

90-906
Spring 2000

FINAL

SOLUTIONS

1. a) We compare the regressions
on pg 2 and pg 5:

$$F\text{-stat} = \frac{(7.67 \times 10'' - 7.45 \times 10'') / 3}{7.45 \times 10'' / 852}$$
$$= 8.39$$

Pick 1% signif and look on the
 $F_{3, \infty}$ table to get 3.80.

We reject linearity & are 99%
confident that the true model is
non-linear.

- 1 b) Salary paths diverge if the coefficient on $\text{age} \times \text{yr school}$ is non-zero. So, we test the null hypothesis that $H_0: \beta_{\text{AGEYRS}} = 0$.

$$t\text{-stat} = \frac{75 - 0}{48} = 1.57$$

At the 5% signif level, the table value is 1.96. We do not reject. There is not enough information in these data to be 95% confident that the salary paths diverge.

1 c) Well,
$$\frac{\partial \text{Salary}}{\partial \text{age}} = \beta_{\text{age}} + 2/\beta_{\text{age}^2} \cdot \text{age} + \beta_{\text{age}^2} \cdot \gamma_{\text{school}}$$

(Note: In light of the results of the previous section, you could omit the β_{age^2} term & use the results on page 4).

So,
$$\widehat{\frac{\partial \text{Salary}}{\partial \text{age}}} = 1067 + 2(-23.4)(37.5) + 75.2(14) = 365$$

We need the variance of $\frac{\partial \text{Salary}}{\partial \text{age}}$, which we get from:

$$\begin{aligned} V\left(\frac{\partial \text{Salary}}{\partial \text{age}}\right) &= V\left((0 \ 0 \ 0 \ 1 \ 75 \ 14) \begin{pmatrix} \hat{\beta}_{\text{age}} \\ \hat{\beta}_{\text{age}^2} \\ \hat{\beta}_{\text{age}^3} \\ \hat{\beta}_{\text{age}^4} \\ \hat{\beta}_{\text{age}^5} \\ \hat{\beta}_{\text{age}^6} \end{pmatrix}\right) \\ &= (0 \ 0 \ 0 \ 1 \ 75 \ 14) V(\hat{\beta}_{\text{OLS}}) \begin{pmatrix} 1000 \\ 75 \\ 14 \end{pmatrix} \\ &= (1 \ 75 \ 14) \begin{bmatrix} 79,1679 & -4,173 & -32,885 \\ -4,173 & 48.8 & 17.9 \\ -32,885 & 17.9 & 2308 \end{bmatrix} \begin{pmatrix} 1 \\ 75 \\ 14 \end{pmatrix} \\ &= 9407 \end{aligned}$$

95% CI:
$$\frac{\partial \text{Salary}}{\partial \text{age}} = 365 \pm 1.96 \sqrt{9407} = 365 \pm 190$$

1 d) Salary is maximized where $\frac{\partial S_{\text{age}}}{\partial \beta_{\text{age}}} = 0$

$$\frac{\partial S_{\text{age}}}{\partial \beta_{\text{age}}} = \beta_{\text{age}} + 2\beta_{\text{age}^2} \text{age} + \beta_{\text{age}^3} \text{yrschool} - 1 = 0$$

$$\widehat{\text{age}} = \frac{1}{-2\beta_{\text{age}^2}} \left(\beta_{\text{age}} + \beta_{\text{age}^3} \text{yrschool} \right)$$

Note: Again, β_{age^3} could be omitted.

$$\widehat{\text{age}} = \frac{1}{-2(-23.4)} \left(1067 + 75.2(14) \right)$$

$$\widehat{\text{age}} = 45.3$$

We must use the delta-method to calculate $V(\widehat{\text{age}})$.

$$\frac{\partial \widehat{\text{age}}}{\partial \beta_{\text{age}}} = -\frac{1}{2\beta_{\text{age}^2}} \quad \frac{\partial \widehat{\text{age}}}{\partial \beta_{\text{age}^3}} = -\frac{1}{2\beta_{\text{age}^2}} \cdot \text{yrschool}$$

$$\frac{\partial \widehat{\text{age}}}{\partial \beta_{\text{age}^2}} = \left(\beta_{\text{age}} + \beta_{\text{age}^3} \text{yrschool} \right) \frac{1}{2(\beta_{\text{age}^2})^2}$$

$$\widehat{\frac{\partial \widehat{\text{age}}}{\partial \beta_{\text{age}}}} = \frac{1}{-2(-23.4)} = \frac{1}{46.8} = 0.021$$

$$\widehat{\frac{\partial \widehat{\text{age}}}{\partial \beta_{\text{age}^3}}} = 0.021(14) = 0.30$$

$$\widehat{\frac{\partial \widehat{\text{age}}}{\partial \beta_{\text{age}^2}}} = (1067 + 75.2(14)) \frac{1}{2(-23.4)^2} = 1.94$$

700, 500, 1000, 1000, 1000

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$$V(\bar{y}_c) = (0.021 \quad 1.94 \quad 0.30) \begin{bmatrix} 791679 & -4173 & -32885 \\ -4173 & 48.8 & 17.9 \\ -32885 & 17.9 & 2308 \end{bmatrix} \begin{pmatrix} 0.021 \\ 1.94 \\ 0.30 \end{pmatrix}$$
$$= 6.98$$

@ 95% CI: $99e = 45.3 \pm 1.96\sqrt{6.98} = 45.3 \pm 5.2$

$$2. \quad a) \quad \hat{\beta}_{OLS} = (X'X)^{-1} X'Y$$

$$= \frac{1}{\sum X^2} \sum XY$$

$$V(\hat{\beta}_{OLS}) = \frac{e'e}{N} (X'X)^{-1}$$

$$= \frac{e'e}{N} \frac{1}{\sum X^2}$$

X	Y	X ²	XY
-3	-3	9	-9
-2	-1	4	-2
-1	-2	1	-2
0	-1	0	0
1	1	1	1
2	0	4	0
3	5	9	15
		<hr/>	<hr/>
		28	29

$$\hat{\beta}_{OLS} = 29/28 \approx 1$$

X	Y	e	e ²
-3	-3	0	0
-2	-1	-1	1
-1	-2	1	1
0	-1	1	1
1	1	0	0
2	0	2	4
3	5	-2	4

$$\frac{e'e}{N} = \frac{11}{7} = 1.57$$

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$$\hat{\beta}_{00} = 1 \quad V(\hat{\beta}_{00}) = 1.57/28 = 0.056$$

$$H_0: \beta = 0 \quad t\text{-stat} = \frac{1}{\sqrt{0.056}} = 4.2$$

We will reject this null at any reasonable signif level.

2. b) Most regress e^z on $X-\bar{X}$ and $X^2-\bar{X}^2$

e^z	$X-\bar{X}$	$X^2-\bar{X}^2$	$(\bar{X}=0, \bar{X}^2=4)$
0	-3	5	
1	-2	0	
1	-1	-3	
1	0	-4	
0	1	-3	
4	2	0	
4	3	5	

$$Z = \begin{bmatrix} 1 & X-\bar{X} & X^2-\bar{X}^2 \end{bmatrix}$$

$$\hat{\alpha}_{OLS} = (Z'Z)^{-1} Z'e^z$$

$$Z'Z = \begin{bmatrix} 7 & 0 & 0 \\ 0 & 28 & 0 \\ 0 & 0 & 88 \end{bmatrix}$$

$(X-\bar{X})^2$	$(X^2-\bar{X}^2)^2$	$(X-\bar{X})(X^2-\bar{X}^2)$
9	25	-15
4	0	0
1	9	3
0	16	0
1	9	-3
4	0	0
9	25	15
<hr/> 28	<hr/> 88	<hr/> 0

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e^z	$(x-\bar{x})$	$(x^2-\bar{x}^2)$	$e^z(x-\bar{x})$	$e^z(x^2-\bar{x}^2)$
0	-3	5	0	0
1	-2	0	-2	0
1	-1	-3	-1	-3
1	0	-4	0	-4
0	1	-3	0	0
4	2	0	8	0
4	3	5	12	20
			<hr/> 17	<hr/> 13

$$Z'e^z = \begin{bmatrix} 11 \\ 17 \\ 13 \end{bmatrix}$$

$$(Z'Z)^{-1}Z'e^z = \begin{bmatrix} \frac{1}{7} & 0 \\ 0 & \frac{1}{28} \\ 0 & 0 \end{bmatrix} \begin{bmatrix} 0 \\ 0 \\ \frac{20}{13} \end{bmatrix} \begin{bmatrix} 11 \\ 17 \\ 13 \end{bmatrix} = \begin{bmatrix} 1.57 \\ 0.61 \\ 0.15 \end{bmatrix}$$

e^z	$x-\bar{x}$	$x^2-\bar{x}^2$	$1.57 + 0.61(x-\bar{x}) + 0.15(x^2-\bar{x}^2)$	Resid
0	-3	5	0.49	-0.49
1	-2	0	0.35	0.65
1	-1	-3	0.51	0.49
1	0	-4	0.97	0.03
0	1	-3	1.73	-1.73
4	2	0	2.79	1.21
4	3	5	4.15	-0.15

$$SSE = (-0.49)^2 + (0.65)^2 + (0.49)^2 + 0.03^2 + (-1.73)^2 + 1.21^2 + (-0.15)^2$$

$$= 5.38$$

$$\bar{e}^2 = 1.57$$

$$TSS = (0-1.57)^2 + (1-1.57)^2 + (1-1.57)^2 + (1-1.57)^2 + (0-1.57)^2 + (4-1.57)^2 + (4-1.57)^2$$

$$= 17.71$$

$$R^2 = \frac{17.71 - 5.38}{17.71} = 0.70$$

$$NR^2 = 7 \cdot 0.70 = 4.9$$

χ^2 -table at 5% is 5.99

So, we would accept the null of homoscedasticity.

2 c) Let's use robust std errors:

$$\widehat{V(\hat{\beta}_{OLS})} = (X'X)^{-1} X' \text{diag}(e^2) X (X'X)^{-1}$$

$$= \frac{1}{\sum X^2} \sum X^2 e^2 \frac{1}{\sum X^2}$$

$$= \frac{1}{28} \sum X^2 e^2 \frac{1}{28}$$

e^2	X^2	$e^2 X^2$
0	9	0
1	4	4
1	1	1
1	0	0
0	1	0
4	4	16
4	9	36
		<hr/>
		57

$$\widehat{V(\hat{\beta}_{OLS})} = \frac{1}{28} 57 \frac{1}{28} = 0.073$$

$$t\text{-stat} = \frac{1}{\sqrt{0.073}} = \frac{1}{0.27} = 3.7$$

Still reject null of no effect X on Y.

- 3 e) Recall the formula for omitted variables:

$$\begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} = \begin{bmatrix} X_1'X_1 & X_1'X_2 \\ X_2'X_1 & X_2'X_2 \end{bmatrix}^{-1} \begin{pmatrix} X_1'Y \\ X_2'Y \end{pmatrix} = \dots$$

$$E \begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \begin{bmatrix} X_1'X_1 & X_1'X_2 \\ X_2'X_1 & X_2'X_2 \end{bmatrix}^{-1} \begin{pmatrix} X_1'X_3 \\ X_2'X_3 \end{pmatrix} \beta_3$$

The fact that $Cov(X_1, X_3) > 0$ only establishes that $X_1'X_3 > 0$. This is little help in signing the bias. The $(X'X)^{-1}$ matrix may scramble $X_1'X_3$ and $X_2'X_3$ making the bias on β_1 +/- or 0.

- b) Similarly, knowing $Cov(X_1, X_3) > 0$ and $Cov(X_2, X_3) > 0$ does not help much. Now, we know that $X_1'X_3 > 0$, $X_2'X_3 > 0$. Of course, we also know that the 11 and 22 elements of $(X'X)^{-1}$ above are > 0 . But, if the 12 and 21 elements are < 0 , still the bias in β_1, β_2 could be +/- or 0.

3. c) Now, we can sign the bias:

$$E \begin{pmatrix} \hat{\beta}_{1OLS} \\ \hat{\beta}_{2OLS} \end{pmatrix} = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \begin{bmatrix} X_1'X_1 & \\ & X_2'X_2 \end{bmatrix}^{-1} \begin{pmatrix} X_1'X_3 \\ X_2'X_3 \end{pmatrix} \beta_3$$

$$= \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \begin{pmatrix} X_1'X_3/X_1'X_1 \\ X_2'X_3/X_2'X_2 \end{pmatrix} \beta_3$$

Both $\hat{\beta}_{1OLS}$ and $\hat{\beta}_{2OLS}$ are biased up
since $X_1'X_3, X_1'X_1, X_2'X_3, X_2'X_2 > 0$.

d) Again, we can sign the bias:

$$E \begin{pmatrix} \hat{\beta}_{1OLS} \\ \hat{\beta}_{2OLS} \end{pmatrix} = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \frac{1}{X_1'X_1 X_2'X_2 - (X_1'X_2)^2} \begin{bmatrix} X_1'X_1 & -X_1'X_2 \\ -X_1'X_2 & X_2'X_2 \end{bmatrix} \begin{pmatrix} X_1'X_3 \\ X_2'X_3 \end{pmatrix} \beta_3$$

$$= \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \frac{\beta_3}{\dots} \begin{pmatrix} X_1'X_1 X_2'X_3 - X_1'X_2 X_1'X_3 \\ X_1'X_2 X_2'X_3 - X_1'X_2 X_1'X_3 \end{pmatrix}$$

Since β_3 and $\frac{1}{\dots}$ are > 0 and $X_1'X_1, X_1'X_3, X_2'X_2, X_2'X_3 > 0$ and $X_1'X_2 < 0$, both $\hat{\beta}_{1OLS}$ and $\hat{\beta}_{2OLS}$ are biased up.

$$\begin{aligned}
 3e) \quad E\left(\begin{matrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{matrix}\right) &= \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \begin{bmatrix} 4 & 3 \\ 3 & 10 \end{bmatrix}^{-1} \begin{pmatrix} 2 \\ -1 \end{pmatrix} \\
 &= \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \frac{1}{40-9} \begin{bmatrix} 10 & -3 \\ -3 & 4 \end{bmatrix} \begin{pmatrix} 2 \\ -1 \end{pmatrix} \\
 &= \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + \frac{1}{31} \begin{pmatrix} 23 \\ -10 \end{pmatrix}
 \end{aligned}$$

$\hat{\beta}_1$ is biased up by $23/31$
 $\hat{\beta}_2$ is biased down by $10/31$

4 a) Yes, it is.

$$\begin{aligned}\overline{\Delta \ln Y} &= \frac{1}{T-1} \sum_{t=2}^T \ln Y_t - \ln Y_{t-1} \\ &= \frac{1}{T-1} \left(\underbrace{\ln Y_2 - \ln Y_1}_{=0} + \underbrace{\ln Y_3 - \ln Y_2}_{=0} + \ln Y_4 - \ln Y_3 + \dots + \ln Y_T - \ln Y_{T-1} \right)\end{aligned}$$

$$= \frac{1}{T-1} (\ln Y_T - \ln Y_1)$$

$$= \frac{1}{T-1} (\beta_1 + \beta_2 T + \epsilon_T - (\beta_1 + \beta_2 1 - \epsilon_1))$$

$$= \frac{1}{T-1} (\beta_2 (T-1) + \epsilon_T - \epsilon_1)$$

$$= \beta_2 + \frac{1}{T-1} (\epsilon_T - \epsilon_1)$$

So, $\overline{\Delta \ln Y}$ is unbiased $E(\overline{\Delta \ln Y}) = \beta_2$.

Also, its variance $\rightarrow 0$ as $T \rightarrow \infty$:

$$V(\overline{\Delta \ln Y}) = \frac{2}{(T-1)^2} \sigma^2$$

4 b) No, it is not. When the model is differenced:

$$\Delta \ln Y_t = \beta_1 + \epsilon_t - \epsilon_{t-1}$$

We introduce serial correlation in the errors. So, the usual OLS estimator will be CAN but usual OLS std errors will be wrong.

4 c) It is given above $\frac{2}{(T-1)}\sigma^2$.

Also, we could use the estimator for the variance of OLS in the GLS model:

$$V(\hat{\beta}_{GLS}) = (X'X)^{-1} X' \Sigma X (X'X)^{-1}$$

$$X'X = (T-1)$$

$$X' \Sigma X = (1, \dots, 1) \begin{bmatrix} 2\sigma^2 - \sigma^2 & 0 & 0 \\ -\sigma^2 & 2\sigma^2 - \sigma^2 & 0 \\ & -\sigma^2 & 2\sigma^2 \\ 0 & & \ddots & -\sigma^2 \\ & 0 & & -\sigma^2 & 2\sigma^2 \end{bmatrix} \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix}$$

~~$$X'X = (T-1) \sigma^2$$

$$X' \Sigma X = (T-1) \sigma^2$$

$$(X' \Sigma X) (X'X)^{-1} = \sigma^2 / (T-1)$$~~

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$$X' \Sigma X = (\sigma^2 \ 0 \ 0 \ \dots \ 0 \ \sigma^2) \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} = 2\sigma^2$$

$$\begin{aligned} (X'X)^{-1} X' \Sigma X (X'X)^{-1} &= \frac{1}{T-1} 2\sigma^2 \frac{1}{T-1} \\ &= 2\sigma^2 / (T-1)^2 \end{aligned}$$

4 d) Yes, by the G-M Theorem.