive-MSE}}(Z)$ instead of $L(Z|\widehat{P})$ (Mazumder et al 2010)

Debiasing MC

I. Construct estimate \widehat{P} of P

Main contribution: New strategy to estimating \widehat{P} with

- Finite sample bounds for $\|\widehat{P}-P\|_F \ \& \ |L(\widehat{S}|\widehat{P})-L_{\mathrm{ideal-MSE}}(\widehat{S})|$
- Competitive empirical performance

No auxiliary information on rows or columns needed!

2. Solve modified version of standard MC problem:

$$\widehat{S} = \underset{Z \in [-1,1]^{m \times n}}{\operatorname{argmin}} \underbrace{\{L(Z|\widehat{P}) + \lambda \|Z\|_*\}}_{\text{nuclear norm}} \quad \text{Convex program}$$

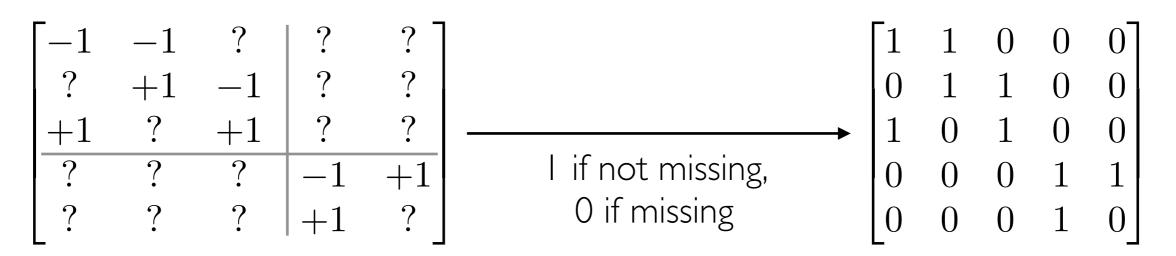
$$\underset{\text{(encourages low rank)}}{\text{nuclear norm}}$$

where

$$L(Z|\widehat{P}) = \frac{1}{mn} \sum_{(i,j)\in\Omega} \frac{(X_{i,j} - Z_{i,j})^2}{\widehat{P}_{i,j}}$$

Standard approach uses $L_{\text{naive-MSE}}(Z)$ instead of $L(Z|\widehat{P})$ (Mazumder et al 2010)

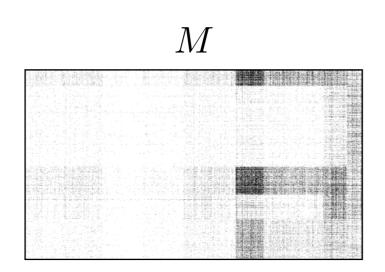
What do missingness patterns look like?



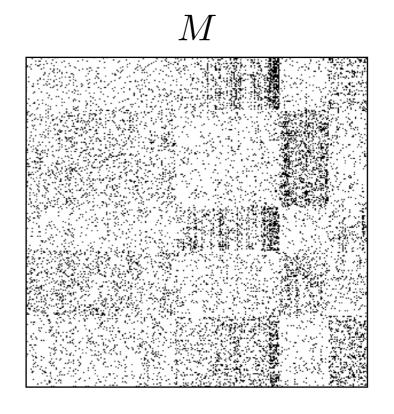
Revealed ratings matrix X

Missingness matrix M

Missingness Matrices in Real Data



MovieLens (Harper and Konstan 2015)



Coat (Schnabel et al 2016)

There is block structure

Low rank

(Can also show that there is topic modeling structure)

Goal: Given M, estimate P under low nuclear norm structure (will in some sense also cover low rank)

General Low Nuclear Norm Structure (Davenport et al 2014)

Parameterize P with user-specified link function $\sigma: \mathbb{R} \to [0,1]$

$$P_{i,j} = \sigma(A_{i,j})$$

Example: standard logistic function $\sigma(x) = 1/(1 + e^{-x})$

for parameter matrix $A \in \mathbb{R}^{m \times n}$

Idea: impose constraints on A instead of P (helpful for theoretical analysis)

Assumption AI: There exists $\theta > 0$ s.t. $||A||_* \le \theta \sqrt{mn}$ (low nuclear norm)

Assumption **A2**: There exists $\alpha>0$ s.t. $\max_{i,j}|A_{i,j}|\leq \alpha$

(bounded probabilities $P_{i,j} \in [\sigma(-\alpha), \sigma(\alpha)]$)

Block structure, clustering, topic models are all special cases!

Any bounded low rank A satisfies AI and A2

Technical detail: with some changes to theory & algorithm, can make upper bound

Algorithm: IbitMC (Davenport et al 2014)

1. Solve a nuclear-norm-regularized maximum likelihood estimation problem:

$$\widehat{A} = \underset{A \in \mathbb{R}^{m \times n}}{\operatorname{argmax:}} \sum_{i=1}^{m} \sum_{j=1}^{n} \underbrace{M_{i,j} \log \sigma(A_{i,j}) + (1 - M_{i,j}) \log(1 - \sigma(A_{i,j}))}_{i=1}$$
 subject to:
$$\underbrace{\|A\|_* \leq \theta \sqrt{mn}, \ \max_{i,j} |A_{i,j}| \leq \alpha}_{i,j} \text{ constraints correspond to Assumptions AI & A2}}_{\text{Assumptions AI & A2}}$$

Convex program depending on choice of σ

2. Estimate propensity scores as follows:

$$\widehat{P}_{i,j} = \sigma(\widehat{A}_{i,j})$$

Davenport et al developed this algorithm for binary matrix completion with MCAR entries

Key idea: apply matrix completion algorithm to fully-observed matrix M to estimate P

We are debiasing matrix completion with more matrix completion!

Can also use other algorithms designed for matrix completion aside from IbitMC to estimate P, such as collaborative filtering

(Technically, we are doing matrix denoising not matrix completion for M)

Theoretical Guarantees

Theorem (IbitMC): Choose link $\sigma(x) = 1/(1 + e^{-x})$.

Under assumptions AI and A2, if # rows m & # cols n are large enough, then with high probability, we simultaneously have:

$$\frac{1}{mn} \sum_{i,j} (\widehat{P}_{i,j} - P_{i,j})^2 \le \mathcal{O}\left(\theta \left[\frac{1}{\sqrt{m}} + \frac{1}{\sqrt{n}}\right]\right)$$
$$|L(\widehat{S}|\widehat{P}) - L_{\text{ideal-MSE}}(\widehat{S})| \le \mathcal{O}\left(\frac{\sqrt{\theta}}{[\sigma(-\alpha)]^2} \left[\frac{1}{m^{1/4}} + \frac{1}{n^{1/4}}\right]\right)$$

 $\begin{array}{cc} m & \\ & \text{if} \\ m \asymp n \end{array}$

Theorem (CF): If there's clustering structure (across rows/columns), we can get faster debiasing rate $m^{-1/2}$ instead of $m^{-1/4}$ (uses collaborative filtering to estimate P instead of I bitMC)

(The collaborative filtering results are in a forthcoming longer version of the paper)

Matrix Completion (MovieLens, Coat)

Coat has its own train/test split Experiment (per dataset): MovieLens: 90/10 split with 10 experimental repeats

- Separate revealed entries into train/test split
- 5-fold cross-validation for hyperparameter selection
- Evaluate prediction error on test entries

Main findings:

- IbitMC debiasing tends to outperform naive Bayes and logistic regression debiasing
- IbitMC debiasing often improves existing methods, at times yielding the best or nearly the best accuracies

Algorithm	Coat		MovieLens-100k		
Augorium	MSE	SNIPS-MSE	MSE	SNIPS-MSE	
PMF	1.000	1.051	0.896 ± 0.013	0.902 ± 0.013	
NB-PMF	1.034	1.117	N/A	N/A	
LR-PMF	1.025	1.107	N/A	N/A	
1BITMC-PMF	0.999	1.052	0.845 ± 0.012	0.853 ± 0.011	
SVD	1.203	1.270	0.862 ± 0.013	0.872 ± 0.012	
NB-SVD	1.246	1.346	N/A	N/A	
LR-SVD	1.234	1.334	N/A	N/A	
1BITMC-SVD	1.202	1.272	0.821 ± 0.011	0.832 ± 0.011	
SVD++	1.208	1.248	0.838 ± 0.013	0.849 ± 0.012	
NB-SVD++	1.488	1.608	N/A	N/A	
LR-SVD++	1.418	1.532	N/A	N/A	
1BITMC-SVD++	1.248	1.274	0.833 ± 0.012	0.843 ± 0.011	
SOFTIMPUTE	1.064	1.150	0.929 ± 0.015	0.950 ± 0.015	
NB-SOFTIMPUTE	1.052	1.138	N/A	N/A	
LR-SOFTIMPUTE	1.069	1.156	N/A	N/A	
1BITMC-SOFTIMPUTE	0.998	1.078	0.933 ± 0.014	0.953 ± 0.014	
MAXNORM	1.168	1.263	0.911 ± 0.011	0.925 ± 0.011	
NB-MAXNORM	1.460	1.578	N/A	N/A	
LR-MAXNORM	1.537	1.662	N/A	N/A	
1BITMC-MAXNORM	1.471	1.590	0.977 ± 0.017	0.992 ± 0.019	
WTN	1.396	1.509	0.939 ± 0.013	0.952 ± 0.013	
NB-WTN	1.329	1.437	N/A	N/A	
LR-WTN	1.340	1.448	N/A	N/A	
1BITMC-WTN	1.396	1.509	0.934 ± 0.013	0.946 ± 0.013	
EXPOMF	2.602	2.813	2.461 ± 0.077	2.558 ± 0.083	

Conclusions & Future Work

Main takeaways:

- We recommend using IbitMC to estimate propensity scores if:
 - I. You don't want parametric assumptions
 - 2. Your data matrix is sufficiently large (e.g., at least hundreds of rows/cols)
- Other MNAR matrix completion methods lack debiasing guarantees; some do not estimate the propensity score matrix (possibly useful for other tasks)

Future directions:

- More robust way to debias MC using propensity score estimates (that neatly handles propensity scores that are 0)
- Handling the case entries are not revealed independently (revealing one entry makes another more/less likely to be revealed)
- Debiasing guarantees for prediction tasks using MNAR imputed features
- Any benefits to using this approach in causal reasoning context?