Distress Dispersion and Systemic Risk in Networks*

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ABSTRACT

I present a model in which heterogeneity in financial distress endogenously generates inefficiencies in network formation, creating excessive systemic risk. Financial firms face costly liquidation and strategically trade assets, thereby forming links. A link with a distressed firm can be socially costly as it increases system-wide liquidation risk. When the cross-sectional dispersion of distress is high, the network composition is distorted in two ways: it features too many links with distressed firms and too few risk-sharing links among non-distressed firms. This inefficiency arises from an externality as bilateral contract terms are not contingent on links faraway in the network. Using insights from the model, I discuss policy implications for financial stability. I also show empirical evidence that the distress dispersion across financial firms provides a novel indicator for systemic risk.

Keywords: Financial network formation, systemic risk, financial distress, network externality.

1 Introduction

The interconnectedness of financial institutions is a key feature of the modern financial system. Linkages are formed by a diverse range of transactions and contracts that connect firms to each other. A growing literature identifies these linkages as a major source of systemic risk (e.g. Allen and Gale (2000), Caballero and Simsek (2013), Brunnermeier (2009), and Acemoglu, Ozdaglar, and Tahbaz-Salehi (2015)). The insights are evident in the financial crisis: initial losses caused the financial distress of a few firms, which then spread via the links that connect the distressed firms with otherwise healthy ones, resulting in systemic failures. Yet, these studies analyze contagion in given network structures and do not consider firms’ strategic formation of links.

In this paper, I focus on endogenous linkage formation which allows firms to strategically build connections for profit and risk diversification purposes. A recent literature examines linkage formation among homogeneous firms and concludes that either over- or under-connections prevail in the financial system (e.g. Castiglioni and Navarro (2011) and Farboodi (2014)). In contrast, this paper studies the linkage formation among firms differing in financial distress levels. Such framework provides novel implications for efficiency and systemic risk by generating over- and under-connections simultaneously.

I show that the endogenously formed network features inefficiencies and leads to systemic risk as measured by the probability of joint failures. A link between two non-distressed firms creates gains from risk-sharing, whereas a link with a distressed firm can be socially costly as it increases systemic risk through balance sheet interdependence. I find that, when the dispersion of distress is high, the network composition is distorted in two ways: there are too many links with distressed firms and too few risk-sharing links among non-distressed firms. The inefficiency arises as firms write bilateral contracts that are not contingent on the entire network structure. Hence, the non-distressed firms have incentives to link with distressed firms for profit, while failing to internalize negative spillovers. Such inefficient network generates contagion and loss in risk-sharing, creating excessive systemic risk. By embedding heterogeneity as a new dimension of links, my model provides unique predictions on the efficiency of network composition.

In my model, financial firms face costly liquidation risks and strategically trade assets,
thereby forming links. There are a finite number of firms financed by short-term debt and each invests in a long-term asset. A random fraction of the asset is liquid and can be used to repay debt. As in Allen, Babus, and Carletti (2012), if the amount of liquid asset falls short of the debt level, a costly liquidation of the illiquid asset is triggered. To hedge the idiosyncratic liquidation risk, firms can strategically enter into bilateral forward contracts to trade liquid assets. A two-sided link in a network is formed when both parties decide to purchase a fraction of each other’s liquid asset claims. A key feature is that firms differ ex ante in how liquid the asset is expected to be, which generates cross-sectional heterogeneity in financial distress levels. Differences in asset liquidity also implies a price of trade in each contract. Motivated by the incomplete contracts literature, I assume that prices in the bilateral trades are not contingent on the entire network structure. Specifically, I consider local contingency, that is, prices are contingent on which firms the two parties directly trade with. Given the network formed, the liquid asset holding of a firm depends not only on who its direct counterparties are, but rather on the entire network structure. As a benchmark for efficiency, I solve for the optimal network that minimizes total bank failures.

The pairwise stable network formed in equilibrium can be inefficient relative to the optimal benchmark: there can be excess links with distressed firms and insufficient risk-sharing links among non-distressed firms. When distress dispersion is high across firms, the optimal network requires that the non-distressed firms form risk-sharing links and that the most distressed firm be isolated. In comparison, the equilibrium network with four or more firms shows that the distressed firm is always connected with the most liquid firm. The suboptimal link between the liquid and the distressed firm (“distress link” hereafter) transmits risky assets in the network and leads to systemic risk, measured by the probability that all firms fail at the same time.

The inefficiency is caused by network externalities. Linking with a distressed firm potentially avoids liquidation, thus is ex ante profitable for the most liquid firm. However, when a firm is too distressed, linking with it can be socially costly because it contaminates the balance sheets of other firms in the network. Hence the decision of a liquid firm to form a distress link imposes a network externality, as distressed assets are then shared jointly by all connected firms.

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2 A firm with a low level of liquid asset has difficulty in repaying short-term debt and hence is distressed.
3 Following Cabrales, Gottardi, and Vega-Redondo (2014), I model this balance sheet interdependence as an iterative swap process which represents asset securitization.
firms. The distress link increases the risk of contagion, which in turn reduces risk-sharing participation among non-distressed firms. As such, two forces reinforce and lead to inefficiency: the transmission of distressed assets that should have been isolated and the insufficient risk-sharing among non-distressed firms.

The necessary ingredients for the externalities are interconnectedness, distress heterogeneity, and local contingency. Interconnectedness transmits risky assets, thereby enabling spillovers. Firm heterogeneity generates distress dispersion and different incentives to form links. When there are only two firms or multiple identical firms, there is no externality. However, when there are trades between multiple firms differing in distress levels, the most liquid firm can profit from trading with the distressed firm and can shift risks away to its direct and indirect counterparties; hence, the most liquid firm has a greater incentive to link with the distressed firm than is socially desirable. But interconnectedness and heterogeneity are not enough. The externalities are not internalized because of local contingency. Firms that bear the externalities cannot jointly give incentives to the liquid firm via contingent payments. This occurs as long as one of the indirect counterparties of the most liquid firm cannot condition payments on the distress link. Thus, the liquid firm fails to internalize negative spillovers and forms the inefficient distress link.

While the prior literature largely focuses on the average soundness of the financial sector, my second primary result identifies a novel indicator for the level of network inefficiency: the distress dispersion across financial firms. In my model, inefficiency arises when the distress dispersion is sufficiently high and increases with the level of dispersion thereafter. This positive relation is due to changes in network composition. When distress dispersion is higher, a wider cross-sectional distribution implies more distressed firms in the left tail and more liquid ones in the right tail. It is precisely then that the most liquid firm has an incentive to form the socially costly distress link. Hence the disparity between individual and social incentives for forming a distress link is greater, which crowds out valuable risk-sharing links and increases inefficiency.

Using insights from the model, I discuss policy implications for financial stability. The links with distressed firms in the model can be interpreted as acquisitions of distressed firms.

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4Atkeson, Eisfeldt, and Weill (2014) measure the median Distance to Insolvency of largest financial firms based on the Leland’s model of credit risk. Rampini and Viswanathan (2014) argue that the net worth of (representative) financial intermediaries is an important state variable affecting the cost of financing. Gilchrist and Zakraysek (2012) show that the average credit spreads on outstanding corporate bonds has predictive power for economic activity.
This interpretation is reasonable because distressed financial firms are commonly acquired by healthier institutions in the same industry.\(^5\) More than 1000 distressed financial firms were acquired during 2000-2013, including Countrywide Financial Corp. and Riggs Bank. The asset size of these acquisitions was $2.2 trillion, about half the size of all current banking deposits. Despite the fact that acquisitions are a prevailing regulatory approach to improve financial stability,\(^6\) my findings imply that excess acquisitions may emerge precisely when more banks are distressed, thus increasing systemic risk rather than reducing failures.

In the context of acquisitions of distressed firms, I show an acquisition tax that varies with the distress dispersion can prevent the inefficient acquisitions and reduce total liquidation costs. Based on this result, regulators can restore efficiency by supervising the acquisitions of distressed firms and using the purchase and assumption (P&A) method for distress resolution. In a model extension that allows for the analysis of ex post policies, I show that if the excess acquisitions are not banned ex ante, the too-connected-to-fail problem arises. In such a scenario, government bailout or subsidized acquisitions are ex post optimal remedies, thereby rationalizing the government interventions observed during the recent financial crisis.

Finally, I provide empirical evidence that the distress dispersion across financial institutions provides a novel indicator for systemic risk. Following Laeven and Levine (2009), I measure distress by estimating Z-scores of financial firms. The time series of distress dispersion displays large variations over time. Moreover, it has a countercyclical pattern and appears to lead recessions. Consistent with the model predictions, the empirical dispersion series significantly comoves with future economic activities and systemic risk, bank failures, acquisitions of distressed firms, and interbank risk sharing. I run forecasting regressions to evaluate whether the dispersion series conveys new information about aggregate indicators beyond what is contained in the average distress and existing systemic risk measures. The estimates confirm that the dispersion series has high predictive power for future indices of systemic risk.

\(^5\) Acharya, Shin, and Yorulmazer (2010) argue that if a bank needs to restructure or be sold, the potential buyers are generally other banks. Almeida, Campello, and Hackbarth (2011) document that distressed firms are acquired by liquid firms in their industries for financial synergies. Such acquisitions are more likely when industry-level asset specificity is high and firm-level asset specificity is low, which applies to the financial sector.

\(^6\) White and Yorulmazer (2014) provide a summary of resolution options for bank distress/failure. An acquisition “imposes the least cost since the franchise value is preserved, there is no disruption to the bank’s customers or the payment system itself, and there are no fiscal costs.” For this reason, acquisition is the primary choice by resolution authorities whenever there are willing acquirers.
1.1 Related Literature

This paper builds on network theory and its applications in economics and finance.\(^7\) Pioneered by Allen and Gale (2000), a growing literature argues that certain network structures among financial institutions can lead to risks of contagion.\(^8\) While powerful for analyzing how risks propagate under different connection properties, this stream of research treats the network structures as given. My paper studies network formation, hence contributes to the analysis of how links evolve in response to changes in policies or aggregate conditions.

The main contribution of this paper is to embed distress heterogeneity in linkage formation and to study the implications on efficiency and systemic risk. As such, my paper belongs to the recent literature on financial network formation, which examines how inefficient networks form due to various frictions.\(^9\) In particular, Castiglionesi and Navarro (2011) demonstrate network fragility when undercapitalized banks gamble with depositors’ money. Similarly, Zawadowski (2013) studies a type of risk shifting stemming from banks’ underinsurance of counterparty risk. Moreover, Gofman (2011) and Farboodi (2014) highlight that bargaining friction and intermediation can result in welfare loss.

In the network formation literature, my paper is closest to Farboodi (2014) who illustrates that a core-periphery intermediation structure arises inefficiently due to a lending constraint and the opportunity to earn intermediation spreads. While my paper also generates excessive systemic risk due to certain types of inefficient links, I differ by studying linkage formation among firms differing in financial distress levels. Inefficiency arises from the incentive of liquid firms to link with distressed firms for profit under contract incompleteness. Moreover, I model links on the asset side of the balance sheet. The resulting asset cross-interdependence structure can be used to evaluate acquisition regulations. Finally, the novel finding that the distress dispersion is a critical state variable allows for a closer link to the data in forecasting systemic risk.

The key friction underlying the network inefficiency in my model is the failure to offer

\(^7\) See surveys by Jackson (2003, 2008) and Allen and Babus (2009).


incentives conditional on the entire network structure. In this sense, my paper is related to the literature on incomplete contracts.\footnote{See for example Hart and Moore (1988, 1999), Tirole (1999), Maskin and Tirole (1999), and Segal (1999).} From Hart and Moore (1988), agents cannot write contracts contingent on states that cannot be clearly specified, even if the states are perfectly foreseeable. The reason is that the states written in the contracts must be verifiable in court. In my setting, given that the links entered by other firms are not specifiable or verifiable, bilateral prices are contingent only on who the two firms directly trade with. This assumption is in line with Acemoglu, Ozdaglar, and Tahbaz-Salehi (2014) who show that inefficient networks can emerge in interbank lending markets with contingency debt covenants.

Finally, this paper adds to the studies on the trade-off between diversification and contagion. Banal-Estanol, Ottaviani, and Winton (2013) evaluate conglomeration with default costs in terms of this trade-off. I follow Cabrales, Gottardi, and Vega-Redondo (2014) and study the trade-off in a network setting. Acharya (2009), Wagner (2010), Ibragimov, Jaffee, and Walden (2011), and Castiglionesi and Wagner (2013) show that diversification may lead to greater systemic risk as banks tend to over-diversify by holding similar portfolios. These papers mostly assume costly joint failures among homogeneous agents. My paper complements these studies by showing that links among heterogeneous firms can result in both over and under diversification.

The rest of the paper proceeds as follows. Section 2 lays out the model environment and defines the equilibrium. Section 3 demonstrates the network inefficiencies and investigates the key friction. Section 4 examines the role of distress dispersion on inefficiency. Section 5 discusses the policy implications in the context of acquisitions of distressed firms. Section 6 presents empirical results, and Section 7 concludes. All proofs are in the Appendix.

2 Model

This section describes a model of network formation in which firms strategically trade assets via bilateral forward swap contracts.

2.1 Environment

Consider a four-date economy with a finite number of debt-financed firms, denoted by \( i = 1, \ldots, N \). All agents are risk neutral and there is no discounting.
At date 0, each firm borrows 1 unit of short-term debt from a continuum of creditors and invests in an asset with fixed return $R$. The asset is subject to liquidity risk. A random component $\tilde{a}_i$ becomes liquid at date 2 and can be used to repay debt, whereas the rest $R - \tilde{a}_i$ is illiquid and matures at date 3. Given this financing structure, a maturity mismatch arises. A firm can be interpreted as a financial institution, e.g., an investment firm investing in a certain class of securities, or a commercial bank issuing an unsecured loan.

At date 1, firms observe the vector $\nu$, which is a public signal about how much liquid asset each firm expects to receive. Then they simultaneously decide to enter into bilateral forward swap contracts for risk-sharing purpose, thus forming links. Each forward swap contract promises a claim to a fraction of each other’s liquid assets.

At date 2, firms observe the amount of their liquid assets, given by $\tilde{a}_i = \nu_i + \sigma \varepsilon_i$. The idiosyncratic shock $\varepsilon_i$ is i.i.d. standard normal and is independent of $\nu_i$.\(^{11}\) Firms fulfill the forward swap contracts, and based on the overall linkage structure, firms obtain potentially diversified liquid asset holdings, which they use to repay debt.\(^{12}\) If the liquid asset holdings fall short of debt, the firm liquidates its illiquid asset with a fixed cost $c$, for instance by selling at a discount to industry outsiders as in Shleifer and Vishny (1992).\(^{13}\)

At date 3, if not liquidated, the illiquid component $R - \tilde{a}_i$ of the asset matures. Using this return, the payments associated with the forward swap contracts are paid in full.

Given the signal $\nu$, firms differ at date 1 in the amount of expected liquid asset. This generates heterogeneity in financial distress. A firm with a highly liquid asset has low liquidation risk, and a firm with a highly illiquid asset has difficulty repaying short-term debt and hence is distressed. I follow Roy (1952) and define a distress statistic, $z_i$, as the number of standard deviations that firm $i$ is expected to be away from liquidation ($z_i \equiv \frac{\nu_i - 1}{\sigma}$). A low $z_i$ indicates high asset illiquidity and hence high financial distress. To highlight the role of heterogeneity, I assume that the vector $z$ has mean $\bar{z}$ and is equally spaced with step size $\delta \geq 0$, i.e.,

\[ z_i = \bar{z} + \frac{N + 1 - 2i}{2} \delta, \quad i = 1, \ldots, N. \]  

\(^{11}\) $\tilde{a}_i$ being negative means further liquidity inflow needed in the asset investment.

\(^{12}\) Introducing debt roll-over, renegotiation, or endogenous default boundary do not change the qualitative features. To separate from risk-shifting due to agency conflict between shareholders and depositors (Jensen and Meckling (1976)), limited liability is not particularly imposed for firm owners.

\(^{13}\) The cost can result from deadweight loss in liquidation due to asset specificity, loss of franchise value, or disruption of credit and payment services associated with relationship banking (see White and Yorulmazer (2014)).
\( \bar{z} \) measures the average distance from liquidation. Let \( \bar{z} > 0 \) so that firms invest in positive NPV projects on average. \( \delta \) is proportional to the cross-sectional standard deviation of \( z_i \) and proxies for the degree of distress dispersion.\(^{14}\) I classify distressed firms as follows.

**Definition 1** Firm \( i \) is distressed if \( z_i < -1 \), i.e. in expectation its amount of liquid asset is below the debt level by more than one standard deviation. Firm \( i \) is non-distressed if \( z_i \geq -1 \), and liquid if \( z_i > 1 \).

### 2.2 Network Formation

At date 1, firms strategically decide to enter into bilateral forward swap contracts. In this network formation game, a strategy of firm \( i \) includes a vector \( l_i = (l_{i1}, \ldots, l_{ii-1}, l_{ii+1}, \ldots, l_{iN}) \) and a vector \( p_i = (p_{i1}, \ldots, p_{ii-1}, p_{ii+1}, \ldots, p_{iN}) \). Firm \( i \) proposes to buy \( l_{ij} \in \{0, \bar{l}\} \),\(^{15}\) where \( \bar{l} \in (0, 1) \), fractions of liquid asset from firm \( j \) at date 2, offering to pay a unit price \( p_{ij} \) at date 3. The prices can be made contingent on the links. Similar to the simultaneous announcement game in Myerson (1991), each non-distressed firm simultaneously proposes to contract with other firms. Each distressed firm proposes to contract with only one non-distressed firm. Hence, if firm \( i \) is distressed, the vector \( l_i \) only has one non-zero element.

A contract is signed (a two-sided link is formed) when both firms decide to swap asset claims at the offered prices. Let the matrix \( L \) represent the linkage structure; its element satisfies

\[
L_{ij} = L_{ji} = \min\{l_{ij}, l_{ji}\}. \tag{2}
\]

Firms \( i \) and \( j \) are directly linked (\( L_{ij} = \bar{l} \)) only if \( l_{ij} = l_{ji} = \bar{l} \). This specification ensures that no firms end up being a net asset seller or buyer so each firm still holds one unit of liquid asset. It also captures an important aspect of the OTC derivatives market: firms have large gross notional positions and small net positions. After the asset swaps, each firm holds a non-negative share of its own asset, i.e. \( l_{ii} = 1 - \sum_{j \neq i} l_{ij} \geq 0 \). As such, \( L \) is a symmetric, doubly stochastic matrix by construction.\(^{16}\) When \( l_{ii} = 1 \), firm \( i \) is isolated.

\(^{14}\) I rank firms by \( z_i \) merely for expository purpose. Distress is modeled as exogenous, while in reality firms choose liquidity holding and risk-taking which endogenously determine distress levels. Acharya, Shin, and Yorulmazer (2010) argue that liquid banks hoard cash for potential gains from asset sales. This implies that an otherwise endogenous setting would generate even bigger heterogeneity during an aggregate liquidity shortage.

\(^{15}\) Firms make binary decisions on the linkage formation. From Lemma 1, all results remain if instead \( l_{ij} \in [0, 1) \).

\(^{16}\) A square matrix is doubly stochastic if all its entries are non-negative and the sum of the entries in each of its rows or columns is 1.
The set of \( N \) firms and the links between them define the network. Depending on the distress level of the two connecting firms, the network is composed of \textit{risk-sharing links} which connect two non-distressed firms, and \textit{distress links} which connect a liquid and a distressed firm.

### 2.3 Payoffs and Firm Value

Firms’ liquid asset holdings, denoted by vector \( \tilde{h}_i(\tilde{a}, L) \), depend on not only their direct counterparties, but rather how firms are interconnected. As such, the linkage creates cross-interdependence from the asset side of firms’ balance sheets. I model links via asset swaps because prior studies highlight that correlated portfolio exposures are the main source of systemic risk in the financial sector.\(^{17}\) In addition, asset swaps simplify the calculation of final asset holdings and systemic risk by avoiding kinks in standard cascade models (e.g. Elliott, Golub, and Jackson (2014)).\(^{18}\)

At date 3, firms deliver payment transfers according to the forward swap contracts. Their final payoffs are thus determined by the liquid asset realizations \( \tilde{a} \), the network \( L \), and the prices \( p \), given by \( \Pi(\tilde{a}, L, p) \),

\[
\Pi_i(\tilde{a}, L, p) = \tilde{h}_i(\tilde{a}, L) + R - \tilde{a}_i - 1 - \sum_{j \neq i} (p_{ij} - p_{ji}) L_{ij},
\]

\[
\Pi_i(\tilde{a}, L, p) = \tilde{h}_i(\tilde{a}, L) + R - \tilde{a}_i - 1 - \sum_{j \neq i} (p_{ij} - p_{ji}) L_{ij}, \tag{3}
\]

(3)

Firm value at date 1 is given by taking the expectation of \( \Pi_i(\tilde{a}, L, p) \),

\[
V_i(z, L, p) = \mathbb{E}_1 \left[ \tilde{h}_i(\tilde{a}, L) \right] + R - \nu_i - 1 - \text{Pr}(\tilde{h}_i(\tilde{a}, L) < 1) c - \sum_{j \neq i} (p_{ij} - p_{ji}) L_{ij}. \tag{4}
\]

### 2.4 Bilateral Prices and Asset Swaps

The key features of a network formation game are the payoff functions and the payment transfers. To further specify these terms in my framework, I next discuss assumptions on bilateral prices and asset swaps process.

**Local Contingency** Which firms have the power to decide on a link between two firms is crucial to linkage formation. The bilateral prices allow for transfer payments among firms, which in turn define the decision power to form links. Given that a link \( L_{ij} \) “alters the payoffs

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\(^{17}\)See for example Elsinger, Lehar, and Summer (2006) and DeYoung and Torna (2013).

\(^{18}\)The asset swaps may capture in a broad sense cross holdings of deposits in Allen and Gale (2000).
to others, it seems reasonable to suppose that other firms, especially the \([\text{direct counterparties} \text{ of}] \text{ firms } i \text{ and } j\) should have some say in the formation of a link between \(i\) and \(j\)" (Goyal (2009)). Following this spirit, I assume prices with local contingency.

**Assumption 1 (Local Contingency)** The bilateral price \(p_{ij}\) is contingent on the direct links entered by the two firms. Let \(L_i\) be the \(i\)-th row of \(L\), then

\[
p_{ij}(L_i, L_j, L_k) = p_{ij}(L_i, L_j, \hat{L}_k), \quad \forall k, \forall \hat{L}_k \neq L_k.
\]

(5)

Under Assumption 1, firm \(i\) offers prices based on its own links \(L_i\) and the links of its direct counterparty \(L_j\). Even if firm \(i\) foresees that it indirectly connects to a third firm \(k\) (\(L_{ij} > 0, L_{jk} > 0\)), the price it offers cannot vary with the links of firm \(k\).

Assumption 1 is the key friction in the model. The motivation lies in an inherent feature of the financial industry: when firms write bilateral contracts in an interconnected setting, it is difficult for institutions to specify in every contract detailed contingencies for every possible network structure. One reason for this is that institutions do not publicly disclose the identities of their counterparties. As in Hart (1993), even if the bilateral relations they form could be foreseeable by other institutions, “they might be difficult to specify in advance in an unambiguous manner. [Hence], a contract that tries to condition on these variables may not be enforceable by a court.” This is essentially one example of *incomplete contracts*.\(^{19}\)

**Price Offering Rules** In each bilateral contract, what matters for firm payoffs is the net transfer payment \((p_{ij} - p_{ji})L_{ij}\). The same net payment can be achieved by a continuum of gross payments; hence, to ensure a unique set of equilibrium prices, I assume that buyer \(i\) proposes price \(p_{ij}\) to \(j\) as a *take-it-or-leave-it offer*. The proposed price cannot be lower than firm \(j\)’s reservation price \(p_{jj}\). Formally,

\[
p_{ij} \geq p_{jj}, \quad \forall i \neq j,
\]

(6)

where \(p_{jj}\) equals \(j\)’s outside option when it cannot form any links, i.e.

\[
p_{jj}(z_j) = V_j(z, L, p | L_j = 0).
\]

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\(^{19}\)An alternative motivation relates to transaction costs à la Williamson (1975). As the size and complexity of the network builds up, it would be prohibitively costly to include all possible structures in each contract for every firm. This is consistent with the fact that we do not observe such types of contracts in practice.
Asset Swap Process  I model the cross-interdependence of liquid asset holdings $\tilde{h}(\tilde{a}, L)$ by an iterative asset swap process: firm $i$ swaps liquid asset with its direct counterparties iteratively. Given the linkage matrix $L$, the vector of asset holdings after the first round of swap is $\tilde{h}^{(1)} = L\tilde{a}$. Applying $L$ to $\tilde{h}^{(1)}$ gives the second round of swap, $\tilde{h}^{(2)} = L\tilde{h}^{(1)} = L^2\tilde{a}$, where $L^2$ denotes $L \times L$. Specifically, I assume that the iteration goes on for infinitely many rounds.

Assumption 2 (Iterative Swap Process) Firms swap liquid assets according to the linkage matrix $L$ iteratively for infinite rounds. The final asset holdings $\tilde{h}$ are given by

$$\tilde{h}(\tilde{a}, L) = \lim_{K \to \infty} L^K \tilde{a}. \quad (8)$$

This iterative process is instantaneous and does not affect the payment of prices. It captures the securitization process such as the origination and trades of asset-backed securities.\[20\]

Under Assumption 2, final holdings $\tilde{h}$ depend on the liquid returns of both direct and indirect counterparties. Take for instance a network with $N = 3$ and $L_{12} = L_{23} = \bar{l}$, $L_{13} = 0$. After the first round, $\tilde{h}_1^{(1)} = (1 - \bar{l}) \tilde{a}_1 + \bar{l}\tilde{a}_2$. After infinite rounds, $\tilde{h}_1 = \tilde{h}_2 = \tilde{h}_3 = \frac{1}{3}\tilde{a}_1 + \frac{1}{3}\tilde{a}_2 + \frac{1}{3}\tilde{a}_3$; hence, firm 1 holds $\frac{1}{3}$ shares of $\tilde{a}_3$ even if it does not directly link with firm 3. The following lemma formalizes this property of the final asset holdings.

Lemma 1 (Complete risk-sharing) For all $L$, $\lim_{K \to \infty} L^K$ is doubly stochastic and coincides with complete risk-sharing among all firms connected in the same component,\[21\] i.e. the holdings of each firm are equally weighted by the liquid assets of all firms directly or indirectly connected to it.

From Lemma 1, it is the linkage structure (whether $L_{ij} = 0$ or $L_{ij} > 0$) rather than the amount of swap that determines the final holdings of each firm. Given Lemma 1, the results still hold if instead $l_{ij} \in [0, 1)$, that is, if we allow firms to make linkage decisions in a continuum space. This rationalizes the simplification that $l_{ij}$ is a binary variable. Moreover, the holding of own asset $L_{ii} = 1 - \sum_{j \neq i} L_{ij} \geq 0$ implies that the maximum number of links a firm can form is $1/\bar{l}$. If $1/\bar{l}$ is very large, the number of possible network structures increases exponentially with $N$.\[22\]

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20\) “The possibly iterative procedure through which each firm exchanges assets on its whole array of asset holdings can be viewed as a securitization process of the firm’s claims” (Cabrales, Gottardi, and Vega-Redondo (2014)).

21\) A component of a network is a maximally connected collection of firms: each firm in the component can reach any other firm in the same component following one or more links.

22\) The number of possible network structures among $N$ heterogeneous firms is $2^{\frac{N(N-1)}{2}}$. 

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To maintain tractability, in what follows I restrict the number of links a firm can form.\footnote{A similar assumption on maximum number of links is made in Allen, Babus, and Carletti (2012).}

**Assumption 3** *(Chain Networks)* each firm can form a maximum of two links, i.e. $\bar{l} = \frac{1}{2}$.

Since a distressed firm can only have one link, the possible network topology is an arbitrary collection of paths,\footnote{A path in a network is a sequence of firms and links that start with firm $i$ and end with another firm $j$.} or chain networks. The number of firms in my model can be interpreted as the largest diameter in an otherwise general network, for instance a core-periphery structure as in Farboodi (2014).

### 2.5 The Equilibrium

After observing the distress vector $z$, firms simultaneously choose linkage decision $l$ and price offerings $p$ to maximize their firm values $V(z, L, p)$. Next I formally define the equilibrium by extending the notion of *pairwise stability* in Jackson and Wolinsky (1996). I embed bilateral prices along the lines of transfer payments in Bloch and Jackson (2007).

**Definition 2** The equilibrium of a network formed by bilateral forward swap contracts is characterized by the linkage structure $L^e$ and the set of bilateral prices $p^e$, such that

- **Optimality:** each firm $i$ takes as given other firms’ strategies $(l_j, p_j), \forall j \neq i$, and chooses its own strategy $(l_i, p_i)$ to optimize its firm value, i.e.

\[
V_i(z, L^e, p^e) = \max_{(l_i, p_i) \in \{0, \bar{l}\}, p_j, j \neq i} V_i(z, L, p),
\]

subject to (2), (4), and constraints (5) - (8).

- **Pairwise stability:** denote $L^e_{-\{ij\}}$ as the matrix $L^e$ by deleting $L^e_{ij}$, and $p^e_{-\{ij,ji\}}$ as the matrix $p^e$ by deleting $p^e_{ij}$ and $p^e_{ji}$. Then $\forall L^e_{ij} > 0$ and $\forall (\hat{p}_{ij}, \hat{p}_{ji}) \neq (p^e_{ij}, p^e_{ji})$,

\[
V_i(z, L^e, p^e) \geq V_i(z, L^e_{-\{ij\}}, L_{ij} = 0, p^e_{-\{ij,ji\}}, \hat{p}_{ij}, \hat{p}_{ji}),
\]

and $\forall L^e_{ij} = 0$ and $\forall (\hat{p}_{ij}, \hat{p}_{ji}) \neq (p^e_{ij}, p^e_{ji})$

\[
V_i\left(z, L^e_{-\{ij\}}, L_{ij} = \bar{l}, p^e_{-\{ij,ji\}}, \hat{p}_{ij}, \hat{p}_{ji}\right) > V_i(z, L^e, p^e),
\]

\[
\Rightarrow V_j(z, L^e_{-\{ij\}}, L_{ij} = \bar{l}, p^e_{-\{ij,ji\}}, \hat{p}_{ij}, \hat{p}_{ji}) < V_j(z, L^e, p^e).
\]
• Feasibility:

\[
L \times 1_{N \times 1} = L^\top \times 1_{N \times 1} = 1_{N \times 1}.
\] (14)

The pairwise stability concept states that two firms connect only if both decide to connect and prefer no other bilateral prices; two firms do not connect only if, for all possible bilateral prices, at least one firm has no incentive to connect. Pairwise stability naturally applies to this setting as the goal here is to understand which networks are likely to arise and remain stable. Moreover, it eliminates the multiplicity of equilibrium networks due to coordination failures under the standard concept of Nash equilibrium.

2.6 Discussions

Synergy from links The two types of links, risk-sharing links and distress links, generate different sources of synergy. A risk-sharing link always generates a positive surplus by reducing the volatility of liquid assets. For example, a link between two \textit{ex ante} identical non-distressed firms reduces the liquidation probability of each firm.\textsuperscript{25} In comparison, a distress link has an extra source of synergy from the distress heterogeneity. For example, let \(\nu_1 = 1.5\), \(\nu_2 = 0.8\). In the forward swap contracts, firm 1 has a claim of \(\frac{1}{2}\bar{a}_2\), and vice versa. Even when \(\sigma = 0\), there is gain as the liquidation of firm 2 can be avoided. The surplus from the reduction of total liquidation costs of firms \(i\) and \(j\) is shown to increase with their distress dispersion \(|z_i - z_j|\).\textsuperscript{26} When \(z_j < -1\), the surplus is positive only if \(z_i > 0 - z_j > 1\); thus, only firms that are liquid enough are able to profit from such a link.

Distress link as an acquisition relation The price offering rule and the fact that a distressed firm only has one link jointly imply that a distress link establishes an \textit{equity ownership} relation between the liquid and the distressed firm, which can be thought of as an \textit{acquisition}. The reason is that the distressed firm \(i\) does not enter other links,\textsuperscript{27} so \(\{p_{ij}, p_{ji}\}\) satisfy \(p_{ij} = p_{jj}\) and \(V_j(z, L, p_{ij} - p_{ji}) = p_{jj}\). \(V_j\) being fixed implies that the liquid firm \(i\) is claiming the entire

\textsuperscript{25}The total expected liquidation costs of two stand alone firms are \(2 \Pr(\bar{a}_i < 1)c = 2\Phi(-z_i)c\). That of two connected firms are \(2\Phi(-2z_i)c\). The total surplus equals \(2(\Phi(-z_i) - \Phi(-2z_i)c > 0\).

\textsuperscript{26}The synergy equals the reduction of liquidation costs of the two firms \(\Phi(-z_i)c + \Phi(-z_j)c - 2\Phi(-z_i - z_j)c\). The derivative of synergy with respect to \(|z_i - z_j|\), holding the sum \(|z_i + z_j|\) fixed, is positive.

\textsuperscript{27}The offered price premium \(p_{ij} - p_{jj}\) endogenously responds to the outside option of firm \(j\) which is in turn determined by the linkage structure \(L\).
surplus value from the bilateral link. In other words, firm $i$ maximizing $V_i$ is equivalent to maximizing $V_i + V_j$, which resembles an acquisition relation.

**Algorithm for linkage formation** There are multiple ways to determine which network emerges given a set of contingent transfer payments (prices). I illustrate the following one. Under rational expectations, firms form a common belief about the equilibrium linkage structure $L^b$. Based on this belief, firms simultaneously submit strategies $l_i(L^b)$ and $p_i = \left( p_{ij}(z, L^b_i, L^b_j) \right)_{j \neq i}$. Given the strategies, the realized equilibrium network is consistent with the common belief, i.e. $L^e = L^b$. An alternative guess-and-verify approach is described in Bloch and Jackson (2007).

**Existence of equilibrium** The existence of the pairwise stable equilibrium in Definition 2 follows from a generalization of Goyal (2009) Proposition 7.1 “For any value function and any allocation function, there exists at least one pairwise stable network or a closed cycle of networks.” I refer the reader to Goyal (2009) for more discussion.

**Payment seniority** The liquid asset obtained from the forward contracts is used to pay debt at date 2, whereas the payments for the forward swap contracts are paid in full at date 3 using yields from the long-term assets. This specification assumes that short-term creditors have seniority over OTC derivative counterparties. The motivation is that derivatives seniority creates an inefficiency in risk-sharing, similar to that illustrated in Bolton and Oehmke (2014). Following the example above, let instead $\nu_1 = 1.2, \nu_2 = 0.8$. Suppose further that $\varepsilon_1 = \varepsilon_2 = 0$, so there are two units of liquid asset in total. Firm 2 has to incur liquidation cost at date 2 whenever it pays a positive net payment (firm 2 is relatively more distressed) to firm 1. In comparison, when net payment is paid at date 3, both firms avoid liquidation. As such, deferring the payments to the final date helps to isolate the network externality mechanism in my model from other potential inefficiencies associated with the derivatives payments.

3 Network Inefficiency

In this section, I examine the efficiency of the equilibrium network relative to a benchmark that minimizes total liquidation costs. Results show that the equilibrium network is inefficient.
Figure 1. Optimal Network. This figure shows the optimal risk-sharing network characterized in Proposition 1 for $N = 4$ and $N = 5$. The horizontal and vertical axes represent the mean and dispersion of firm distress statistic $z$. In the white region, all firms are linked in one component. In the dark region ($\bar{z} \in [\bar{z}_2, \bar{z}_1], \delta > \delta_1(\bar{z})$), firm $N$ is isolated.

when the dispersion of financial distress levels is high: there are more distress links and fewer risk-sharing links. Lastly, I discuss the key friction that drives the network inefficiency.

3.1 Optimal Network

Under the model specifications for links and the asset swap process, the social planner chooses the optimal linkage structure that minimizes total liquidation costs (maximizes total bank values).

Definition 3 The optimal network $L^*$ minimizes total expected liquidation costs, i.e.

$$L^* = \arg \min_{L_{ij} \in \{0,1\}} \sum_i \Pr(\tilde{h}_i < 1) c,$$

subject to the conditions of two-sided links $L_{ij} = L_{ji}$, iterative procedure (8), and feasibility (14).

Based on this definition, next I solve Problem (P1) and characterize the properties of $L^*$.

Proposition 1 (Optimal Network) $\exists \bar{z}_1, \bar{z}_2, \bar{z}_1 > \bar{z}_2 \geq 0, \exists$ cutoff function $\delta_1(\bar{z}) > 0$ such that

- for $\bar{z} \geq \bar{z}_1, \delta \geq 0$ or $\bar{z} \in [\bar{z}_2, \bar{z}_1], \delta \in [0, \delta_1(\bar{z})]$, all firms are connected in one component; formally, either $L^*_{ij} > 0$ or there exists a path between $i$ and $j$, i.e. $L^*_{ik_1}, \ldots, L^*_{km,j} > 0$;
- for $\bar{z} \in [\bar{z}_2, \bar{z}_1], \delta > \delta_1(\bar{z})$, the distressed firm $N$ is isolated ($L^*_{NN} = 1$), whereas all other firms are connected in one component.
Proposition 1 states that the optimal network can be characterized by the two moments of the distress distribution, \{\bar{z}, \delta\}. All firms fully diversify by connecting in one component in an economy with high enough average \bar{z} (low average distress), or with low \bar{z} and low enough distress dispersion \delta. In comparison, when distress dispersion \delta is high and \bar{z} is not sufficiently high, the most distressed firm \text{N} should be isolated, whereas all other firms are connected in one component. These patterns are shown in Figure 1 for \text{N} = 4 and \text{N} = 5.\footnote{The cutoff value \(\bar{z}_2\) is zero for \(\text{N} = 4\), and is positive for \(\text{N} \geq 5\). For \(\bar{z} < \bar{z}_2\), there are regions when \(L^*\) isolates more than one firm. For instance, in Panel B Figure 1, both firms 4 and 5 are isolated in the hump-shaped region in the lower left corner. As \delta increases further, \(L^*\) switches from isolating two firms to one firm. This is because the total expected liquidity of firms 1 to \(\text{N} - 1\) increases with \delta which mechanically results from the assumption of symmetric cross-sectional distribution in Equation 1. Similar patterns display for \(\text{N} > 5\).}

The intuition for Proposition 1 is the trade-off between diversification and risks of contagion.\footnote{The trade-off between risk-sharing and contagion is in line with Cabrales, Gottardi, and Vega-Redondo (2014), who find that, when shock distribution has thin tails, firms should be connected in one component, whereas when shock distribution has fat tails, maximum segmentation into small components is optimal.} In an economy with high dispersion \(\delta\) and low average \(\bar{z}\), firm \text{N} is heavily distressed from Equation (1). The contamination cost of linking firm \text{N} with all other firms dominates the risk-sharing benefit, which rationalizes isolating it.

The model specifications on links and asset swaps do not deviate the optimal network from the best possible risk-sharing outcome. In Appendix A.2, I show that under the iterative swap procedure, the asset holdings implied by the optimal network, \(\bar{h}^* = (L^*)^\infty \bar{a}\), are equivalent to the optimal allocations when the social planner directly chooses asset holdings for each firm. Hence, total liquidation costs achieve the minimum as long as the network is optimal.

### 3.2 Excess Distress Link

The question I address next is whether the optimal network can be decentralized in the network formation, and if not, in which ways the equilibrium network is inefficient.

**Proposition 2 (Excess Distress Link)** For \(\text{N} = 4\), all firms are connected in one chain in equilibrium and the distressed firm 4 is linked with the most liquid firm 1; formally, \(\forall i, \bar{z}, \delta, L_{ii}^e < 1 \text{ and } \forall j \neq i, \text{ either } L_{ij}^e > 0 \text{ or there exists a path between } i \text{ and } j, \text{ i.e. } L_{ik_1}^e ... L_{km}^e L_{mj}^e > 0\).

Proposition 2 states that for all parameter values, all firms are connected in one component at equilibrium including the most distressed firm via a distress link. Comparing Propositions 1 and 2, when the average \(\bar{z}\) is low and dispersion \(\delta\) is high, the optimal network has no distress link.
Figure 2. Equilibrium Network \((N = 4)\). This figure shows the equilibrium four-firm chain network. Firms are ranked by the level of distress, and firm 4 is distressed. A solid line represents a link between two firms. A $ arrow indicates the direction of net payment transfers via bilateral prices.

\((L^*_{NN} = 1)\); however, the equilibrium network features over-connection, \(\sum_{i \neq N} L^e_{iN} - \sum_{i \neq N} L^*_iN > 0\).

The excess distress link implies that equilibrium network is inefficient.

Figure 2 illustrates the intuition. Under reservation prices \(p_{ij} = p_{jj}\), firm 1 deviates to link with firm 4 to obtain a large profit (as \(p_{14} = p_{44}\)). Firm 2 has incentive to sever the 1 – 2 link as the cost of indirectly holding a faction of \(\tilde{a}_4\) is too high. In order to keep firm 2 staying connected, firm 1 offers a premium price \(p_{12}\) by sharing part of the profit from \(L_{14}\). This premium price matches the value of firm 2 to the same value that firm 2 gets when it withdraws. This way, there is over-connection at equilibrium: the distressed firm 4 should have been isolated but is linked into the network. Firm 2 cannot afford to pay a premium price high enough to prevent 1 from connecting with 4. This is because the benefit of isolating \(\tilde{a}_4\) is shared between firms 2 and 3, and so firm 2 would be worse-off paying the required premium fully on its own.

### 3.3 Risk Sharing Loss

As the chain network gets longer, the excess distress link can crowd out valuable risk-sharing links, thus giving rise to an additional channel of inefficiency from the loss of risk-sharing.

**Proposition 3** (Risk Sharing Loss) For \(N \geq 5\), 3 cutoff function \(\delta_2(\tilde{z})\) such that when \(\tilde{z} \in [\tilde{z}_2, \tilde{z}_1]\) and \(\delta > \delta_1(\tilde{z})\), there is excess distress link, \(\sum_{i \neq N} L^e_{iN} - \sum_{i \neq N} L^*_iN > 0\). In particular,

- when \(\delta \in [\delta_1(\tilde{z}), \delta_2(\tilde{z})]\), all firms are connected in one component, so there is over-connection due to the distress link;

- when \(\delta > \max \{\delta_1(\tilde{z}), \delta_2(\tilde{z})\}\), the non-distressed firms are not connected in one component: the network has inefficient composition due to both excess distress link and insufficient risk-sharing.
**Figure 3. Equilibrium Network** ($N = 5$). This figure shows the equilibrium five-firm network. The horizontal and vertical axes represent the mean and dispersion of firm distress statistic $z$. In colored regions, the optimal network isolates firm 5. Blue (lighter) region denotes over-connection, and orange (darker) region denotes inefficient network composition.

**Figure 4. Inefficient Network Composition** ($N = 5$). This figure shows the equilibrium connection structure $\{1 - 5, 2 - 3, 4\}$ with inefficient network composition that features both over- and under-connections.

Proposition 3 formalizes two channels of inefficiency: one from the excess distress link (over-connection), and the other from risk-sharing loss (under-connection). When the average $\bar{z}$ is low and dispersion $\delta$ is high, the distressed firm $N$, which should be isolated, is linked by firm 1 at equilibrium, generating the excess distress link. This result occurs in the colored regions in Figure 3 where $\bar{z} \in [\bar{z}_2, \bar{z}_1]$ and $\delta > \delta_1(\bar{z})$. Specifically, if the value of dispersion is in a middle range ($\delta \in [\delta_1(\bar{z}), \delta_2(\bar{z})]$), all firms are linked in one component, so inefficiency only results from over-connection. When the dispersion increases further ($\delta > \max\{\delta_1(\bar{z}), \delta_2(\bar{z})\}$), some risk-sharing link severs: a non-distressed firm becomes isolated or the non-distressed firms separate into multiple components. The externality from the distress link crowds out potential gains from risk-sharing. In this case, the inefficient network features inefficient composition featuring
Figure 5. Complete Contingent Contracts ($N = 4$). This figure shows the optimal network is decentralized at equilibrium under complete contingent contracts. Firms are ranked by the level of distress, and firm 4 is distressed. A solid line represents a link between two firms. A $\$\$\$\$ arrow indicates the direction of net payment transfers via bilateral prices.

over- and under-connections simultaneously.

Take a $N = 5$ chain network as an example. Without loss of generality, firms start from the chain $1 \rightarrow 2 \rightarrow 3 \rightarrow 4 \rightarrow 5$. As dispersion $\delta$ increases, firm 5 becomes distressed. The $4 \rightarrow 5$ link terminates and the distress link $1 \rightarrow 5$ forms: equilibrium network $1 \rightarrow 2 \rightarrow 3 \rightarrow 4 \rightarrow 5$ generates over-connection. As $\delta$ rises further, firm 2 is worse off staying in the network: $1 \rightarrow 2$ link severs and $2 \rightarrow 4$ link forms, as shown in Figure 4. Notice that various initial sequences in the stable risk-sharing chain at $\delta = 0$ imply different outside options and deviation incentives for each firm. Different initial sequences therefore lead to different equilibria, all of which share the same inefficiency feature. Detailed analysis is included in Appendix A.3.

3.4 The Key Friction

The inefficiency is caused by network externalities. Due to local contingency specified in Assumption 1, the liquid firm fails to internalize the negative externalities to its direct and indirect counterparties. When Assumption 1 is relaxed, bilateral prices $p_{ij}(z,L)$ can induce the efficient network, which indicates that the incomplete contingency on the network structure is the mere underlying friction.

Recall the $N = 4$ example. When $\delta$ is high, linking with the distressed firm 4 by 1 imposes an externality to both 2 and 3. To prevent this distress link, firms 2 and 3 need to jointly offer incentives to 1. In Appendix A.4, I formally elaborate that there exist unique premium prices $p^{*}_{21}$ and $p^{*}_{32}$ such that $L^{*}_{14} = 0$ if and only if $L^{*}_{14} = 0$. In particular, $p^{*}_{32}$ is a function of $L_{14}$ and firm 3 pays a premium price when $L_{14} = 0$. Put differently, the price offered by firm 3 depends not only on the links of 2 and 3, but also on the links of the counterparty’s counterparty (see Figure 5).
4 The Distress Dispersion

In this section, I investigate factors that indicate the level of network inefficiency. While prior literature has largely focused on the first moment of financial distress, I show that heterogeneity in firm distress measured by the dispersion $\delta$ is a critical indicator for inefficiency. Both inefficiency indicators, value loss and systemic risk, increase with dispersion $\delta$. Using comparative statics, I explain this positive relation by associating the network inefficiency to changes in the network composition.

4.1 Measures of Inefficiency and Dispersion

In the model, I measure network inefficiency by value loss and systemic risk. Define value loss, $\Delta V$, as the difference in total expected firm values between the optimal and the equilibrium networks. Then let $\Delta V\%$ be the percentage value loss, which is simply the percentage of value loss over total optimal firm values.

$$
\Delta V = \sum_{i=1}^{N} V_i(z, L^*, p^*) - \sum_{i=1}^{N} V_i(z, L^e, p^e); \quad \Delta V\% = \frac{\Delta V}{\sum_{i=1}^{N} V_i(z, L^*, p^*)}.
$$

(15)

Under the feasibility condition of asset swaps in Equation (14), value loss equals the increment of total liquidation costs. Next, I characterize the properties of value loss as a function of the two moments of firm distress distribution, $(\bar{z}, \delta)$.

**Proposition 4** (Value Loss) Value loss decreases with average $\bar{z}$ and increases with dispersion $\delta$. It increases with $\delta$ faster when $\bar{z}$ is lower. Formally, $\frac{\partial \Delta V}{\partial \bar{z}} \leq 0$, $\frac{\partial \Delta V}{\partial \delta} \geq 0$, and $\frac{\partial^2 \Delta V}{\partial \bar{z} \partial \delta} \leq 0$.

From Proposition 4, value loss is bigger when the average distress is higher or when the dispersion is higher. In such scenarios, firm $N$ is so distressed that linking it with other firms generates large contagion risk. Consequently, the cost from such a distress link causes higher loss in total firm values.

Next I explore an alternative measure for inefficiency: systemic risk denoted as $Pr_{sys}^L$. It is defined as the probability that all firms liquidate at the same time. In a network where all firms are linked in one component, systemic risk equals the liquidation probability of one firm because all firms hold exactly the same diversified asset, i.e.

$$
Pr_{sys}^{all\ connect} = Pr \left( \frac{1}{N} \sum_{i=1}^{N} \tilde{a}_i < 1 \right).
$$

(16)
Figure 6. Excess Systemic Risk. This figure plots the excess systemic risk against average $\bar{z}$ and dispersion $\delta$ for the equilibrium four-firm chain network.

In a network that isolates the distressed firm, systemic risk is the probability that the isolated firm liquidates at the same time when all non-distressed firms in one connected component liquidate,

$$Pr_{\text{isolate}} N = Pr \left( \frac{1}{N-1} \sum_{i=1}^{N-1} \tilde{a}_i < 1 \right) \times Pr (\tilde{a}_N < 1).$$

(17)

Define excess systemic risk, $\Delta Pr_{\text{sys}} = Pr_{\text{sys}}^{L^*} - Pr_{\text{sys}}^{L^r}$, i.e. the difference between systemic risk at the equilibrium network compared to the optimal network. In the example of $N = 4$, the excess systemic risk is positive whenever the network is inefficient. That is, $\Delta Pr_{\text{sys}}(N = 4) > 0$ in the inefficient region ($\bar{z} \in [\bar{z}_2, \bar{z}_1], \delta > \delta_1(\bar{z})$).³⁰

Figure 6 plots excess systemic risk as a function of the mean (Panel A) and dispersion of $z$ (Panel B). Excess systemic risk is positive when the average distress is sufficiently high and firm distress is dispersed. $\Delta Pr_{\text{sys}}$ decreases with $\bar{z}$; and as long as the dispersion $\delta$ is high enough, it increases with $\delta$ at a steeper rate when $\bar{z}$ is lower. The similarity of these patterns with Proposition 4 suggests that excess systemic risk serves as an alternative measure for inefficiency.

4.2 Comparative Statics: dispersion, inefficiency, and network composition

The above analysis shows that firm distress dispersion $\delta$ is a key indicator for both measures of inefficiency. To inspect the mechanism, I analyze how the equilibrium network responds to changes in $\bar{z}$ and $\delta$, relative to the optimal network. Especially, I look at the two inefficiency measures, $\Delta V$ and $\Delta Pr_{\text{sys}}$, together with changes in the network composition in terms of distress.

³⁰For example, when $\bar{z} = 0.2$ and $\delta = 1.5$, $\Delta Pr_{\text{sys}} = 0.34 - 0.05 = 0.29$. 
Figure 7. Increase in Average Distress under High Distress Dispersion. This figure shows the properties of the five-firm chain network with high $\delta$ when we lower $\bar{z}$. The horizontal axis $\Delta \bar{z}$ is the reduction in $\bar{z}$. I plot the values in the equilibrium network (solid) and the optimal network (dashed).

In the first comparative statics, I lower the level of $\bar{z}$ in two cases when $\delta$ takes a low and a high value. When firms are similar in financial distress ($\delta$ is low), all firms linking in a single component is optimal and pairwise stable. As we lower $\bar{z}$, the optimal network remains unchanged and is also stable. Consequently, both $\Delta V$ and $\Delta Pr_{sys}$ equal zero.

Results are different when firms are dispersed in financial distress ($\delta$ is high): a decrease in $\bar{z}$ affects the optimal and the equilibrium network differently. Figure 7 plots the value loss (Panel A), systemic risk (Panel B), distress links (Panel C), and risk-sharing links (Panel D) as functions of the reduction in $\bar{z}$ in a five-firm network, starting from $\delta = 1$ and $\bar{z} = 0.5$.\footnote{I consider a chain network $1 - 2 - 3 - 4 - 5$ of which the optimal and equilibrium networks are analyzed in Subsection 3.3 and in Figure 3. In particular, I set $A = 4$, $c = 2$, so that when $\delta = 1$ and $\bar{z} = 0.5$, the average liquidation cost amounts to 8% of total firm value.} As $\bar{z}$
Figure 8. Increase in Dispersion. This figure shows the properties of the five-firm chain network when we raise dispersion $\delta$ while adjusting $\bar{z}$ so total firm values at $L^*$ remain constant. I plot the values in the equilibrium network (solid) and the optimal network (dashed).

reduces, both value loss $\Delta V$ and excess systemic risk $\Delta Pr_{sys}$ (the difference of the solid and the dashed curves in Panel B) rise. Corresponding to where the inefficiency occurs, Panels C and D show that the equilibrium network has one extra distress link between 1 and 5, and one fewer risk-sharing link between 1 and 2. This exercise has two implications. First, comparing the two cases when $\delta$ takes a low and a high value, $\delta$ only matters to inefficiency when it is sufficiently large. Second, the observed positive relation of inefficiency and $\delta$ is associated with changes in the network composition.

In the second comparative statics, I study how the equilibrium network changes with dispersion. However, when firms form links optimally, increasing dispersion alone increases total firm values, as the total liquidation costs decrease monotonically.\textsuperscript{32} For this reason, in the following

\textsuperscript{32}When all firms are linked in a single component, total liquidation costs equal $N\Phi \left[ \sqrt{N} (-\bar{z}) \right] c$, independent of $\delta$. When the most distressed firm is optimally isolated, total liquidation costs, $(N-1)\Phi \left[ \sqrt{N-1} (-\bar{z} - \frac{1}{2}\delta) \right] +$
exercise, I increase $\delta$ while also adjusting $\bar{z}$ such that the total firm values in the optimal network remains constant. This allows me to conduct a “fair” comparison across states when identical firm values can possibly be achieved. Figure 8 plots the inefficiency measures and linkages of the same chain network as before. From Proposition 4, both measures of inefficiency (see in Panels A and B) increase with dispersion. When $\delta$ is large enough, inefficiency becomes positive and increases thereafter. Particularly, systemic risk at equilibrium increases with $\delta$, except for the drop when the 1-2 risk-sharing link severs, which reduces asset correlations.

These patterns are due to over-connection at high values of dispersion and wrong network composition when dispersion gets even higher. This can be seen by comparing the number of distress links and risk-sharing links in Panels C and D. The optimal network isolates firm 5 before it becomes distressed, so the only jump on the dashed curve in Panel C is when firm 4 falls into distress. In comparison, firms 4 and 5 are always connected at equilibrium (see the two steps in the solid curve), which results in over-connection. Shown in Panel D, when $\delta$ is high, the optimal network has one more risk-sharing link than the equilibrium network (dashed minus solid curves). The severance of the 1-2 risk-sharing link implies wrong network composition, which creates an extra channel for inefficiency.

To summarize, the above two comparative statics conclude that a decrease in $\bar{z}$ when $\delta$ is high, or an increase in $\delta$ (together with a decrease in $\bar{z}$) is associated with: (1) higher value loss and higher systemic risk, (2) more distress links, (3) fewer risk-sharing links. In both exercises, the cross-sectional distribution of firm distress has high dispersion.

5 Policy Implications on the Acquisitions of Distressed Firms

In this section, I apply the model to the case where links with distressed firms are interpreted as acquisitions. There are two reasons for this particular application. First, in the data, a major example of the links with distressed firms is acquisitions. Acharya, Shin, and Yorulmazer (2010) and Almeida, Campello, and Hackbarth (2011) provide evidence that liquid firms acquire distressed firms for potential gains from asset sales or advantageous bargaining position. Second, compared with OTC derivative contracts that are challenging to supervise, acquisitions $\Phi [-\bar{z} - \frac{1-N}{N} \delta]$, decrease monotonically with $\delta$. With no linkages, however, liquidation costs increase monotonically with dispersion $\delta$ as more firms are distressed.
in the financial sector are subject to regulatory approval, which makes it relevant for policy interventions.

Based on the model result, regulations that prevent the inefficient distress links can generate social gains. I begin by proposing one such regulation using an acquisition tax to supervise acquisitions. Then I study an extension of the model that allows for the analysis of optimal government policies both before and after the linkage formation. Results indicate that the too-connected-to-fail problem arises if the excess acquisition is not effectively prevented ex ante. In this case, liquidating the distressed firm is too costly due to spillovers to its existing counterparties. Using the extended model, I discuss, respectively, the options of government bailout, subsidized acquisition, and pushed acquisitions. I find that these are ex post optimal remedies, thereby rationalizing the government interventions observed during the crisis.

5.1 Acquisition Tax

Current authorities consider acquisition as the primary approach to resolve firm distress as it incurs the least fiscal cost. However, my results imply that acquisitions of distressed firms should rather be regulated accounting for the externalities in the financial linkage formation. If the regulators are able to provide incentives by imposing taxes, then a tax formula that varies with the distress distribution can induce the optimal level of acquisitions and restore the efficient network. Next I formally characterize the tax rate.

Proposition 5 (Acquisition Tax) In an $N$-firm chain, the optimal network can be decentralized by a tax $\tau$ imposed to firm 1 upon its acquisition of the distressed firm $N$,

$$
\tau = \left[ N \Phi \left( \sqrt{N} (-\bar{z}) \right) - (N-2) \Phi \left( \sqrt{N-1} (-\bar{z} - \frac{1}{2} \delta) \right) - \Phi(-z_1) - \Phi(-z_N) \right] c. \tag{18}
$$

where $\Phi(.)$ is the CDF of the standard normal distribution. Furthermore, $\tau > 0 \iff L_{NN}^* = 1$. $\tau$ satisfies $\frac{\partial \tau}{\partial \delta} > 0$ and $\frac{\partial \tau}{\partial \bar{z}} < 0$.

Proposition 5 states that the acquisition tax is positive if and only if the most distressed firm should be isolated in the optimal network. Moreover, the acquisition tax increases with the average and dispersion of distress. The intuition is as follows. The acquisition tax equals precisely the negative externalities to all other non-distressed firms $i = 2, \ldots, N - 1$. Hence, it exactly aligns the individual motivation with the social incentive for acquisition. Accounting
for negative spillovers, the acquisition tax is a function of the cross sectional distribution of firm distress in terms of \( \{N, \bar{z}, \delta\} \). When dispersion is higher, the negative externalities are bigger; hence, we require bigger incentive to correct for the externality. A similar argument holds for the relation with the average distress. Note that the tax is only imposed conditional on the excess acquisition. Therefore, no tax will be physically collected from the acquirers because the inefficient acquisition is effectively prevented.

The model provides a sharp theoretical guidance on how to regulate acquisitions. In particular, the novel insight of considering firm distress distribution complements the current metrics in regulatory decisions. The “financial stability” factor has been included for the first time for processing firm acquisitions by the Dodd-Frank Wall Street Reform and Consumer Protection Act in Section 604(d). This section amended Section 3(c) of the Bank Holding Company Act of 1956 and it requires the Fed to consider “the extent to which a proposed acquisition, merger or consolidation would result in greater or more concentrated risks to the stability of the United States banking or financial system.” In the orders on approving recent acquisitions, for instance Capital One’s acquisition of ING Bank, the Fed illustrates the new financial stability metrics in response to Dodd-Frank’s mandate, including size, substitutability, interconnectedness, complexity, and cross-border activity.\(^{33}\) The discussion regarding the interconnectedness factor, however, only covers the degree of interconnectedness of the resulting firm, rather than considering the entire linkage structure and possible externalities through indirect linkages.

The key issue is how to implement such acquisition tax. From Equation (18), the regulators need to account for the distribution of financial distress. One feasible approach detailed in Section 6 is to estimate quarterly Z-scores of all financial firms. Among the limitations of this measurement are the low frequency and the opacity of balance sheets. Using exclusive regulatory data, the banking supervisors can potentially achieve better estimates by using observations with higher frequency or alternative models such as CAMELS ratings.

Once the excess acquisitions are prevented, alternative resolution methods in case of failure include liquidation or the Purchase and Assumption (P&A) transactions. The Federal Deposit Insurance Corporation Improvement Act of 1991 mandates the FDIC to choose the resolution method least costly to the Deposit Insurance Fund. To comply with this mandate, the FDIC

\(^{33}\)www.federalreserve.gov/newsevents/press/orders/2012orders.htm
chose P&A transactions as the resolution method for a great majority of failing banks (about 95%). My results hence indicate that P&As are preferred to relying on private sector solutions which give rise to network externalities and the potential build-up of systemic risk.

5.2 Ex post Policies

Several acquisition cases observed during the recent financial crisis render the baseline model counterfactual, including the acquisitions of Bear Stearns, Merrill Lynch, and National City. These cases differ from the baseline setting in several dimensions: links with the target institutions were formed before the distress conditions were fully disclosed. Additionally, government interventions such as bailout or pushed/subsidized acquisitions took place. For example, counterparties did not immediately pull back from trading with Bear Stearns after the failure of its two funds in 2007. When Bear Stearns suffered severe financial distress on March 2008, the Fed provided assistance in the form of a non-recourse loan of $29 billion to JP Morgan to make the acquisition. To rationalize the observed government interventions of such kind, I next consider extensions of the baseline model, and the key deviation is that the timing of the network formation does not coincide with the observation of distress.

Suppose the linkage cannot be severed once formed at $t = 1$ after $\nu$ is learned. Further, assume that the liquid return $\tilde{a}_i$ satisfies

$$\tilde{a}_i = \nu_i + \theta_i + \sigma \varepsilon_i, \ i = 1, \ldots, N,$$

where the additional term $\theta_i$ is realized after links are formed. Hence, $\nu_i$ and $\theta_i$ jointly determine the amount of liquid value firm $i$ expects to receive. Let $\theta$ be a vector with $\theta_i = 0, \forall i = 1, \ldots, N-1$, and $\theta_N = -k\bar{z}\sigma$.\(^{35}\) Further let $\bar{z} \in [\bar{z}_2, \bar{z}_1]$ and $\delta > \delta_1(\bar{z})$ such that the distress firm $N$ should be isolated (Proposition 1). Nonetheless, in the absence of the acquisition tax, all firms are connected at equilibrium (Proposition 2). Now, assume firm $N$ receives a second bad liquidity shock $\theta_N$ with $k > N$ such that it drags down the average distress of all firms below zero. In this case, the links do not generate positive risk-sharing surplus, thus total liquidation costs are higher than without any links among firms.

\(^{34}\)For detailed institutional background on bank failures see White and Yorulmazer (2014), Granja, Matvos, and Seru (2014), and the the Guidance for Developing Effective Deposit Insurance Systems from FDIC, at http://www.fdic.gov/deposit/deposits/international/guidance/guidance/FailedResolution.pdf.

\(^{35}\)In practice, distress signals are released gradually. The negative $\theta_N$ captures persistence in liquidity conditions.
5.2.1 Government Bailout

Next I analyze conditions when government bailout is *ex post* optimal and how total costs compare to those under the *ex ante* optimal policies (imposing acquisition tax). For this purpose, let us enable the option of government bailout in the form of costly liquidity injection. Specifically, let $B\sigma$ denote the amount of government liquidity injection to the heavily distressed firm $N$. Since all firms are connected and each has the same diversified asset holdings, they share the same probability of liquidation $\Phi\left[\frac{1}{\sqrt{N}}(N\bar{z} - k\bar{z} + B)\right]$. Here, the total costs incurred include expenses both in liquidation and bailout.  

I find that positive government bailout is *ex post* optimal in an over-connected network as long as the liquidation cost is not very small. The formal analysis is provided in Appendix A.5, Proposition 7. When the liquidation cost satisfies $c > \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}$, a positive government bailout that matches at least the total expected liquid value shortfall ($B^* > (k - N)\bar{z}$) is *ex post* optimal. This lower bound of liquidation cost is smaller when the distressed firm has more counterparties or when asset volatility is lower. Now, suppose the second shock $\theta_N$ to firm $N$ is not sufficiently bad, the lower bound of liquidation cost that justifies government bailout will be higher. In other words, the worse shock the connected banking system gets, the more likely government bailout is *ex post* optimal. This relation is consistent with the empirical observation that bailout only occurs in rare occasions with severe distress.

Despite the fact that government bailout can be *ex post* optimal, it is likely to be more costly than preventing the excess acquisition *ex ante*. I show that, as long as the bailout cost is not sufficiently low, total costs from *ex post* government bailout is higher than regulating the links *ex ante* using the acquisition tax (see Proposition 8 in Appendix A.5). This result captures one critical aspect of inefficiency in the current policy making: the time-inconsistency problem.

When a liquid firm observes the distress of some institution, it acquires the distressed target while generating externalities. Precisely owing to the excess acquisition link, liquidation of the distressed firm gets too costly. In consequence, government bailout becomes *ex post* optimal.

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36 The total costs incurred equal $N\Phi\left[\frac{1}{\sqrt{N}}(N\bar{z} - k\bar{z} + B)\right] + B\sigma$.

37 Formally, if $0 \leq k \leq N$ in $\theta_N = -k\bar{z}\sigma$ instead, the average distress $\frac{1}{\sqrt{N}}(N - k)\bar{z}$ is then positive. And the lower bound for liquidation cost is higher than the case of $k > N$, i.e. $c \geq \frac{\sqrt{2\pi}e^{(N-k)^2/2N}}{\sqrt{N}} > \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}$.

38 For other discussions on the time-inconsistency issue, see Acharya and Yorulmazer (2007), Spatt (2009), Chari and Kehoe (2013), and Gimber (2013).
and \textit{ex ante} inefficient.

### 5.2.2 Government Subsidized Acquisition

Back to the Bear Stearns case, instead of injecting capital directly, the Fed provided assistance to the acquirer JP Morgan in the form of a non-recourse loan.\(^{39}\) With a slight variation, the extended framework can explain this behavior. I show that, when there exist healthier institutions currently not connected with the distressed firm, government subsidized acquisition can reduce total liquidation costs.

Consider another group of connected firms that are separate from the existing firms. Suppose there are \(N\) firms \(i = N + 1, \ldots, 2N\) with the same average \(\bar{z} > 0\) and dispersion \(\delta = 0\), such that a complete risk-sharing network optimally emerges.\(^{40}\) Let the additional signal \(\theta_i\) be \(\theta_{N+1} = \hat{k}\bar{z}\sigma\) and \(\theta_i = 0, \forall i = N + 2, \ldots, 2N\), so the \((N + 1)\)th firm gets a positive shock in the liquid return. The question I address next is whether firm \(N + 1\) has the incentive to acquire the distressed firm \(N\) after the realization of \(\theta\), and whether the \textit{ex post} acquisition is socially optimal.

The answer to this question depends on how the liquidity surplus of firm \(N + 1\) compares with the liquidity shortage of firm \(N\). In Corollary 1 of Appendix A.5, I show that the \textit{ex post} distressed acquisition is efficient and it occurs at equilibrium if and only if the average distress is above zero \((\hat{k} > k - 2N)\). However, if the adverse liquidity shock \(k\) is considerably large \((k \geq \hat{k} + 2N)\), the acquisition has negative surplus, and firm \(N + 1\) does not have incentive to acquire. In this case, subsidized acquisition in the form of liquidity injection to the acquirer is \textit{ex post} optimal as long as the liquidation cost is not very small \((c > \sqrt{\pi\sigma}\sqrt{N})\). The intuition is that risk-sharing among the two groups of firms reduces total liquidation costs only when the total expected liquidity is positive. And both acquisition subsidy and government bailout can push the average liquidity above zero. I find that the required optimal government subsidy is lower when the positive liquidity shock of the potential acquirer \((\hat{k})\) is higher. This result rationalizes the observation that the subsidized acquirers during the financial crisis, for instance JP Morgan and PNC (respectively acquirers of Stearns and National City), are relatively more liquid firms.

Comparing the two types of \textit{ex post} policy remedies, the government subsidized acquisition

\(^{39}\)On March 14, 2008, the New York Fed agreed to provide a $25 billion collateralized loan to Bear Stearns for up to 28 days, but later decided that the loan was unavailable to them.

\(^{40}\)The results are robust to \(\delta > 0\). I leave the robustness on the number of firms in the two groups to the next subsection.
generates lower total costs than government bailout, thus is always preferred. This result holds even when the acquisition alone is socially costly. Nonetheless, if the excess link with the distressed firm was prevented in the first place, liquidation would not be as expensive; hence, neither subsidized acquisition nor bailout would be necessary.

5.2.3 Government Pushed Acquisition

I have shown that when the two groups of firms have the same cardinality, the acquisition link forms at equilibrium if and only if it generates value gains. However, this “if and only if” condition does not hold when the cardinality of the two groups differs. Specifically, if the additional healthier group has fewer firms, the acquisition might not occur even if it is ex post socially valuable, which motivates direct government interventions.

The relative cardinality of the two groups determines the sign of the bilateral surplus and implies whether the ex post acquisition occurs at equilibrium or not. When the potential acquirer in the second group has more counterparties, there are more firms to share the cost of the acquisition than there are in the original distressed group to share the benefit. The bilateral surplus from the acquisition is greater than the social surplus, hence the acquisition link forms ex post whenever it is socially valuable. When the cardinality of the two groups are the same, the sign of the bilateral surplus matches that of the social surplus, and we are back to the special case in Section 5.2.2.

If instead the distressed firm \( N \) has more counterparties, the bilateral acquisition surplus is smaller than the social surplus. Especially, the bilateral surplus can be negative even when the social surplus is positive. Hence, the ex post socially valuable acquisition does not occur at equilibrium. In such circumstances, government pushed acquisition is socially value improving.

For a detailed analysis see Proposition 9, Appendix A.5.

There are many ways in which a government intervention can take place. One approach is by exerting pressure to the potential acquirers. Examples include the Fed pressuring Bank of America to acquire the distressed Merrill Lynch.\(^{41}\) The regulators can also aim to correct the sign of the bilateral surplus by subsidizing the acquirer using fund collected from the counterparties of the distressed firm.

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\(^{41}\)As discussed in Spatt (2010), “secretary of the Treasury Henry Paulson indicated to [Bank of America CEO] Lewis that banking supervisors would question his suitability to lead Bank of America if BoA backed out of the merger and then needed more federal support, while federal authorities agreed to provide ‘ring-fencing’ of difficult to value Merrill Lynch assets if Bank of America went ahead with the merger.”
the distressed firm. Alternatively, the regulators can provide a coordination device for collective decision making: let the potential acquirer and all the counterparties of the distressed firm bargain over the payments. One such example is the initiation of collective bailout of LTCM by the New York Fed in 1998.42

6 Empirical Evidence

In this section, I document evidence that the distribution of distress across financial institutions provides a novel measure for systemic risk and aggregate failures in the financial sector. I establish this result by first examining how the cross-sectional mean and dispersion of distress correlate with indicators for aggregate systemic risk, liquidation costs, distress links through acquisitions, and interbank risk-sharing. I then confirm the findings using predictive regressions.

6.1 Measurement

The sample of financial institutions I consider includes bank holding companies and all Federal Deposit Insurance Corporation (FDIC) insured commercial banks and savings institutions. The quarterly accounting data of bank holding companies for the period of 1986-2013 are taken from FR Y-9C filings provided by the Chicago Fed. The quarterly accounting data for commercial banks (Call Reports) and savings institutions (Thrift Financial Reports) are taken from the FDIC’s Statistics on Depository Institutions, available for 1976-2013. Next, I discuss the method for estimating the distress measure Z-scores and identifying the acquisitions of distressed firms.

6.1.1 Z-score

The quarterly accounting data provide the basis for measuring financial distress and identifying acquisitions of distressed institutions. I measure financial distress by estimating the Z-score, which has been widely used in the recent literature (e.g. Stiroh (2004), Boyd and De Nicolo (2005) and Laeven and Levine (2009)) as an indicator for a institution’s distance from insolvency (Roy (1952)). The Z-score is defined as the return on assets plus the capital-asset ratio divided by the standard deviation of return on assets. Simply put, it equals the number of standard deviations that an institution’s return on assets has to drop below the expected value before

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42On Sept 23 1998, the New York Fed arranged a meeting for a group of LTCM’s major creditors at one of its conference rooms. During this historic meeting, the creditors worked out a restructuring deal that recapitalized LTCM and avoided its bankruptcy.
Figure 9. Log Z-score Moments across Financial Institutions. This figure plots the quarterly time series of dispersion, mean, and the 10-90 percentile range of log Z-score across all financial institutions over the period of 1978-2013. The series are normalized such that both the dispersion and the mean are centered around one. Shaded bars indicate NBER recessions.

The Z-score combines accounting measures of profitability, leverage and volatility. In particular, it is estimated according to the formula

\[
Z\text{-score}_{i,t} = \frac{1}{T} \sum_{\tau = 0}^{T-1} ROA_{i,t-\tau} + \frac{1}{T} \sum_{\tau = 0}^{T-1} CAR_{i,t-\tau},
\]

(20)

where \(ROA_{i,t}\) and \(CAR_{i,t}\) are respectively the return on assets (net income over total assets) and capital asset ratio (total equity capital over total assets) for firm \(i\) in quarter \(t\). In my analysis, the Z-score is computed considering a rolling window of eight observations, i.e. \(T = 8\). The estimated Z-score is highly skewed; hence, I follow Laeven and Levine (2009) and Houston, Lin, Lin, and Ma (2010) and adopt the natural logarithm of the Z-score as the distress measure.

The time series of the mean and dispersion of log Z-score are estimated by taking the average and standard deviation across all financial firms in each quarter. Figure 9 plots the quarterly series of dispersion, mean, and the 10-90 percentile range of log Z-score over the period of 1978-2013. For the purpose of visualization, the series are normalized such that both the dispersion and the mean are centered around one. The shaded bars indicate NBER recession dates.

From Figure 9, we can make the following observations. First, relative to the cross-sectional
mean, the dispersion of log Z-score displays a fair amount of variation and has an increasing overall trend. Second, the dispersion series demonstrates a countercyclical pattern: it increases during the Savings and Loan crisis, the Dot-com crash and the recession afterwards, as well as during the 2007-2009 financial crisis. Based on the comparative statics in Section 4.2, precisely during the crises spell, network inefficiency is more pronounced, which potentially aggravates the crises and increases systemic risk. Finally, the dispersion series appears to lead recessions. Take the most recent crisis for instance, the dispersion starts to increase since 2005, and by the time financial firms enter the crisis in the 3rd quarter of 2007, they already show significant dispersion in financial distress. These features combined suggest that the time series of dispersion can potentially signal economic changes and systemic risk, which I will test at the end of this section.

While the Z-score provides a quantitative measure for distress, it is worth noting a few limitations. First, the quarterly accounting data are an endogenous outcome of certain degrees of risk diversification, thus are not exogenous to firms as assumed in my model. Nonetheless, the Z-score gives the best available proxy for the distress shock in the static framework because it is estimated using past data, which are taken as given by firms to make decisions onwards. The Z-score indicates firm stability well also because, as shown by Acharya, Shin, and Yorulmazer (2010), initially liquid firms tend to hoard liquidity or deleverage for potential gains from asset sales, whereas risk management tools for an initially distressed firm are limited. Hence, the ranks of the estimated Z-score across firms can reflect the ranks of initial distress. The second limitation pertains to the estimation of Z-score using accounting data. It omits off-balance sheet activities, and thus possibly gives a biased measure of firm risk. However, off-balance sheet usages are only relevant for a few institutions, hence do not necessarily affect the entire distribution.

6.1.2 Acquisitions of Distressed Firms

Based on the above measure, an acquisition of a distressed firm occurs when the target has a low Z-score. This enables us to proxy for the acquisition links with the distressed firms in the model. The acquisition transactions are taken from the Chicago Fed Mergers and Acquisitions dataset. The dataset records all the acquisition transactions of banks and bank holding companies since
1976, keeping track of both the target and acquirer entities at the merger completion date. I drop the observations that are failures or restructurings. I then match the dataset with quarterly accounting data using RSSD ID of the target firm two quarters ahead. Around 86% (17,930) of the observations are matched. Out of the matched sample, I identify a distressed acquisition if the target firm reports a negative net income two quarters prior to the acquisition completion date, or if the target firm has a log Z-score of below 2.35 (two standard deviations below the sample mean) at least once, two to four quarters before the acquisition completes.

Using this strategy, around 20% (3,153) of the matched sample acquisitions are classified as distressed acquisitions, whereas the rest mostly took place during the merger wave in the 2000s after the Gramm-Leach-Bliley Act, which enabled mergers among investment banks, commercial banks, and insurance companies. Among the identified distressed acquisitions, some notable examples include Countrywide by Bank of America, Riggs and Sterling by PNC, and Wachovia by Wells Fargo.

Figure 10(a) plots the quarterly percentage of distressed acquisitions over total number of acquisitions. This figure plots the quarterly (asset-weighted) distressed acquisition rate for 1978-2013 (left) and compares the distressed acquisition rate to the total acquisition rate (right). Shaded bars indicate NBER recessions.

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43 Failures refer to transactions with Termination Reason Code = 5. Restructurings occur when the target entities and the acquirer entities have exactly the same entity name but different Federal Reserve RSSD IDs.

44 To match as many entities as possible, in this step, I include the FR Y-9LP and FR Y-9SP fillings for bank holding companies. However, since these non-consolidated parent banks only report semiannually, I do not include them when computing the Z-score distributions. I match the quarterly accounting dataset two quarters ahead because the merger date in Chicago Fed M&A dataset represents the completion date and is usually later than the last quarter when the non-survivor firm files quarterly report.
Table 1. Distressed Acquisition Likelihood and Log Z-score

<table>
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<tr>
<th></th>
<th>Pr(Completing an Acquisition of a Distressed Firm)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
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<tr>
<td>Log Z-score</td>
<td>0.153*</td>
</tr>
<tr>
<td></td>
<td>[0.070]</td>
</tr>
<tr>
<td>Firm Controls</td>
<td>yes</td>
</tr>
<tr>
<td>Year Fixed-Effects</td>
<td>yes</td>
</tr>
<tr>
<td>2006-2013</td>
<td>yes</td>
</tr>
<tr>
<td>Observations</td>
<td>57,035</td>
</tr>
<tr>
<td>Firm Fixed-Effects</td>
<td>yes</td>
</tr>
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</table>

Notes: This table reports the results from a fixed-effects logit regression. The sample includes commercial banks, savings institutions and bank holding companies. The dependent variable Pr(Completing an Acquisition of a Distressed Firm) takes the value of one if institution \( i \) completes an acquisition of a distressed firm at time \( t + 4 \), and zero otherwise. Firm controls include quarterly CAR, ROA, and asset size. Regression coefficients are reported with standard errors in the square bracket. *, **, *** denote statistical significance at the 5%, 1%, and 0.1% level.

financial institutions as well as the distressed acquisition rate weighted by the asset size of the targets. From the plots, the distressed acquisition rates are countercyclical. Two periods with clustered acquisitions are the Savings and Loan crisis and the 2007-2009 financial crisis. The asset-weighted acquisition rate displays significant spikes (some spikes reach as high as 3%, while the plots are trimmed at 2.5%).\(^{45}\) Panel 10(b) compares the distressed acquisition rate to the total acquisition rate. The insignificant comovement between the two curves shows that variations in distressed acquisitions are unlikely driven by merger waves.

6.1.3 Model Assumptions on Distressed Acquisitions

To confirm the assumption made in the model that more liquid firms acquire the distressed firms, I match the quarterly firm-level data with the acquisition dataset using the acquirer entities and acquisition completion dates, and perform fixed-effects logit regressions. The dependent variable is a dummy indicating whether a firm conducts a distressed acquisition at a certain quarter. I assume that an acquisition takes on average four quarters to complete, so it starts four quarters prior to the merger completion date recorded in the Chicago Fed dataset. The independent variable of interest is the firm’s estimated log Z-score. Results reported in Table 1 confirm that a firm with higher log Z-score has a higher likelihood of acquiring a distressed firm. For a one-

\(^{45}\)The spikes include one in the 2nd quarter of 1992 due to the acquisition of Security Pacific, one in 2007-2008 mostly due to the acquisitions of Lasalle bank (10/01/2007), Countrywide (01/11/2008), National City (10/24/2008), and Wachovia (12/31/2008).
Figure 11. *Log Z-score and Asset Size: Acquirer - Target.* This figure plots the distribution of log Z-score and log asset size of the acquirer-target wedge for the identified 3,153 distressed acquisitions in 1983-2013. Shaded bars indicate NBER recessions.

A standard-deviation increase in log Z-score (.58), the log odds ratio of a distressed acquisition increases by 0.09 (=0.153 × 0.58). The economic and statistical significance of the coefficient is robust to including firm-level controls, year fixed effects, and only considering the post-2006 period.

Among the identified 3,153 distressed acquisitions, a clear pattern emerges among the acquirer-target pairs: the acquirer has higher Z-score and bigger asset size relative to the target. The results are depicted in Figure 11. The plots show the distributions of the acquirer-minus-target log Z-score (Panel 11(a)) and log asset size (Panel 11(b)). Both distributions are significantly above zero, implying that more stable firms acquire smaller and distressed targets.

In the theoretical analysis, a link with the distressed firm is modeled as a bilateral forward swap contract, which increases the financial distress of the acquirer and thus negatively affects its Z-score. To confirm this assumption, I perform fixed-effects regressions of growth rate in log Z-score on target log Z-score, and the dummy variables representing acquisition and distressed acquisition, controlling for firm-level characteristics. The regression results summarized in Table 2 show strong support for the model assumption. The estimates suggest that the effect of the log Z-score of the targets on the growth rate of Z-score of the acquirers is positive and significant. The economic magnitude of the effect is sizable: a one-standard-deviation decrease in target log Z-score decreases future log Z-score of the acquirer by 0.16, more than four times the magnitude.
Table 2. Effect of Target log Z-score on Acquirers’ Future Z-score

<table>
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<th>(2)</th>
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<th>(4)</th>
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<td>0.310***</td>
<td>0.326*</td>
<td>0.291**</td>
<td>0.250***</td>
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<td></td>
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<td>[0.064]</td>
<td>[0.130]</td>
<td>[0.095]</td>
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<td>0.624***</td>
<td>0.884***</td>
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<td>[0.464]</td>
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<td>Yes</td>
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<tr>
<td>Firm Fixed-Effects</td>
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<td>Year-quarter Dummy</td>
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<td></td>
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<td>Yes</td>
<td>Yes</td>
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Notes: This table reports the coefficients from a fixed-effects regression. The sample includes commercial banks, savings institutions and bank holding companies. The dependent variable log $z_{i,t+1} - log z_{i,t}$ is the growth rate of log Z-score for firm $i$ at quarter $t$. The target log Z-score is the level of log Z-score of the target firm at the acquisition completion date if firm $i$ has an acquisition at quarter $t$. The dummy variables take 1 (and 0 otherwise) if firm $i$ has an acquisition or a distressed acquisition at quarter $t$. Firm controls include total assets, total equity, net income, and current level log Z-score. Regression coefficients are reported with standard errors in the square bracket. *, **, *** denote statistical significance at the 5%, 1%, and 0.1% level.

6.2 Model Predictions

As shown in the comparative statics in Section 4.2, an increase in dispersion (together with a decrease in average Z-score) is associated with higher systemic risk, more liquidations, more (excess) distress links through acquisitions, and fewer risk-sharing links. Next, I illustrate that patterns in the data provide suggestive evidence for these model-predicted relations.

6.2.1 Aggregate Indicators

The goal is to provide aggregate level evidence that distress dispersion is indicative of economic activity and financial stability. To measure macroeconomic activity, I use the Chicago Fed...
Figure 12. Bank Failure Rates. This figure plots the quarterly failure rate and asset-weighted failure rate of commercial banks and savings institutions for 1978-2013. Shaded bars indicate NBER recession dates.

National Activity Index (CFNAI),\textsuperscript{46} which is adopted in Giglio, Kelly, and Pruitt (2013) to evaluate the predictive power of various systemic risk measures. As an indicator for systemic risk, I take the Chicago Fed’s National Financial Conditions Index (NFCI).

Failures are aggregated from the FDIC Failure and Assistance Transaction Reports of all commercial banks and savings institutions in 1976-2013. I append this sample using the failures of bank holding companies, i.e. those in the Chicago Fed Mergers and Acquisitions dataset with Termination Code = 5 (failure). In total, I obtain 3,473 failures with an asset value of 1.84 trillion in 2010 dollars. I construct the quarterly failure rates (numbers of failures over the numbers of total financial institutions) as well as the failure rates weighted by the failing institution’s asset size. As depicted in Figure 12, failure rates are strongly countercyclical: the majority of bank failures took place during the Savings and Loan crisis and the 2007-2009 crisis.

Regarding the linkage composition, the model predicts that non-distressed firms that do not engage in distressed acquisitions withdraw from risk-sharing contracts as a consequence of network externalities. Direct evidence on this prediction would be obtained if full information on individual level linkage is available. Instead, I consider the lending and interbank lending behavior of small to medium-sized commercial banks as proxies for risk-sharing contracts since

\textsuperscript{46}The CFNAI is designed to gauge overall economic activity and related inflationary pressure. It includes the following subcomponents: production and income (P&I), sales, orders, and inventories (SO&I), employment, unemployment, and hours (EU&H), and personal consumption and housing (C&H).
these institutions are more likely to be the non-distressed and non-acquirer firms in the model. In particular, using data from the Fed’s H.8 release, I construct the fractions of bank credit and Fed funds and reverse Repos with banks over total assets for small to medium-sized (beyond top 25) commercial banks.

### 6.2.2 Univariate Correlations

Table 3 provides the summary statistics of the above series as well as their univariate correlation coefficients with the mean and dispersion of financials’ log Z-scores. Both the mean and dispersion series are rescaled such that the two series are centered around one. The distress dispersion displays higher variation over time and does not significantly correlate with the mean of distress, thereby confirming that dispersion provides new information not captured by the mean.

Well aligned with the theoretical findings, dispersion series correlate negatively with the economic activity index CFNAI and positively with the systemic risk index NFCI. In other words, high dispersions relate to bad economic times and low financial stability. As the model predicts, the failure rates and distressed acquisition rates are significantly higher when the dispersion is higher or when the average Z-score is lower. Additionally, the distressed acquisitions as a fraction of total acquisitions correlate even more significantly with the log Z-score moments, ruling out the possibility that the variations in distressed acquisitions are due to changes in total acquisition rates. These patterns all corroborate that high dispersion is associated with more distressed acquisitions and consequently, more failures. Last but not least, indicators for lending and interbank lending have negatively significant correlation with dispersion. Small and medium-sized commercial banks reduce interbank lending and exposures with other banks in the Fed Funds and Reverse Repos market, with significance at the 0.001 level. This finding supports that certain risk-sharing contracts terminate as dispersion increases.

### 6.2.3 Predictive Regressions

Evidence from the univariate correlations provides a strong indication that the distress dispersion comoves with aggregate indicators. However, contemporaneous correlations do not necessarily imply that the distress dispersion is able to forecast systemic risk. Hence, the next goal is to
Table 3. Summary Statistics and Univariate Correlations

<table>
<thead>
<tr>
<th></th>
<th>Mean</th>
<th>StDev</th>
<th>Sacf</th>
<th>Correlations w/ log Z-score</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean of Log Z-score</td>
<td>1.00</td>
<td>0.03</td>
<td>0.90</td>
<td></td>
</tr>
<tr>
<td>Dispersion of Log Z-score</td>
<td>1.00</td>
<td>0.22</td>
<td>0.97</td>
<td>-0.13</td>
</tr>
</tbody>
</table>

**A. Economic activity and systemic risk**

Chicago Fed National Activity Index (CFNAI) | -0.11 | 0.72 | 0.80 | -0.03 | -0.30**
National Financial Conditions Index (NFCI)  | -0.34 | 0.54 | 0.84 | -0.25** | 0.37***

**B. Bank failures**

Failure Rate (%)                       | 0.18  | 0.25  | 0.72 | -0.60*** | 0.45***
Asset-weighted Failure Rate (%)        | 0.11  | 0.25  | 0.34 | -0.38*** | 0.17*

**C. Distressed acquisitions**

Distressed Acquisition Rate (%)        | 0.21  | 0.09  | 0.64 | -0.41*** | 0.60***
Distressed over Total Acquisition Rate | 0.19  | 0.13  | 0.71 | -0.44*** | 0.68***

**D. Lending and interbank lending**

Small Comm. Bk Credit over Assets       | 0.88  | 0.02  | 0.94 | -0.26**  | -0.73***
Small Comm. Bk Fed Funds Loan over Assets| 0.02  | 0.01  | 0.85 | -0.09    | -0.53***

*Notes: This table reports summary statistics for the quarterly cross-sectional mean and dispersion of log Z-score, indicators for economic activity and systemic risk (A), bank failures (B), distressed acquisitions (C), and lending and interbank lending (D). Group A series are from FRED. Series in groups B and C are aggregated based on data from the FDIC and the Chicago Fed. Group D series are constructed from the Fed’s Z.1 and H.8 release. Data availability on bank holding companies restricts the analysis to 1986-2013. Sacf is the first-order sample autocorrelation coefficient. The last two columns report the correlation coefficients between cross-sectional mean and dispersion of log Z-score and each series in groups A-D. *, **, *** denote statistical significance at the 5%, 1%, and 0.1% level.*

To evaluate whether the distress dispersion has predictive power of aggregate indicators by providing additional information beyond what is contained in the average distress and existing systemic risk measures.

To this end, I run forecasting regressions of the above introduced aggregate indicators on the dispersion and mean of log Z-score controlling for moments including the term spread used in Giglio, Kelly, and Pruitt (2013), the leverage of both financial business and the security broker-dealers as in Adrian, Etula, and Muir (2014), and the growth rate of non-financial corporate liability as a measure of aggregate credit creation. The forecasting horizons range from one to four quarters and the data cover the years 1986-2013. To overcome correlation and autocorrelations in the time series, I calculate Newey-West standard errors.

Table 4 reports the coefficient estimates on the dispersion and mean of log Z-score, the values of $R^2$ when I run the regressions with and without the dispersion series. The regression results
Table 4. Predictive Regressions using Distress Dispersion

<table>
<thead>
<tr>
<th>Quarters</th>
<th>A. CFNAI</th>
<th>B. Failure Rate(%)</th>
<th>Asset-weighted Failure Rate(%)</th>
<th>C. Acquisition Rate(%)</th>
<th>Distressed over Total Acquisition Rate</th>
<th>D. Sml Bk Credit over Assets</th>
<th>Sml Bk Fed Funds over Assets</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Dispersion</td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
<td>Dispersion</td>
<td>1 2 3 4</td>
<td>Dispersion</td>
<td>1 2 3 4</td>
</tr>
<tr>
<td></td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
<td>1 2 3 4</td>
</tr>
<tr>
<td>Forecasting</td>
<td>A. CFNAI</td>
<td>NFCI</td>
<td>Forecasting</td>
<td>B. Failure Rate(%)</td>
<td>Asset-weighted Failure Rate(%)</td>
<td>Forecasting</td>
<td>C. Acquisition Rate(%)</td>
</tr>
<tr>
<td>Dispersion</td>
<td>-2.09***</td>
<td>-4.04***</td>
<td>-5.85***</td>
<td>-7.50***</td>
<td>1.52**</td>
<td>2.77**</td>
<td>3.83**</td>
</tr>
<tr>
<td>Mean</td>
<td>2.75</td>
<td>6.74</td>
<td>8.66</td>
<td>7.92</td>
<td>-8.95***</td>
<td>-17.80***</td>
<td>-25.73***</td>
</tr>
<tr>
<td>R²</td>
<td>44.85</td>
<td>52.03</td>
<td>54.05</td>
<td>52.48</td>
<td>53.22</td>
<td>53.40</td>
<td>52.24</td>
</tr>
<tr>
<td>R² w/o disp</td>
<td>28.15</td>
<td>34.78</td>
<td>36.47</td>
<td>34.74</td>
<td>37.54</td>
<td>39.28</td>
<td>39.42</td>
</tr>
</tbody>
</table>

Notes: This table summarizes the ability of distress dispersion to forecast future economic activity, systemic risk, failure rates, distressed acquisition rates, and bank lending behavior. Aggregate indicators in groups A-D are regressed respectively on the cross-sectional dispersion and mean of log Z-score controlling for the term spread, the leverage of financial business and security broker-dealers, and the growth rate of real non-financial corporate liability. Forecasting horizons range from one to four quarters and the data cover the years of 1986-2013. The table reports the regression coefficients of the dispersion and mean of log Z-score, the \( R^2 \), as well as the \( R^2 \) when the regressions are run without the dispersion series. *, **, *** denote statistical significance (based on Newey-West standard errors) at the 5%, 1%, and 0.1% level.

Echo those from the correlations and indicate striking predictive power of the dispersion series to forecast economic activity and systemic risk, failures, distressed acquisitions, and interbank lending. The predictive power is evidenced by both the economic significance of the regression coefficients and the differences in the \( R^2 \)s with and without dispersion in the regressors. For example, the estimates in the forecasting regression of CFNAI imply that (holding the mean fixed) a one-standard-deviation increase in Dispersion (=0.22) relates to a 0.46 (= 0.22 × 2.09) decrease in CFNAI. Notably, the national activity index CFNAI, the credit and loans and the
interbank lending of small and medium-sized commercial banks all respond negatively to an increase in distress dispersion, but not to changes in the mean of distress. Overall, these results paint a clear picture: the second moment of the cross-sectional distress distribution conveys new information about future activities in the financial sector in terms of systemic risk, failures, acquisitions, as well as interbank lending behavior.

7 Conclusion

Given the importance of financial interconnectedness, policies on financial stability and distress resolution should not analyze institutions in isolation. This paper has developed a network formation model to highlight a novel channel of systemic risk due to externalities via financial links.

Adding to the recent literature on financial network formation, this paper embeds firm heterogeneity in financial distress and examines how the linkage formation affects efficiency and systemic risk. I have shown that, when firms display high distress dispersion, the equilibrium network features too many links with the distressed firms and too few risk-sharing links among liquid firms. The reason is that the relatively more liquid firms have incentives to connect with distressed firms for profit while shifting risks away to their direct and indirect counterparties via the links. Particularly, these liquid firms fail to internalize the negative externalities when prices in the bilateral contracts cannot be contingent on the overall network structure. The inefficient link with the distressed firm not only generates risks of contagion but also crowd out valuable risk-sharing links, thereby increasing systemic risk. Notably, this inefficiency is shown to be more severe when institutions are more dispersed in financial distress.

While detailed data on the precise linkages among financial institutions are yet to be collected, this paper draws a relation between the degree of network inefficiency and the cross-sectional distribution of fundamentals, thus contributing to the measurement and forecast of systemic risk. The test can be extended along the lines of Giglio, Kelly, and Pruitt (2013) by comparing the distress dispersion to existing systemic risk measures such as CoVaR (Brunnermeier and Adrian (2011)) and Marginal and Systemic Expected Shortfall (Acharya, Pedersen, 47For current challenges in measuring linkages and systemic risk, see for example Bisias, Flood, Lo, and Valavanis (2012), Hansen (2013), and Yellen (2013).
Philippon, and Richardson (2010)). Additionally, my model predicts that links between firms with different distress levels respond differently to an aggregate dispersion increase. With possibly better data access in the future, more work is needed to test these qualitative predictions.

My model provides new insights on policies for financial stability. The links with distressed firms in the model can be interpreted as acquisitions of such firms. In this context, my results call for regulations to eliminate the network inefficiencies associated with acquisitions of distressed firms. The task of the regulators is to oversee the acquisitions of distressed firms, especially those by highly interconnected acquirers when the distress dispersion is high across institutions. Rather than relying on acquisitions as the preferred private sector solution, regulators should instead adopt resolution methods such as purchase and assumption (P&A) for these distressed targets in case of failure.

References


Spatt, Chester, 2009, Economic principles, government policy and the market crisis, Keynote Remark at the 2009 Western Finance Association meetings.


Appendix

A Technical Appendix

A.1 Notation Summary

General Notation:
- $t$: event dates, $t = 0, 1, 2, 3$.
- $i$: firm index, $i \in \{1, \ldots, N\}$.
- $R$: fixed return of asset.
- $\tilde{a}_i$: amount of liquid asset component.
- $\nu_i$: expectation of $\tilde{a}_i$ at $t = 1$.
- $\nu$: the vector of $\nu_i$.
- $\varepsilon_i$: idiosyncratic term in the liquid return, i.i.d. standard normal.
- $\sigma$: conditional volatility of liquid return.
- $c$: fixed liquidation cost.
- $z_i$: the distress statistic of firm $i$.
- $z$: the vector of $z_i$.
- $\bar{z}$: average distance to liquidation.
- $\delta$: distress dispersion across firms.
- $l_{ij}$: the fraction of firm $j$’s liquid asset that firm $i$ proposes to buy, $l_{ij} \in \{0, 1\}$.
- $\bar{l}$: fixed share of bilateral asset swaps.
- $l_i$: i’s strategy ($l_{i1}, \ldots, l_{i,i-1}, l_{i,i+1}, \ldots, l_{iN}$).
- $p_{ij}$: the unit price of firm $j$’s asset offered by buyer firm $i$.
- $p_i$: i’s strategy ($p_{i1}, \ldots, p_{i,i-1}, p_{i,i+1}, \ldots, p_{iN}$).
- $p_{jj}$: reservation price of firm $j$.
- $L$: the matrix representing two-sided links.
- $L_i$: the $i$-th row of matrix $L$.
- $\tilde{h}_i$: the final liquid asset holdings.
- $\Pi_i(z, L, p)$: firm payoff at date 3.
- $V_i(z, L, p)$: firm value under network $L$ and price $p$ at date 1.

Additional notation in Section 3:
- $L^*$: the optimal network.
- $\tilde{h}^*$: the optimal asset holdings under the optimal network $L^*$.
- $\bar{z}_1, \bar{z}_2$: the cutoff values of $\bar{z}$.
- $\delta_1(\bar{z})$: the cutoff function of $\delta(\bar{z})$ where the optimal network changes from a single connected component to isolating the distressed firm.
- $\delta_2(\bar{z})$: the cutoff function of $\delta(\bar{z})$ where the equilibrium network changes from over connection to inefficient composition.

Additional notation in Section 4:
- $\Delta V$: the value loss.
- $\Delta V\%$: the percentage value loss.
- $\Pr_{\text{sys}}^L$: the systemic risk in network $L$.
- $\Delta \Pr_{\text{sys}}$: the excess systemic risk.

Additional notation in Section 5:
- $\tau$: the acquisition tax.
- $\theta_i$: part of the expected liquid return realized after linkages are formed.
- $k$: a scalar indicating the negative shock to the distressed firm $N$, $\theta_N = -k\bar{z}\sigma$.
- $B\sigma$: government liquidity injection after linkages are formed.
- $B^*$: the optimal government liquidity injection policy.
- $\hat{k}$: a scalar indicating the positive shock to the potential acquirer firm $N + 1$, $\theta_{N+1} = \hat{k}\bar{z}\sigma$. 


A.2 Optimal Risk Sharing Allocation

This section provides the technical results for subsection 3.1. I show that under the iterative asset swap procedure, the asset holdings resulting from the optimal network $L^*$ is equivalent to the allocations when the social planner is allowed to choose the asset allocations directly.

**Definition 4** Let $H$ be an asset holding matrix such that firms’ liquid asset holdings are $\tilde{h} = H\tilde{a}$. The optimal asset allocation $H^*$ is feasible and minimizes total expected liquidation costs,

$$H^* = \arg \min_H \sum_i E\left[\tilde{h}_i < 1\right]c,$$

subject to

$$H \times 1_{N \times 1} = H^\top \times 1_{N \times 1} = 1_{N \times 1},$$}

where Equation (21) imposes the feasibility constraint for asset allocation. $H$ being a doubly stochastic matrix ensures that no assets are created or lost from asset pooling and that each firm still holds one unit of assets. The following lemma states that if it is optimal for a firm to have a diversified asset holding, then it is more likely to be a relatively liquid firm. Moreover, its optimal asset holding is full risk-sharing, i.e. it holds the equally weighted asset composed of assets of all firms that participate in risk-sharing.

**Lemma 2** If $\nexists i$ with $\tilde{h}_i^* = \tilde{a}_i$, then $\tilde{h}_i^* = \frac{1}{N} \sum_{j=1}^N \tilde{a}_j$, $\forall i \leq N$. If $\exists i$ with $\tilde{h}_i^* = \tilde{a}_i$ and $\tilde{h}_{i-1}^* \neq \tilde{a}_{i-1}$, then $\tilde{h}_j^* = \tilde{a}_j$, $\forall j \geq i$ and $\tilde{h}_j^* = \frac{1}{i-1} \sum_{k=1}^{i-1} \tilde{a}_k$, $\forall j \leq i - 1$.

**Proof** The optimal connection minimizes total expected liquidation costs (or equivalently default probabilities). Let the number of firms that participate in risk-sharing and have diversified asset holdings $\tilde{h} = H\tilde{a}$ be $M$. The total expected liquidation costs equal

$$\sum_{i=1}^M \Pr(\tilde{h}_i \leq 1)c = \sum_{i=1}^M \Phi\left(\frac{-\bar{z} - (1 - \sum_j H_{ij})\nu_i - \sum_j H_{ij}\nu_j}{\sqrt{(1 - \sum_j H_{ij})^2 + \sum_j H^2_{ij}\sigma}}\right)c.$$

(22)

The first order condition with respect to $H_{ij}$ is

$$\frac{\partial \sum_{i=1}^M \Pr(\tilde{h}_i \leq 1)c}{\partial H_{ij}} = \frac{\partial \Pr(\tilde{h}_i \leq 1)c}{\partial H_{ij}} + \frac{\partial \Pr(\tilde{h}_j \leq 1)c}{\partial H_{ji}}.$$

(23)

In particular, the derivative for firm $i$ is

$$\frac{\partial \Pr(\tilde{h}_i \leq 1)c}{\partial H_{ij}} = \Phi'\left(\frac{-\bar{z} - H_{ii}\nu_i - \sum_j H_{ij}\nu_j}{\sqrt{H_{ii}^2 + \sum_j H^2_{ij}\sigma}}\right)c \times$$

$$(\nu_i - \nu_j)\sqrt{H_{ii}^2 + \sum_j H^2_{ij}\sigma} + \left(-\bar{z} - H_{ii}\nu_i - \sum_j H_{ij}\nu_j\right)\sigma \left(H_{ii}^2 + \sum_j H^2_{ij}\right)^{-\frac{1}{2}}(H_{ii} - H_{ij}).$$

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Similarly, write out the symmetric equation for firm $j$ with respect to $H_{ji} = H_{ij}$, and plug into Equation (23) we obtain

$$
\frac{\partial \sum_{i=1}^{M} \Pr(\tilde{h}_i \leq 1)c}{\partial H_{ij}} \bigg|_{H_{ii}} = H_{ij} = \frac{1}{M}, \forall i \neq j.
$$

This implies that the first order conditions with respect to asset holdings equal zero when each element of $H$ is evaluated at $\frac{1}{M}$, thus achieving the optimal allocation. $H_{ij} = \frac{1}{M}$ indicates full risk-sharing. It is worth noting that the above result holds if we relabel $\Phi$ as a rather general distribution function even through the above proof explicitly uses normal distribution $\Phi$. The only condition necessarily required is that $\varepsilon_i$ is distributed independently across firms.

From Lemma 2, if no firm holds entirely idiosyncratic assets, then all firms share risks fully by having the same holding equally weighted by the liquid assets of all firms. If there exist firms who hold only their original assets, they must be the relatively distressed ones, while all other more liquid firms pool liquid assets equally. As such, the optimal asset holdings boil down to determining who should participate in risk-sharing and who should stay isolated.

Lemma 1 shows that the asset composition matrix $H^*$ implied by the optimal network $L^*$ also coincides with full risk-sharing among all connected firms. In this regard, under the iterative asset swap process, the optimal network $H^*$ in (P1) achieves the best asset allocation matrix $H^*$ in (P2). Importantly, the iterative feature of the asset swap itself does not deviate equilibrium from the optimal allocation.

A.3 Multiple Equilibria for $N \geq 5$

In Section 3.3, when analyzing the inefficiency in risk-sharing loss, I have focused on one specific equilibrium when firms start from the chain $\{1-2-3-4-5\}$ before distress dispersion increases.

In fact, when $N \geq 5$, different initial sequences in the stable risk-sharing chain at $\delta = 0$ imply different outside options and deviation incentives for each firm. As a result, we can have various equilibrium networks. The four panels in Figure A.1 illustrate different equilibrium connection structures in the $(\bar{z}, \delta)$ space. Importantly, across all potential equilibria, a general pattern displays: equilibrium network switches from optimal connection to over-connection to over- and under-connections simultaneously as $\delta$ increases.

A.4 Full Contingent Contracts

This section provides the technical results for subsection 3.4. I use an example of $N = 4$ and show that a complete set of contracts contingent on the entire network structure decentralizes the efficient network.

**Proposition 6** The efficient network is decentralized by bilateral contracts contingent on the entire network structure.
**Figure A.I. Equilibria \((N = 5)\).** This figure shows the equilibrium five-firm network. The horizontal and vertical axes represent the mean and dispersion of firm distress statistic \(z\). The four panels show different stable networks with different initial sequences along the chain. In colored regions, firm 5 is isolated in the optimal network. Blue (lighter) region denotes over-connection; orange (darkest) and yellow (lightest) regions denote inefficient network composition.

For \(N = 4\), the prices between 1 and 4 are \((p_{41}, p_{44})\) if 1 links with 2 and \((p_{11}, p_{44})\) otherwise; prices between 1 and 2 are \((p_{11}, p_{12})\) if 1 links with 4 and \((p_{21}, p_{22})\) otherwise; prices between 2 and 3 are \((p_{33}, p_{32})\) if 1 links with 4 and \((p_{33}, p_{32})\) otherwise, where

\[
p_{41} = p_{11} + \Phi(-z_1)c + \Phi(-z_4)c - 2\Pr(\hat{h}_4 < 1)c;
\]

\[
p_{12} = p_{22} + \max \left\{ \delta + 2\Pr(\hat{h}_4 < 1)c - 2\Pr\left(\frac{\tilde{a}_2 + \tilde{a}_3}{2} < 1\right)c + \Phi(-z_2)c - \Phi(-z_1)c, 0 \right\};
\]

\[
p_{21} = p_{11} + \delta\sigma + 2\Pr\left(\frac{\tilde{a}_2 + \tilde{a}_3}{2} < 1\right)c + 2\Pr\left(\frac{\tilde{a}_1 + \tilde{a}_2}{3} < 1\right)c
- 6\Pr(\hat{h}_4 < 1)c + \Phi(-z_1)c + 2\Phi(-z_2)c - \Phi(-z_4)c;
\]

\[
p_{32} = p_{22} + \delta\sigma + 4\Pr\left(\frac{\tilde{a}_1 + \tilde{a}_2 + \tilde{a}_3}{3} < 1\right)c - 6\Pr(\hat{h}_4 < 1)c + 2\Phi(-z_4)c,
\]

and \(\hat{h}_4 = \frac{\tilde{a}_1 + \tilde{a}_2 + \tilde{a}_3 + \tilde{a}_4}{4}\). Under these conditional prices \\{\(p_{41}, p_{12}, p_{21}, p_{32}\)\}, when \(\bar{z} \geq \bar{z}_1\) or when
\(\bar{z} \in [\bar{z}_2, \bar{z}_1]\) and \(\delta < \bar{\delta}(\bar{z})\), all firms are connected in the equilibrium network; when \(\bar{z} \in [\bar{z}_2, \bar{z}_1]\) and \(\delta > \bar{\delta}(\bar{z})\), distressed firm 4 is isolated in the equilibrium network.

**Proof** The proof is equivalent to show that under the bilateral contracts \(\{(p_{41}, p_{14}), (p_{11}, p_{44})\}, \{(p_{11}, p_{12}), (p_{21}, p_{22})\}\), and \(\{(p_{33}, p_{22}), (p_{33}, p_{33})\}\), \(L^c = \{1 - 2 - 3, 4\} \iff L^* = \{1 - 2 - 3, 4\}\), and \(L^c = \{4 - 1 - 2 - 3\} \iff L^* = \{4 - 1 - 2 - 3\}\). Next I check that firms have no incentive to deviate in the optimal network. In what follows, let \(\hat{V}^L_i\) denote the value of firm \(i\) in network \(L\) under reservation prices \(\{p_{ji}\}\), and \(V^L_i\) the value of firm \(i\) in network \(L\) under bilateral prices \(\{p_{ji}\}\).

If \(L^* = \{4 - 1 - 2 - 3\}\), under bilateral prices, firm values are

\[
\hat{V}_1^{4123}((p_{41}, p_{44}), (p_{11}, p_{12})) = \sum_{i=1,2,4} \hat{V}_i^{4123} - \hat{V}_4^a - \hat{V}_2^{2-3};
\]

\[
\hat{V}_2^{4123}((p_{41}, p_{44}), (p_{11}, p_{12})) = \hat{V}_2^{2-3};
\]

\[
\hat{V}_3^{4123}((p_{41}, p_{44}), (p_{11}, p_{12})) = \hat{V}_3^{4123};
\]

\[
\hat{V}_4^{4123}(p_{41}, p_{44}) = V^a_4.
\]

If \(L^* = \{1 - 2 - 3, 4\}\), under bilateral prices, firm values are

\[
\hat{V}_1^{4123}((p_{11}, p_{44}), (p_{21}, p_{22})) = \sum_{i=1,2,4} \hat{V}_i^{4123} - \hat{V}_4^a - \hat{V}_2^{2-3};
\]

\[
\hat{V}_2^{4123}((p_{21}, p_{22}), (p_{33}, p_{33})) = \hat{V}_2^{2-3};
\]

\[
\hat{V}_3^{4123}(p_{33}, p_{32}) = \sum_{i=1}^3 \hat{V}_i^{123} + V^a_4 - \sum_{i=1,2,4} \hat{V}_i^{4123};
\]

\[
\hat{V}_4^{4123} = V^a_4.
\]

When \(L^* = \{4 - 1 - 2 - 3\}\), no firm would deviate, so \(\{4 - 1 - 2 - 3\}\) is an equilibrium. In fact, by checking that all the other possible connection structures are not stable under these bilateral prices, full risk-sharing is the unique, stable, efficient equilibrium. When \(L^* = \{1 - 2 - 3, 4\}\), \(\sum_{i=1}^4 V_i^{123} + V^a_4 - \sum_{i=1,2,4} \hat{V}_i^{4123} > \hat{V}_3^{4123}\). This implies that firm 3 is willing to pay for the premium price \(p_{32}\). Similarly, by checking that all the other possible connection structures are not stable under these bilateral prices, full risk-sharing is the unique, stable, efficient equilibrium. Therefore, the above bilateral prices are able to decentralize the optimal network structures.

Next I show that under the price offering rules, these are the unique profit maximizing prices to decentralize the optimal networks. If \(L^* = \{4 - 1 - 2 - 3\}\), we have \(\sum_{i=1}^4 \hat{V}_i^{123} \geq \sum_{i=1}^2 \hat{V}_i^{123,4} + V_4^a\). Under the outside prices, there is a large region in the \(\{\bar{z}, \delta\}\) space in which 2 is better off to withdraw to form a link with 3. If this is the case, \(\hat{V}_2^{2-3} > \hat{V}_2^{4123}\), we require that firm 1 pays a premium \(\frac{1}{2} (p_{12} - p_{22}) = \hat{V}_2^{2-3} - \hat{V}_2^{4123}\) for \(L_{12} = \frac{1}{2}\) share of asset swap. Since firm 1 is offering a take-it-or-leave it offer to firm 2, the profit maximization behavior of firm 1 implies that

\[p_{12} = p_{22} + \max \left\{2 \left(\hat{V}_2^{2-3} - \hat{V}_2^{4123}\right), 0\right\}.
\]
The participation constraint for firm 4 implies that
\[
\frac{1}{2} (p_{41} - p_{11}) \leq \hat{V}_4^{14123} - V_4^a. \tag{24}
\]

For the equilibrium not to have the 1 – 4 link, we require that neither 2 nor 3 be worse off offering price premiums nor 1 is worse off accepting the prices. This means the sum of value of 1 and 2 and 3 is higher with 1 – 4 link than without,
\[
\sum_{i=1}^3 \hat{V}_i^{123} \leq \sum_{i=1}^3 \hat{V}_i^{14123} + \frac{1}{2} (p_{41} - p_{11})
\]

Combining with \(\sum_{i=1}^4 \hat{V}_i^{14123} \geq \sum_{i=1}^2 \hat{V}_i^{123,4} + V_4^a\), we require that
\[
\frac{1}{2} (p_{41} - p_{11}) \geq \hat{V}_4^{14123} - \hat{V}_4^a.
\]

Combining with Equation (24), we obtain
\[
\frac{1}{2} (p_{41} - p_{11}) = \hat{V}_4^{14123} - V_4^a.
\]

Therefore, the value of firm 1 under \(\{4 – 1 - 2 - 3\}\) is
\[
V_1^{4123} ((p_{41}, p_{44}), (p_{11}, p_{12})) = \hat{V}_1^{123} + \hat{V}_4^{14123} - V_4^a - \frac{1}{2} (p_{12} - p_{22}) \leq \sum_{i=1,2,4} \hat{V}_i^{14123} - V_4^a - \hat{V}_2^{2-3}.
\]

If instead \(L^* = \{1 – 2 – 3, 4\}\), \(\sum_{i=1}^2 \hat{V}_i^{123,4} + V_4^a \geq \sum_{i=1}^4 \hat{V}_i^{14123}\). It is sufficient to ensure pairwise stability if the following conditions hold: (1) 1 severs link with 4, (2) 2 stays link with 1, (3) 3 stays link with 2.
\[
\hat{V}_1^{123} + \frac{1}{2} (p_{21} - p_{11}) \geq \hat{V}_1^{14123} + \hat{V}_4^{14123} - V_4^a - \frac{1}{2} (p_{12} - p_{22})
\]
\[
\hat{V}_2^{123} - \frac{1}{2} (p_{21} - p_{11}) + \frac{1}{2} (p_{32} - p_{23}) \geq \hat{V}_2^{14123} + \frac{1}{2} (p_{12} - p_{22}) \geq \hat{V}_2^{2-3}
\]
\[
\hat{V}_3^{123} - \frac{1}{2} (p_{32} - p_{23}) \geq \hat{V}_3^{14123} \geq V_3^a
\]

Since firm 2 is offering the price premium, the minimum possible \(p_{21}\) is
\[
\frac{1}{2} (p_{12} - p_{22}) = \hat{V}_1^{14123} + \hat{V}_4^{14123} - V_4^a - \frac{1}{2} (p_{12} - p_{22}) - \hat{V}_1^{123} = \hat{V}_1^{14123} + \hat{V}_4^{14123} - V_4^a - \hat{V}_1^{123} - \hat{V}_2^{2-3} + \hat{V}_2^{4123}.
\]

Firm 3 needs to offer price premium \(p_{32}\) so that
\[
\hat{V}_2^{123} - \frac{1}{2} (p_{21} - p_{11}) + \frac{1}{2} (p_{32} - p_{23}) \geq \hat{V}_2^{14123} + \frac{1}{2} (p_{12} - p_{22}) = \hat{V}_2^{2-3}.
\]

So the minimum possible \(p_{32}\) is
\[
\frac{1}{2} (p_{32} - p_{23}) = \hat{V}_2^{2-3} - \hat{V}_2^{123} + \hat{V}_1^{14123} + \hat{V}_4^{14123} - V_4^a - \hat{V}_1^{123} - \hat{V}_2^{2-3} + \hat{V}_2^{4123}.
\]
Based on all the above analysis, the required profit-maximizing prices are uniquely given by
\[
\frac{1}{2} (p_{21} - p_{11}) = \hat{V}_1^{4123} + \hat{V}_4^{1123} - V_4^a - \hat{V}_1^{123} - \hat{V}_2^{123} - \hat{V}_2^{2-3} + \hat{V}_2^{123};
\]
\[
\frac{1}{2} (p_{32} - p_{23}) = \hat{V}_1^{4123} + \hat{V}_4^{1123} + \hat{V}_2^{1123} - \hat{V}_2^{123} - \hat{V}_4^a;
\]
\[
\frac{1}{2} (p_{12} - p_{22}) = \max \left\{ \left( \hat{V}_2^{2-3} - \hat{V}_2^{123} \right), 0 \right\};
\]
\[
\frac{1}{2} (p_{41} - p_{11}) = \hat{V}_4^{123} - V_4^a.
\]
Substituting the values, we recover the prices in Proposition 6.

A.5 Extension with Government Bailout

This section provides the technical results for Section 5.2. I consider slight variations of the baseline model where the timing of the network formation does not coincide with the observation of distress. Under the set up in Section 5.2, if the regulators had optimally isolated the distressed \( N \), the total liquidation cost is
\[
C_{\text{iso-N}} = (N - 1) \Phi \left[ \sqrt{N - 1} \left( -\bar{z} - \frac{1}{2} \delta \right) \right] c + \Phi \left[ (k - 1) \bar{z} + \frac{N - 1}{2} \right] c. \quad (25)
\]
In the absence of the acquisition tax, all firms are connected and the liquidation costs are
\[
C = \sum_{i=1}^{N} \Pr \left( \tilde{h}_i < 1 \right) c = N \Phi \left[ \frac{k - N}{\sqrt{N}} \bar{z} \right] c. \quad (26)
\]
When we enable the option of ex post government bailout as in 5.2.1, the costs equal liquidation plus bailout costs,
\[
C_{\text{GB}} = \sum_{i=1}^{N} \Pr \left( \tilde{h}_i < 1 \right) c + B \sigma = N \Phi \left[ \frac{(k - N) \bar{z} - B}{\sqrt{N}} \right] c + B \sigma. \quad (27)
\]
Notice that \( C = C_{\text{GB}} (B = 0) \). The gain from government bailout is \( C - C_{\text{GB}} \). The next proposition shows that as long as the fixed liquidation cost \( c \) is large enough, a positive government bailout that at least matches the expected liquid value shortfall is ex post optimal.

**Proposition 7** If \( c > \sqrt{\frac{2\pi \sigma}{Nc}} \), \( k > N \), the government bailout \( B^* \sigma \) generates positive surplus, where
\[
B^* = (k - N) \bar{z} + \sqrt{N} \left[ -2 \log \left( \frac{\sqrt{2\pi \sigma}}{\sqrt{Nc}} \right) \right].
\]

**Proof** The optimal liquid value injection policy \( B^* \) minimizes total costs \( C_{\text{GB}} \) and thus satisfies the first order condition \( \frac{\partial C_{\text{GB}}}{\partial B} = 0 \), i.e. \( N \Phi' \left[ \frac{(k - N) \bar{z} - B}{\sqrt{N}} \right] c \left( -\frac{1}{\sqrt{Nc}} \right) + \sigma = 0 \). This gives
\[
\Phi' \left[ \frac{(k - N) \bar{z} - B^*}{\sqrt{N}} \right] = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \left( \frac{(k - N) \bar{z} - B^*}{\sqrt{Nc}} \right)^2} = \frac{\sigma}{\sqrt{Nc}}. \quad (29)
\]
Solving for $B^*$ gives (28). Given $e^{-\frac{(N-k)^2}{2\pi}} \leq 1$, (29) implies $\frac{\sigma}{\sqrt{Nc}} \leq \frac{1}{\sqrt{2\pi}}$. Hence $c \geq \sqrt{2\pi\sigma}/\sqrt{N}$, i.e. the liquidation cost needs to be large enough.

Further, in order that $B^*$ archives the global minimum of $C_{GB}(B)$, we take the second derivative of $B$,

$$\frac{\partial^2 C_{GB}}{\partial B^2} = \Phi\left[\frac{(k-N)\bar{z}-B^*}{\sqrt{N}}\right] \geq 0, \quad \forall (k-N)\bar{z} - B^* \leq 0.$$

The second derivative is positive which ensures that $B^*$ archives the global minimum of $C_{GB}(B)$, so the bailout surplus is positive, $C - C_{GB} = C_{GB}(B = 0) - C_{GB}(B = B^*) > 0$. $B^* \geq (k-N)\bar{z}$ requires that $B^*$ at least matches the expected liquid value short fall, $B^*\sigma > (k-N)\bar{z}\sigma$. The extra liquidity injection depends on the uncertainty and cost tradeoff.

From Equation (28), $B^*\sigma$ at least matches the expected liquid value shortfall, $B^*\sigma > (k-N)\bar{z}\sigma$. $c > \sqrt{2\pi\sigma}/\sqrt{N}$ ensures that $B^*$ is non-zero. This requirement is easier to be satisfied when there are more counterparties to the distressed, and when uncertainty is lower. The extra liquidity injection, $\sqrt{N}\left(-2\log\left[\frac{\sqrt{2\pi\sigma}/\sqrt{N}}{\Phi}\right]\right)$, depends on the trade-off between cost and uncertainty. $\frac{\partial B^*}{\partial c} > 0$ implies that the bigger the liquidation cost is, the higher the optimal government bailout is; from $\frac{\partial B^*}{\partial \sigma} < 0$, optimal government bailout decreases with asset uncertainty.

If instead $0 \leq k \leq N$, the average distress after $\theta$ shock is positive. From Equation (28), a positive government bailout requires that $c \geq \sqrt{2\pi\sigma}/\sqrt{N} e^{(N-k)2/2\pi} > \sqrt{2\pi\sigma}/\sqrt{N}$. Plugging Equation (28) into (27), the total costs under optimal bailout policy $B^*$ is

$$C_{GB}^* = (k-N)\bar{z}\sigma + N\Phi\left[-\sqrt{-2\log\left[\frac{\sqrt{2\pi\sigma}}{\sqrt{Nc}}\right]}\right] + \sqrt{N\sigma}\sqrt{-2\log\left[\frac{\sqrt{2\pi\sigma}}{\sqrt{Nc}}\right]}.$$

Although $C_{GB}^*$ improves upon $C$, it is important to compare $C_{GB}^*$ with the cost when the acquisition link had been prevented *ex ante*.

**Proposition 8** There exists $\bar{c} > \sqrt{2\pi\sigma}/\sqrt{N}$, such that $C_{GB}^* > C_{isoN}$ for $c \in \left[\sqrt{2\pi\sigma}/\sqrt{N}, \bar{c}\right]$ and for all $\delta \geq 0$, where $c = (N-1)\Phi\left[-\sqrt{N-1}\bar{z}\right] + \Phi\left[(k-1)\bar{z}\right] - N\Phi\left[\frac{(k-N+1)\bar{z}}{\sqrt{N}}\right]$, and $B^*$ is given by Equation (28).

Therefore, when liquidation cost is bounded by $\bar{c}$, $C_{GB}^*$ is more costly than $C_{isoN}$.

Next I consider the optimal policy when there are healthier institutions currently not connected with the distressed. Denote the existing firms $i = 1,...,N$ as group one. Now, consider group two of $N$ other firms $i = N+1,...,2N$ with the same liquid value structure, $\bar{z} > 0, \sigma > 0$. For simplicity, let the dispersion $\delta$ among these firms be zero, so *ex ante* an optimal full risk-sharing network is formed. Let the additional signal $\theta_i$ be $\theta_{N+1} = k\bar{z}\sigma$ and $\theta_i = 0, \forall i = N+2,...,2N$, so *ex post* the $N+1$th firm gets a positive shock in the liquid value. The next corollary examines whether the *ex post* acquisition of heavily distressed $N$ by the liquid $N+1$ can reduce total liquidation costs, and if not, whether subsidized acquisition is value increasing.
Corollary 1 With no subsidy, the liquid firm $N + 1$ acquires the heavily distressed $N$ if and only if $\hat{k} \geq k - 2N$. Government subsidized acquisition is ex post optimal if $\hat{k} < k - 2N$ and $c > \frac{\sqrt{\pi \sigma}}{\sqrt{N}}$; the optimal subsidy to the acquirer firm $N + 1$ upon acquisition is $B^*_A \sigma$, where

$$B^*_A = (k - \hat{k} - 2N) \bar{z} + \sqrt{2N} \sqrt{-2 \log \left( \frac{\sqrt{\pi \sigma}}{\sqrt{Nc}} \right)}.$$  \hspace{1cm} (31)

When there exist healthier institutions, ex post subsidized acquisition is always preferred to ex post government bailout.

When the cardinality of the two groups differs, pushed acquisition could be ex post optimal. Denote $N_1$ (instead of $N$) the number of the group one firms including the heavily distressed $\theta_{N_1} = -k \bar{z} \sigma$. Consider $N_2$ other firms (group two), with the same $\bar{z} > 0$, $\sigma > 0$, but $\delta = 0$ for simplicity. Ex ante an optimal full risk-sharing network is formed among $N_2$ firms. The additional signal is $\theta_{N_1 + 1} = \hat{k} \bar{z} \sigma$, $\theta_i = 0$, $\forall i = N_1 + 2, \ldots N_1 + N_2$. Hence firm $i = N_1 + 1$ has the highest liquid value ex post. Suppose after $t = 1$ when links in each group are formed and prices are exchanged, the most liquid firm can acquire the heavily distressed.

Proposition 9 The social surplus of the acquisition is positive when the liquidity shocks satisfy

$$\hat{k} > \max \left[ \sqrt{\frac{N_1 + N_2}{N_1 + N_2 - N_1}}(k - N_1) - N_2, k - N_1 - N_2 \right]. \hspace{1cm} (32)$$

Under (32),

- when $N_2 \geq N_1$ the bilateral surplus is positive;
- when $N_2 < N_1$ the bilateral surplus is negative when

$$2\Phi \left[ \frac{k - \hat{k} - (N_1 + N_2)}{\sqrt{N_1 + N_2}} \bar{z} \right] c > \Phi \left[ \frac{k - N_1}{\sqrt{N_1}} \bar{z} \right] c + \Phi \left[ \frac{-\hat{k} - N_2}{\sqrt{N_2}} \bar{z} \right] c + \frac{(N_2 - N_1)(N_2 k + N_1 \hat{k})}{N_1 N_2 (N_1 + N_2)} \bar{z} \sigma. \hspace{1cm} (33)$$

As a sufficient condition for a positive social surplus, (32) sets a lower bound for the positive liquidity shock $\hat{k}$. The relative cardinality of the two groups of firms is essential in determining the sign of the bilateral surplus. When $N_2 > N_1$, on average, the pair of $i = \{N_1, N_1 + 1\}$ gets bigger surplus than an average bank. When $N_1 = N_2$, we recover the case in subsection 5.2.2, so the sign of the bilateral surplus matches that of the social surplus. When $N_1 > N_2$, under condition (33), bilateral surplus can be negative even if social surplus is positive. (33) implies an upper bound for $\hat{k}$, hence is especially relevant when the potential acquirer does not have an abundant supply of liquidity.

B  Proofs

B.1 Proof of Lemma 1

Before showing the properties of the asset composition matrix $L^\infty$, let us first check the features of the initial asset swap matrix $L$.  

55
Claim 1  The initial asset swap matrix $L$ is a doubly stochastic matrix. Its largest eigenvalue is 1, and all the other eigenvalues lie within the unit circle.

Proof  By construction, $L \times 1_{N \times 1} = 1_{N \times 1}$. Thus $L$ is a doubly stochastic matrix, $\lambda = 1$ is its eigenvalue with eigenvector $1_{N \times 1}$. Suppose for contradiction that there exists an eigenvalue $\lambda > 1$. Then there exists a non-zero vector $x$ such that $Lx = \lambda x > x$. However since the rows of $L$ are non-negative and sum to 1, each element of vector $Lx$ is a convex combination of the components of $x$. This implies that $\max[Lx] \leq \max[x]$, which contradicts with $\max[\lambda x] > \max[x]$. Hence all the eigenvalues cannot exceed 1 in absolute value.

More formally, we can resort to the properties of self-consistent norm. Let $\lambda$ be the eigenvalues and $x$ be the corresponding eigenvector. For any self-consistent matrix norm $\|\|$, we have

$$|\lambda| \times \|x\| = \|\lambda x\| = \|Lx\| \leq \|L\| \|x\|.$$ 

Because $x$ is non-zero, $|\lambda| \leq \|L\| = \max_j (\sum_i L_{ij}) = 1$.

Lastly, we need to show that $\lambda = -1$ is not an eigenvalue of $L$. It is equivalent to show that the matrix $L + I$ is non-singular. This can be seen from

$$\det(L + I) = \det \left( \begin{array}{cccc}
2 - \sum_{j\neq 1} L_{ij} & L_{12} & \ldots & L_{1M} \\
L_{21} & 2 - \sum_{j\neq 2} L_{ij} & \ldots & \ldots \\
\ldots & \ldots & \ldots & \ldots \\
L_{M1} & \ldots & \ldots & 2 - \sum_{j\neq M-1} L_{ij} \\
\end{array} \right)$$

All the off-diagonal elements are within 0 and 1. All the diagonal elements are within 1 and 2. For any column or row, the largest element is on the diagonal. Hence there are no columns or rows that are zero or linearly dependent. Therefore $\det(L + I) > 0$, and $\lambda = -1$ cannot be an eigenvalue of $L$.

Next we use the result from Claim 1 to show the limiting properties of $H = L^\infty$. Since $L$ is a symmetric matrix, all the eigenvalues $\{\lambda_1, \lambda_2, \ldots, \lambda_M\}$ are real. And there exists an orthogonal matrix $Q$ with $Q' = Q^{-1}$ such that $L^\infty = Q \Lambda Q^{-1}$, where $\Lambda = \text{diag}(\lambda_1, \lambda_2, \ldots, \lambda_M)$. And the columns of $Q$ are eigenvectors of unit length corresponding to the eigenvalues $\lambda_1, \lambda_2, \ldots, \lambda_M$.

Without loss of generality, we rank the eigenvalues $\lambda_i \geq \lambda_{i+1}$, then

$L^\infty = Q \Lambda Q^{-1} \ldots Q \Lambda Q^{-1} = Q \Lambda^\infty Q^{-1} = Q \left[ \begin{array}{cccc}
\lambda_1^\infty & 0 & \ldots & 0 \\
0 & \lambda_2^\infty & \ldots & \ldots \\
\ldots & \ldots & \ldots & \ldots \\
0 & \ldots & \ldots & \lambda_M^\infty \\
\end{array} \right] Q^{-1} \rightarrow Q \left[ \begin{array}{cccc}
1 & 0 & \ldots & 0 \\
0 & 0 & \ldots & \ldots \\
\ldots & \ldots & \ldots & \ldots \\
0 & \ldots & \ldots & 0 \\
\end{array} \right] Q^{-1}$,

where the last step follows from $\lambda_i < 1$ and $\lim \lambda_i^\infty = 0, \forall i \neq 1$. Let the first column of $Q$, which is the unit length eigenvector corresponding to $\lambda_1 = 1$ be $x_1$, then
positive whenever \( \bar{z} \geq 0 \). Therefore \( \Phi \) given by

\[
\begin{pmatrix}
1 - \sum_{j \neq 1} H_{ij} & H_{12} & \ldots & \ldots & H_{1M} \\
H_{21} & 1 - \sum_{j \neq 2} H_{ij} & \ldots & \ldots & \ldots \\
\vdots & \vdots & \ddots & \ddots & \vdots \\
H_{M1} & \ldots & \ldots & 1 - \sum_{j \neq M-1} H_{ij} & \ldots \\
\end{pmatrix}
\begin{pmatrix}
x_{11} \\
x_{12} \\
\vdots \\
x_{1M} \\
\end{pmatrix} =
\begin{pmatrix}
x_{11} \\
x_{12} \\
\vdots \\
x_{1M} \\
\end{pmatrix}
\]

(34)

Combining Equations (34) and (35), we can solve for the unit length eigenvectors as \( x_{11} = x_{12} = \ldots = x_{1M} = \frac{1}{\sqrt{M}} \). In this case,

\[
L^\infty = Q^{-1} = \begin{pmatrix}
1 & 0 & \ldots & 0 \\
0 & 0 & \ldots & \ldots \\
\vdots & \vdots & \ddots & \ddots \\
0 & \ldots & \ldots & 0 \\
\end{pmatrix}
\]

Finally, since \( L \) is a doubly stochastic matrix, \( L \times 1_{N \times 1} = 1_{N \times 1} \), \( L^\top \times 1_{N \times 1} = 1_{N \times 1} \). Then \( L^N \times 1_{N \times 1} = L^{N-1} \times L \times 1_{N \times 1} = L^{N-1} \times 1_{N \times 1} = 1_{N \times 1} \). Similarly \( L^{TN} \times 1_{N \times 1} = 1_{N \times 1} \), so \( H = L^\infty \) is also a doubly stochastic matrix. Q.E.D.

B.2 Proof of Proposition 1

Let us start by analyzing the risk-sharing decision of \( N = 2 \) in the following lemma.

**Lemma 3** The risk-sharing surplus for \( N = 2 \) is positive if and only if \( \bar{z} \geq 0 \); the risk-sharing surplus increases monotonically with \( \delta \).

**Proof** The total liquidation cost for two separate firms with distress \( \{\bar{z} + \frac{1}{2}\delta, \bar{z} - \frac{1}{2}\delta\} \) is \( \Pr (\bar{a}_1 \leq 1) c + \Pr (\bar{a}_2 \leq 1) c = \Phi [-\bar{z} - \frac{1}{2}\delta] c + \Phi [-\bar{z} + \frac{1}{2}\delta] c \). The total liquidation cost for the two firms to fully share risk is \( \Pr (\bar{h}_1 \leq 1) c + \Pr (\bar{h}_2 \leq 1) c = 2\Phi [-\sqrt{2\bar{z}}] c \). The bilateral risk-sharing surplus is given by

\[
\Pr (\bar{a}_1 \leq 1) c + \Pr (\bar{a}_2 \leq 1) c - 2 \Pr (\bar{h}_1 \leq 1) c = \Phi [-\bar{z} - \frac{1}{2}\delta] c + \Phi [-\bar{z} + \frac{1}{2}\delta] c - 2\Phi [-\sqrt{2\bar{z}}] c.
\]

Function \( \Phi(x) \) is monotonically increasing for all \( x \), and is convex for \( x < 0 \) \((\Phi'' > 0, \forall x < 0)\). Therefore \( \Phi [-\bar{z} - \frac{1}{2}\delta] c + \Phi [-\bar{z} + \frac{1}{2}\delta] c \geq 2\Phi [-\bar{z}] c \geq 2\Phi [-\sqrt{2\bar{z}}] c \), i.e. the surplus is positive whenever \( \bar{z} > 0, \delta > 0 \). The first derivative with respect to \( \delta \) is \( -\frac{1}{2} \Phi' [-\bar{z} - \frac{1}{2}\delta] c + \frac{1}{2} \Phi' [-\bar{z} + \frac{1}{2}\delta] c = \frac{\epsilon}{2} (\Phi' [-\bar{z} + \frac{1}{2}\delta] - \Phi' [-\bar{z} - \frac{1}{2}\delta]) > 0 \), for \(-\bar{z} + \frac{1}{2}\delta > -\bar{z} - \frac{1}{2}\delta\), i.e. \( \delta > 0 \). □
Next I show that $\delta$ matters for the optimal risk-sharing policy for $N \geq 3$. The total default probability in a full risk-sharing network with $N$ firms is

$$\sum_{i=1}^{N} \Pr(\tilde{h}_i \leq 1) = N \times \Phi \left[ \sqrt{N}(-\bar{z}) \right].$$

The total default probability when the best $N - 1$ firms fully share risk, isolating the most distressed firm is

$$\sum_{i=1}^{N-1} \Pr(\tilde{h}_i \leq 1) + \Pr(\tilde{a}_N \leq 1) = (N - 1) \times \Phi \left[ \sqrt{N-1}(-\bar{z} - \frac{1}{2}\delta) \right] + \Phi \left[ -\bar{z} - \frac{1 - N}{2}\delta \right].$$

The difference between the above two terms is

$$\sum_{i=1}^{N-1} \Pr(\tilde{h}_i \leq 1) + \Pr(\tilde{a}_N \leq 1) - \sum_{i=1}^{N} \Pr(\tilde{h}_i \leq 1) = (N - 1) \times \Phi \left[ \sqrt{N-1}(-\bar{z} - \frac{1}{2}\delta) \right] + \Phi \left[ -\bar{z} - \frac{1 - N}{2}\delta \right] - N \times \Phi \left[ \sqrt{N}(-\bar{z}) \right]. \quad (36)$$

When $\delta \to \infty$, the limit becomes

$$\lim_{\delta \to \infty} \sum_{i=1}^{N-1} \Pr(\tilde{h}_i \leq 1) + \Pr(\tilde{a}_N \leq 1) - \sum_{i=1}^{N} \Pr(\tilde{h}_i \leq 1) = 1 - N \times \Phi \left[ \sqrt{N}(-\bar{z}) \right]. \quad (37)$$

Then as long as $\bar{z}$ is large enough that $\Phi \left[ \sqrt{N}(-\bar{z}) \right] < \frac{1}{N}$, full risk-sharing dominates, i.e. $\bar{z}_1 = -\frac{1}{N} \Phi^{-1}(\frac{1}{N})$. If we consider an upper bound on $\delta$ in our analysis, say the upper bound is $\delta \leq 2$, then equating Equation (36) to zero and plugging in $\delta = 2$, $\bar{z}_1$ solves

$$(N - 1) \Phi \left[ \sqrt{N-1}(-\bar{z}) \right] + \Phi \left[ N - \bar{z} \right] = N \Phi \left[ \sqrt{N}(-\bar{z}) \right].$$

When the best $N - 2$ firms fully share risk whereas the two most distressed firms are isolated, we have that the total default probability becomes

$$\sum_{i=1}^{N-2} \Pr(\tilde{h}_i \leq 1) + \Pr(\tilde{a}_{N-1} \leq 1) + \Pr(\tilde{a}_N \leq 1)$$

$$= (N - 2) \Phi \left[ \sqrt{N-2}(-\bar{z} - \delta) \right] + \Phi \left[ -\bar{z} - \frac{3 - N}{2}\delta \right] + \Phi \left[ -\bar{z} - \frac{1 - N}{2}\delta \right].$$

The difference between isolating one or two distressed firms is

$$\sum_{i=1}^{N-2} \Pr(\tilde{h}_i \leq 1) + \Pr(\tilde{a}_{N-1} \leq 1) + \Pr(\tilde{a}_N \leq 1) - \sum_{i=1}^{N-1} \Pr(\tilde{h}_i \leq 1) - \Pr(\tilde{a}_N \leq 1)$$

$$= (N - 2) \Phi \left[ \sqrt{N-2}(-\bar{z} - \delta) \right] + \Phi \left[ -\bar{z} - \frac{3 - N}{2}\delta \right] - (N - 1) \times \Phi \left[ \sqrt{N-1}(-\bar{z} - \frac{1}{2}\delta) \right].$$
When \( \delta \to \infty \), the limit of the function is 1. When \( \tilde{z} = 0 \), the RHS becomes \( (N - 2)\Phi \left[ \sqrt{N - 2}(-\delta) \right] + \Phi \left[ \frac{N-3}{2}\delta \right] - (N - 1) \times \Phi \left[ \sqrt{N - 2}(-\frac{1}{2}\delta) \right] < 0 \) for small values of \( \delta \) and when \( N > 4 \). The curve \( (N - 2)\Phi \left[ \sqrt{N - 2}(-\tilde{z} - \delta) \right] + \Phi \left[ -\tilde{z} - \frac{3-N}{2}\delta \right] = (N - 1)\Phi \left[ \sqrt{N - 1}(-\tilde{z} - \frac{1}{2}\delta) \right] \) is concave with \( \delta \) and convex with \( \tilde{z} \). Denote \( \tilde{z}_2 \) the maximum value of \( \tilde{z} \) on this curve. Then for \( \tilde{z} > \tilde{z}_2 \), isolating one distressed firm is preferred to isolating two firms. In this case, the cutoff curve \( \delta_1(\tilde{z}) \) is defined by

\[
(N - 1) \times \Phi \left[ \sqrt{N - 1}(-\tilde{z} - \frac{1}{2}\delta_1(\tilde{z})) \right] + \Phi \left[ -\tilde{z} - \frac{1 - N}{2}\delta_1(\tilde{z}) \right] = N \times \Phi \left[ \sqrt{N}(-\tilde{z}) \right].
\]

From the implicit function theorem, the curve is well-defined, and

\[
\frac{\partial \delta_1(\tilde{z})}{\partial \tilde{z}} = -\frac{N\sqrt{N}\Phi' \left[ \sqrt{N}(-\tilde{z}) \right] - (N - 1)\sqrt{N - 1}\Phi' \left[ \sqrt{N - 1}(-\tilde{z} - \frac{1}{2}\delta) \right] - \Phi' \left[ -\tilde{z} - \frac{1 - N}{2}\delta \right]}{\frac{N - 1}{2} (\Phi' \left[ -\tilde{z} + \frac{N - 1}{2}\delta \right] - \sqrt{N - 1}\Phi' \left[ \sqrt{N - 1}(-\tilde{z} - \frac{1}{2}\delta) \right])} > 0.
\]

Q.E.D.

**B.3 Proof of Proposition 2**

The proof is equivalent to show that there do not exist bilateral prices \( (p_{21}, p_{12}), (p_{41}, p_{14}), (p_{32}, p_{23}) \) that can decentralize the optimal network in the parameter region \( \{ \tilde{z} \in [1, \tilde{z}_1], \delta > \delta_1(\tilde{z}) \} \). In other words, when \( L^* = \{4, 1 - 2 - 3\} \), there does not exist a feasible premium price \( p_{21} \) offered by firm 2 that prevents firm 1 from linking with 4.

In what follows, let \( \hat{V}_i^L \) denote the value of firm \( i \) in network \( L \) under reservation prices \( \{p_{ii}\} \), and \( V_i^L \) the value of firm \( i \) in network \( L \) under bilateral prices \( \{p_{ji}\} \).

If \( L^* = \{4 - 1 - 2 - 3\} \), we have \( \sum_{i=1}^{4} \hat{V}_i^{4-1-2-3} \geq \sum_{i=1}^{3} \hat{V}_i^{1-2-3} + V_4^a \). Under the outside prices, there is a large region in the \( \{\tilde{z}, \delta\} \) space in which 2 is better off to withdraw and form risk-sharing pair with 3. If this is the case, \( \hat{V}_2^{2-3} > \hat{V}_2^{4-1-2-3} \), we require that firm 1 pays at least a premium \( \frac{1}{2} (p_{12} - p_{22}) = \hat{V}_2^{2-3} - \hat{V}_2^{4-1-2-3} \) for \( L_{12} = \frac{1}{2} \) share of asset swap so that \( V_2^{4-1-2-3} (p_{12}, p_{11}) = \hat{V}_2^{2-3} \).

\[
V_1^{4-1-2-3} ((p_{41}, p_{14}), (p_{22}, p_{12})) = \hat{V}_1^{4-1-2-3} + \frac{1}{2} (p_{41} - p_{11}) - \frac{1}{2} (p_{12} - p_{22}) \geq \hat{V}_1^{1-2-3}; \quad (38)
\]

\[
V_2^{4-1-2-3} ((p_{22}, p_{12}), (p_{32}, p_{23})) = \hat{V}_2^{4-1-2-3} + \frac{1}{2} (p_{12} - p_{22}) + \frac{1}{2} (p_{32} - p_{23}) \geq \hat{V}_2^{2-3}; \quad (39)
\]

\[
V_3^{4-1-2-3} (p_{32}, p_{23}) = \hat{V}_3^{4-1-2-3} - \frac{1}{2} (p_{32} - p_{23}) \geq V_3^a. \quad (40)
\]

\[
V_4^{4-1-2-3} (p_{41}, p_{14}) = \hat{V}_4^{4-1-2-3} - \frac{1}{2} (p_{41} - p_{11}) \geq V_4^a. \quad (41)
\]

From (39), the minimum price offered by 1 is

\[
\frac{1}{2} (p_{12} - p_{22}) = \hat{V}_2^{2-3} - \hat{V}_2^{4-1-2-3} - \frac{1}{2} (p_{32} - p_{23}). \quad (42)
\]

Let us pick the upper bound of prices \( \frac{1}{2} (p_{32} - p_{23}) \) and \( \frac{1}{2} (p_{41} - p_{11}) \) from the participation
constraints (40) and (41), the value firm 1 gets by linking with 4 is

\[ V_1^{4-1-2-3} ((p_{41}, p_{44}), (p_{22}, p_{12})) \]
\[ = \hat{V}_1^{4-1-2-3} + V_2^{4-1-2-3} + \frac{1}{2} (p_{32} - p_{23}) + \frac{1}{2} (p_{41} - p_{11}) - \hat{V}_2^{2-3} \]
\[ = \hat{V}_1^{4-1-2-3} + V_2^{4-1-2-3} + \hat{V}_4^{4-1-2-3} - V_4^a + \hat{V}_3^{4-1-2-3} - \hat{V}_3^a - \hat{V}_2^{2-3} \]
\[ > \sum_{i=1}^4 \hat{V}_i^{4-1-2-3} - V_4^a - \hat{V}_3^a - \hat{V}_2^{2-3} > \sum_{i=1}^4 \hat{V}_i^{4-1-2-3} - V_4^a - \hat{V}_3^{2-3} - \hat{V}_2^{2-3} \]
\[ > \hat{V}_1^{1-4} + \hat{V}_4^{1-4} - V_4^a. \]

This shows that paying the premium (42) to prevent 2 from withdrawing is always a dominating strategy for firm 1. The value of 1 in \( L = \{4-1-2-3\} \) is

\[ V_1^{4-1-2-3} ((p_{41}, p_{44}), (p_{22}, p_{12})) \]
\[ = \hat{V}_1^{4-1-2-3} + V_2^{4-1-2-3} + \frac{1}{2} (p_{32} - p_{23}) + \hat{V}_4^{4-1-2-3} - V_4^a - \hat{V}_2^{2-3}. \]

And equilibrium replicates the optimal connection \( L^* = L = \{4-1-2-3\} \).

If \( L^* = \{4, 1-2-3\} \), we have \( \sum_{i=1}^3 \hat{V}_i^{1-2-3} + V_4^a \geq \sum_{i=1}^4 \hat{V}_i^{4-1-2-3} \). Under the outside prices, for all the region, firm 1 wants to link with 4 and firm 2 wants to withdraw. We require that firm 2 pays at least a premium \( \frac{1}{2} (p_{21} - p_{11}) \) to prevent 1 from linking with 4.

\[ V_1^{4,1-2-3} (p_{11}, p_{21}) = \hat{V}_1^{1-2-3} + \frac{1}{2} (p_{21} - p_{11}) \geq V_1^{4-1-2-3} ((p_{41}, p_{44}), (p_{22}, p_{12})); \]

\[ V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})) = \hat{V}_2^{1-2-3} - \frac{1}{2} (p_{21} - p_{11}) + \frac{1}{2} (p_{32} - p_{23}) \geq \hat{V}_2^{2-3}; \]

\[ V_3^{4,1-2-3} (p_{32}, p_{23}) = \hat{V}_3^{1-2-3} - \frac{1}{2} (p_{32} - p_{23}) \geq V_3^a. \]

\[ V_4^{4,1-2-3} = V_4^a. \]

Notice that \((p_{32}, p_{23})\) is not contingent on the link of 1–4, thus it has the same value in both structures. From (43), the minimum required incentive offered by firm 2 to 1 is

\[ \frac{1}{2} (p_{21} - p_{11}) = \hat{V}_1^{4-1-2-3} + \hat{V}_2^{4-1-2-3} + \frac{1}{2} (p_{32} - p_{23}) + \hat{V}_4^{4-1-2-3} - V_4^a - \hat{V}_2^{2-3} - \hat{V}_1^{1-2-3}. \]

Plugging into (44), the value of firm 2 then becomes

\[ V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})) = \hat{V}_2^{2-3} + \hat{V}_2^{1-2-3} + V_4^a + \hat{V}_1^{4-1-2-3} - \left( \hat{V}_1^{4-1-2-3} + \hat{V}_2^{1-2-3} + \hat{V}_4^{4-1-2-3} \right). \]

The total surplus of firms 1,2,4

\[ \hat{V}_1^{1-2-3} + V_4^a + \hat{V}_1^{1-2-3} - \left( \hat{V}_1^{4-1-2-3} + \hat{V}_2^{1-2-3} + \hat{V}_4^{4-1-2-3} \right) \]

is the group surplus of 1,2,4 in \( \{4, 1-2-3\} \) compared to that in \( \{4-1-2-3\} \). The surplus can be expressed as
We require (38), (39), (43), and (45). When evaluated at $\delta = 0$, the surplus is $-2\Phi(\bar{z}) + 2\Phi(-\sqrt{3}\bar{z}) + 3\Phi(-2\bar{z}) < 0$. Take the derivative of $\delta$, it establishes that

$$\sqrt{3}\Phi'\left(-\sqrt{3}\bar{z} - \frac{\sqrt{3}}{2}\delta\right) - \frac{3}{2}\Phi'\left(-\bar{z} + \frac{3}{2}\delta\right) - \frac{1}{2} - 6\Phi'[-2\bar{z}] < 0, \forall \delta > 0, \bar{z} > 0.$$ 

which follows from $\Phi'\left[-\bar{z} + \frac{3}{2}\delta\right] + \frac{1}{2} + 6\Phi'[-2\bar{z}] > \Phi'(0) + \Phi'[-2\bar{z}] > 2\Phi'[-\bar{z}] > \sqrt{3}\Phi'[-\bar{z}] > \sqrt{3}\Phi'\left[-\sqrt{3}\bar{z} - \frac{\sqrt{3}}{2}\delta\right]$. Therefore,

$$\hat{V}_2^{1-2-3} + \hat{V}_3^{1-2-3} < \hat{V}_1^{4-1-2-3} + \hat{V}_2^{4-1-2-3} + \hat{V}_4^{4-1-2-3} + \hat{V}_3^{1-2-3}.$$ 

This further implies

$$V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})) < \hat{V}_2^{2-3} = V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})).$$

Firm 2 is worse off providing the required premium price $p_{21}$ than staying in the full connection $L = \{4 - 1 - 2 - 3\}$. Therefore, the efficient network is not stable. In other words, the equilibrium fails to replicate the optimal connection $L^e = \{4 - 1 - 2 - 3\} \neq L^* = \{4, 1 - 2 - 3\}$. This further implies

$$V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})) < \hat{V}_2^{2-3} = V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})).$$

I next show that even if we relax the price offering rule, there still do not exist feasible bilateral prices between 1 and 2 to effectively prevent the 4 - 1 link. For $L^* = \{4, 1 - 2 - 3\}$ to be stable, we require (38), (39), (43), and

$$V_2^{4,1-2-3} ((p_{11}, p_{21}), (p_{32}, p_{23})) = \hat{V}_2^{1-2-3} - \frac{1}{2} (p_{21} - p_{11}) + \frac{1}{2} (p_{32} - p_{23}) \geq V_2^{4,1-2-3} ((p_{22}, p_{12}), (p_{32}, p_{23})).$$

Combining all these conditions, we require

$$\hat{V}_2^{1-2-3} + \hat{V}_1^{1-2-3} \geq \hat{V}_2^{4,1-2-3} + \hat{V}_1^{4-1-2-3} + \hat{V}_4^{4-1-2-3} + \frac{1}{2} (p_{41} - p_{11}).$$

From (41), we require

$$\hat{V}_2^{1-2-3} + \hat{V}_1^{1-2-3} \geq \hat{V}_2^{4,1-2-3} + \hat{V}_1^{4-1-2-3} + \hat{V}_4^{4-1-2-3} - V_4^a. \quad (45)$$

Consider the region around cutoff curve $\delta_1(\bar{z})$, where $\sum_{z=1}^{3} \hat{V}_i^{1-2-3} + V_4^a = \epsilon + \sum_{z=1}^{4} \hat{V}_i^{1-2-3}$. The total values under $L^*$ is only slightly greater than that under $L = \{4 - 1 - 2 - 3\}$, but the value difference for firm 3 is big especially when dispersion $\delta$ is large,

$$\hat{V}_3^{1-2-3} - \hat{V}_3^{4-1-2-3} = \frac{1}{2} \delta + \Pr(\hat{h}_4 < 1) \frac{\delta_1 + \tilde{a}_2 + \tilde{a}_3}{3} c > \epsilon,$$

In this case, (45) does not hold: there does not exist bilateral price $(p_{21}, p_{12})$ to prevent the formation of the 4 - 1 link. Q.E.D.
B.4 Proof of Proposition 3

First consider the case with $N = 5$ firms. The distress vector is $z = \{\bar{z} + 2\delta, \bar{z} + \delta, \bar{z}, \bar{z} - \delta, \bar{z} - 2\delta\}$. I next show that the proposition holds in different equilibrium networks, $\{5 - 1 - 2 - 3 - 4\}$, $\{5 - 1 - 3 - 2 - 4\}$, $\{5 - 1 - 4 - 3 - 2\}$, $\{5 - 1 - 4 - 2 - 3\}$. In all these structures, I start from the optimal full risk-sharing network and solve for the bilateral prices that decentralize the full risk-sharing network. Then I fix the contracts between 2, 3, 4 (because these contract should not change with the link 1 – 5), and check whether any agent deviates from full connection.

Case 1. I check the stability and efficiency of the chain $\{5 - 1 - 2 - 3 - 4\}$. Firm values under outside prices are given by $\{\hat{V}_{151234}, \hat{V}_{251234}, \hat{V}_{351234}, \hat{V}_{451234}, \hat{V}_{551234}\}$.

Notice that the bilateral prices between 2 – 3 – 4 are not contingent on $L_{15}$. To decentralize the optimal full network, let $p_{43} - p_{34} = \frac{1}{2}\delta\sigma$, $p_{32} - p_{23} = \frac{1}{2}\delta\sigma$, $\frac{1}{2}(p_{12}(L_{15}) - p_{22}) = \sum_{i=2}^{4}\hat{V}_{i234} - V_{3}^{a} - V_{4}^{a} - \hat{V}_{251234}$, so that $V_{2}^{51234} = \sum_{i=2}^{4}\hat{V}_{i234} - V_{3}^{a} - V_{4}^{a}$ (outside option of 2 in $\{4 - 2 - 3\}$). As a result, $V_{151234} = \hat{V}_{151234} - V_{2}^{a} + \hat{V}_{251234} - \sum_{i=2}^{4}\hat{V}_{i234} + V_{3}^{a} + V_{4}^{a}$, $V_{3}^{51234} = \hat{V}_{351234}$, $V_{4}^{51234} = \hat{V}_{451234} - \frac{1}{2}\delta\sigma$. I then check that it is not feasible for 2 to offer incentive to 1 not to link with 5, i.e. $L^{*} = \{5, 1 - 2 - 3 - 4\}$ cannot be decentralized.

Next I check if the full-connection is pairwise stable. Since both 2 and 5 are offered by the contingent contracts exactly their respective outside options, we only need to check the deviation incentives for firm 1, 3, 4. Results show that (1) neither 3 or 4 deviates; (2) there is a region in which 1 is better off severing the linkage with 2 so $L^{c} = \{5 - 1, 3 - 2 - 4\}$; (3) for the rest regions, full connection is stable so $L^{c} = \{5 - 1 - 2 - 3 - 4\}$.

Case 2. I check the stability of $\{5 - 1 - 3 - 2 - 4\}$ using the same logic. Firm values under outside prices are given by $\{\hat{V}_{151324}, \hat{V}_{251324}, \hat{V}_{351324}, \hat{V}_{451324}, \hat{V}_{551324}\}$.

Let the bilateral prices be $p_{23} - p_{32} = \frac{1}{2}\delta\sigma$, $\frac{1}{2}(p_{3}(L_{15}) - p_{33}) = V_{3}^{a} - \hat{V}_{351324} - \frac{1}{2}\delta\sigma$, so that $V_{3}^{51324} = V_{3}^{a}$ (outside option of 3 in $\{4 - 2 - 3\}$). So $V_{151324} = \hat{V}_{151324} + \hat{V}_{551324} - V_{5}^{a} + \hat{V}_{351324} - V_{3}^{a} + \hat{V}_{251324} - \frac{1}{2}\delta\sigma$, $V_{3}^{51324} = \hat{V}_{351324}$, $V_{4}^{51324} = \hat{V}_{451324}$, $V_{5}^{51324} = \hat{V}_{551324}$.

In this case, it is not feasible for 3 to offer incentive to 1 not to link with 5, i.e. $L^{*} = \{5, 1 - 3 - 2 - 4\}$ cannot be decentralized.

I then check if full-connection (all banks connected in one component) is stable. After computing the deviation incentives of 1, 2, 4, results show that (1) neither banks 2 or 4 deviates; (2) there is a region in which 1 is better off severing the linkage with 3 and so $L^{c} = \{5 - 1, 3 - 2 - 4\}$; (3) for the rest regions, full connection is stable so $L^{c} = \{5 - 1 - 3 - 2 - 4\}$.

Case 3. I check the stability of $\{5 - 1 - 4 - 3 - 2\}$ using the same logic. Firm values under outside prices are given by $\{\hat{V}_{151432}, \hat{V}_{251432}, \hat{V}_{351432}, \hat{V}_{451432}, \hat{V}_{551432}\}$.
In order to decentralize the full risk-sharing network, we require that the bilateral prices between 1 and 4 be contingent on \(L_{15}\) to prevent 4 from withdrawing, and that the bilateral prices between 4 and 3, 3 and 2 be independent of \(L_{15}\) link.

Let \(p_{43} - p_{34} = \frac{1}{2}\delta\sigma, p_{32} - p_{23} = \frac{1}{2}\delta\sigma, \) and \(\frac{1}{2}(p_{44}(L_{15}) - p_{44}) = V^a_4 - \hat{V}^{5,14,32}_4 + \frac{1}{2}\delta\sigma, \) so that \(V^{5,14,32}_4 = V^a_4 \) (outside option of 4 in \(\{4 \rightarrow 2 \rightarrow 3\}\)). So \(V^{5,14,32}_1 = \hat{V}^{5,14,32}_1 + V^a_3 - V^a_5 + \hat{V}^{5,14,32}_5 - \frac{1}{2}\delta\sigma - V^a_4, V^{5,14,32}_2 = \hat{V}^{5,14,32}_2 + \frac{1}{2}\delta\sigma, V^{5,14,32}_3 = \hat{V}^{5,14,32}_3, V^{5,14,32}_4 = V^a_4, V^{5,14,32}_5 = V^a_5. \) Similarly, it is not feasible for 4 to offer incentive to 1 not to link with 5, i.e. \(L^e = \{5, 1 - 4 - 2 - 3\}\) cannot be decentralized.

Next check whether the full-connection is pairwise stable by computing the deviation incentives for 1, 2, 3. It turns out that for a large region, \(V^{5,14,32}_2 < V^a_2\) and firm 2 withdraws from the end of the chain. Given that 2 withdraws, we need to check if \(\{5 - 1 - 4 - 3\}\) is stable, we compare \(V^{5-1-4-3}_1 = \hat{V}^{5-1-4}_1 + \hat{V}^{5-1-4}_4 + V^{5-1-4}_3 - V^a_4 - V^a_5\) with \(\hat{V}^{5-1-4}_1\). If \(V^{5-1-4-3}_1 > \hat{V}^{5-1-4}_1\), \(L^e = \{5 - 1 - 4 - 3, 2\}\); otherwise \(L^e = \{5 - 1, 4 - 2 - 3\}\), which has wrong network composition compared to \(L^e\).

**Case 4.** We move to check the stability of \(\{5 - 1 - 4 - 2 - 3\}\). Firm values under outside prices are given by \(\{\hat{V}^{5,14,23}_1, V^{5,14,23}_2, V^{5,14,23}_3, \hat{V}^{5,14,23}_4, V^{5,14,23}_5\}\).

Let \(p_{23} - p_{32} = \frac{1}{2}\delta\sigma, \) and \(\frac{1}{2}(p_{14}(L_{15}) - p_{44}) = V^a_4 - \hat{V}^{5,14,32}_4, \) so that \(V^{5,14,32}_1 = V^a_4 \) (outside option of 4 in \(\{4 \rightarrow 2 \rightarrow 3\}\)). So \(V^{5,14,32}_1 = \hat{V}^{5,14,32}_1 + V^a_3 + V^{5,14,32}_4 - V^a_5 + V^{5,14,32}_4 - \frac{1}{2}\delta\sigma, V^{5,14,32}_2 = \hat{V}^{5,14,32}_2 + \frac{1}{2}\delta\sigma, V^{5,14,32}_3 = \hat{V}^{5,14,32}_3, V^{5,14,32}_4 = V^a_4, V^{5,14,32}_5 = V^a_5. \)

Compared to the previous case, now firm 3 is at the end of the chain. In most of the regions, 3 withdraws when \(\hat{V}^{5,14,32}_3 - V^a_3 < 0.\) Given this, the outside option of 2 is to form a pair with 3, i.e. \(V^{outside}_2 = 1 + \bar{z} + \delta\sigma - \Pr\left(\frac{\bar{a}_1 + \bar{a}_3}{2} < 1\right) - \Phi(-z_2) - \hat{\Phi}(-z_2).\) So we then check whether 2 deviates by comparing \(\bar{V}^{5,14,23}_2\) and \(V^{outside}_2.\) If 2 does not withdraw, \(L^e = \{5 - 1 - 4 - 2, 3\}\). When 2 withdraws, \(\{5 - 1 - 4\}\) is not stable, and the equilibrium network becomes \(L^e = \{5 - 1, 4 - 2 - 3\}.\) Q.E.D.

**B.5 Proof of Proposition 4**

I prove this proposition in a four-firm network setting. The inefficiency occurs in the region \((\bar{z} \in [1, \bar{z}], \delta > \delta_1(\bar{z})),\) where \(L^* = \{4, 1 - 2 - 3\}\) and \(L^e = \{4 - 1 - 2 - 3\}.\) The value loss equals the difference of the total firm values at \(L^*\) compared to \(L^e\),

\[
V^{loss} = \sum_{i=1}^{3} V^{1-2-3}_i + V^a_4 - \sum_{i=1}^{4} V^{4-1-2-3}_i
= 4\bar{z} - 3\Pr\left(\frac{\bar{a}_1 + \bar{a}_2 + \bar{a}_3}{3} < 1\right) - \Phi(-z_4) - \left(4\bar{z} - 4\Pr\left(\bar{h}_4 < 1\right)\right)
= 4 \times \Phi\left[\hat{2}(\bar{z})\right] - 3\Phi\left[\sqrt{3}(\bar{z} - \frac{1}{2}\delta)\right] - \Phi\left[-\bar{z} + \frac{3}{2}\delta\right].
\]

The value loss has the following properties. First, from Proposition 2, \(V^{loss} > 0,\) \(\bar{z} \in\)
Figure A.II. Percentage Value Loss. This figure plots the percentage value loss against $\bar{z}$ and $\delta$ for the equilibrium four-firm chain network.

$$\frac{\partial V^{loss}}{\partial \delta} = \frac{3\sqrt{3}}{2} \Phi' \left[ \sqrt{3}(\bar{z} - \frac{1}{2}\delta) \right] - \frac{3}{2} \Phi' \left[ -\bar{z} + \frac{3}{2}\delta \right]$$

$$= \frac{3\sqrt{3}}{2} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\bar{z} - \frac{1}{2}\delta)^2} - \frac{3}{2} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\bar{z} + \frac{3}{2}\delta)^2}$$

$$= \frac{3}{2} \frac{1}{\sqrt{2\pi}} \left( \sqrt{3} e^{-\frac{3}{2}(\bar{z} - \frac{1}{2}\delta)^2} - e^{-\frac{1}{2}(\bar{z} + \frac{3}{2}\delta)^2} \right) > 0, \quad \bar{z} \in [0, \bar{z}_1], \delta > \delta_1(\bar{z}).$$

So $V^{loss}$ increases with $\delta$. Third, the value loss also decreases with average financial distress $\bar{z}$,

$$\frac{\partial V^{loss}}{\partial \bar{z}} = -8\Phi' [2(\bar{z})] + 3\sqrt{3}\Phi' \left[ \sqrt{3}(\bar{z} - \frac{1}{2}\delta) \right] + \Phi' \left[ -\bar{z} + \frac{3}{2}\delta \right] < 0.$$ 

Finally, the cross-derivative of value loss with respect $\bar{z}$ and $\delta$ is negative

$$\frac{\partial^2 V^{loss}}{\partial \delta \partial \bar{z}} = \frac{3}{2} \Phi'' \left[ -\bar{z} + \frac{3}{2}\delta \right] - \frac{9}{2} \Phi'' \left[ \sqrt{3}(\bar{z} - \frac{1}{2}\delta) \right] < 0,$$ 

which means that the value loss increases faster with $\delta$ when $\bar{z}$ is lower.

Figure A.II illustrates these patterns for $N = 4$. The left panel plots the percentage value loss against $\bar{z}$ when evaluating $\delta = 1$. For $\delta > 0$, the value loss is positive until $\bar{z}$ is large enough. The right panel plots the value loss as a function of the dispersion parameter $\delta$ for $\bar{z} = 0.2$ and $\bar{z} = 0.3$. It shows that value loss increases with $\delta$, and the slope is steeper when $\bar{z}$ is smaller. The negative cross-derivative implies that the impact of heterogeneity in network inefficiency is more pronounced during episodes of banking distress. Q.E.D.

B.6 Proof of Proposition 5

I show that the acquisition tax $\tau$ aligns the social incentive for acquisition with that of firm 1. $\tau$ equals precisely the negative externality imposed by the acquisition behavior of acquiring firm

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1 to all the other non-distressed banks, \( i = 2, \ldots, N - 1 \). Under the required tax payment \( \tau \), the value of 1 upon acquisition is

\[
V_1(\tau) = \hat{V}_{12}^{12,N} - \tau = \sum_{i=1}^{N} \hat{V}_i^{12,N} - \sum_{i=2}^{N-1} \hat{V}_i^{12,N-1} - V_a^N.
\]

Bank 1 chooses to acquire if and only if \( V_1(\tau) \) is larger than the value of not acquiring, i.e.,

\[
V_1(\tau) \geq V_1^{12,N-1}.
\]

Plugging in \( \tau \), this condition is equivalent to

\[
\sum_{i=1}^{N} \hat{V}_i^{12,N} - \sum_{i=2}^{N-1} \hat{V}_i^{12,N-1} - V_a^N \geq \sum_{i=2}^{N-1} \hat{V}_i^{12,N-1} + V_a^N \iff N \text{ is linked into the network.}
\]

Plugging in the following values

\[
\sum_{i=2}^{N-1} \hat{V}_i^{12,N-1} = (N - 2)(1 + \bar{z}) - (N - 2) \Phi \left[ \sqrt{N - 1}(\bar{z} - \frac{1}{2} \delta) \right] c,
\]

\[
\sum_{i=2}^{N-1} \hat{V}_i^{12,N} = N(1 + \bar{z}) - N \Phi \left[ \sqrt{N}(-\bar{z}) \right] c - \left( 1 + \bar{z} + \frac{N - 1}{2} \delta\sigma - \Phi \left[ -\bar{z} - \frac{N - 1}{2} \delta\sigma \right] \right) c;
\]

\[
V_a^N = 1 + \bar{z} + \frac{1 - N}{2} \delta\sigma - \Phi \left[ -\bar{z} - \frac{1 - N}{2} \delta \right] c,
\]

we get that \( \tau \) is a function of \( \{N, \bar{z}, \delta\} \).

\[
\tau_1^A = \left( N \Phi \left[ \sqrt{N}(-\bar{z}) \right] - (N - 2) \Phi \left[ \sqrt{N - 1}(\bar{z} - \frac{1}{2} \delta) \right] - \Phi(-z_1) - p_N \right) c.
\]

Whenever the most distressed firm should be optimally isolated, we have

\[
\sum_{i=1}^{N-1} \hat{V}_i^{12,N-1} + V_a^N - \sum_{i=1}^{N} \hat{V}_i^{12,N} > 0. \tag{46}
\]

Combining the optimal condition (46) and the acquisition incentive of firm 1, i.e. \( \hat{V}_1^{12,N} - \hat{V}_1^{12,N-1} > 0 \).

\[
\tau = \sum_{i=2}^{N-1} \hat{V}_i^{12,N-1} + V_a^N - \sum_{i=2}^{N-1} \hat{V}_i^{12,N}
\]

\[
= \sum_{i=1}^{N-1} \hat{V}_i^{12,N-1} + V_a^N - \sum_{i=1}^{N} \hat{V}_i^{12,N} + (V_1^{12,N} - V_1^{12,N-1}) > 0.
\]

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Further, $\tau$ increases with dispersion $\delta$, decreases with mean $\bar{z}$. To see this, we take the derivatives of $\bar{z}$, $\delta$, and the cross-derivative of $\bar{z}$ and $\delta$.

\[
\frac{\partial \tau_i^A}{\partial \bar{z}} = \sqrt{N - 1}(N - 2)\Phi' \left[ \sqrt{N - 1}(\bar{z} - \frac{1}{2}\delta) \right] c + \Phi' \left[ \bar{z} - \frac{N - 1}{2}\delta \right] c - N\sqrt{N}\Phi' \left[ \sqrt{N}(-\bar{z}) \right] c < 0,
\]

\[
\frac{\partial \tau_i^A}{\partial \delta} = \frac{1}{2}\sqrt{N - 1}(N - 2)\Phi' \left[ \sqrt{N - 1}(\bar{z} - \frac{1}{2}\delta) \right] c + \frac{N - 1}{2} \left( \Phi' \left[ \bar{z} - \frac{N - 1}{2}\delta \right] - \Phi' \left[ \bar{z} - \frac{1}{2}\delta \right] \right) c > 0.
\]

And \( \frac{\partial^2 \tau_i^A}{\partial \bar{z} \partial \delta} < 0\). Q.E.D.

### B.7 Proof of Proposition 8

In this proof, I first show that $C_{isoN}$ decreases monotonically with $\delta$, hence it achieves the maximum at $C_{isoN}(\delta = 0)$. Then I show that $C_{isoN}^*$ is a concave function: $C_{isoN}^* > C_{isoN}(\delta = 0)$ at the minimum value for cost $c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}$, $C_{isoN}^*$ crosses the linear function $C_{isoN}(\delta = 0)$ at $\bar{c}$.

Accordingly, $C_{isoN}^*$ is greater than in the region $c \in \left[ \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}, \bar{c} \right]$.

**Step 1:** $\frac{\partial C_{isoN}}{\partial \delta} < 0$, so $C_{isoN}$ decreases with $\delta$. Take the derivative of $C_{isoN}$ with respect to $\delta$,

\[
\frac{\partial C_{isoN}}{\partial \delta} = \frac{(N - 1)c}{2} \left( \Phi' \left[ (k - 1)\bar{z} + \frac{N - 1}{2}\delta \right] - \sqrt{N - 1}\Phi' \left[ \sqrt{N - 1} \left( \bar{z} - \frac{1}{2}\delta \right) \right] \right). \tag{47}
\]

Notice that $(k - 1)\bar{z} + \frac{N - 1}{2}\delta > 0$, $\sqrt{N - 1} \left( \bar{z} - \frac{1}{2}\delta \right) < 0$, and we can also show that $(k - 1)\bar{z} + \frac{N - 1}{2}\delta > -\sqrt{N - 1} \left( \bar{z} - \frac{1}{2}\delta \right)$.$^{48}$ Accordingly, $\Phi'(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2}$ implies that

\[
\Phi' \left[ (k - 1)\bar{z} + \frac{N - 1}{2}\delta \right] < \Phi' \left[ \sqrt{N - 1} \left( \bar{z} - \frac{1}{2}\delta \right) \right] < \sqrt{N - 1}\Phi' \left[ \sqrt{N - 1} \left( \bar{z} - \frac{1}{2}\delta \right) \right]
\]

Plugging into Equation (47), we have $\frac{\partial C_{isoN}}{\partial \delta} < 0$. Evaluate $C_{isoN}$ at $\delta = 0$, we obtain a linear function of $c$,

\[
C_{isoN}(\delta = 0) = (N - 1)\Phi \left[ -\sqrt{N - 1}\bar{z} \right] c + \Phi \left[ (k - 1)\bar{z} \right] c.
\]

**Step 2:** $C_{isoN}^*$ is a concave function.

---

$^{48}$We take the difference of the squares, $(\bar{z} + \frac{N - 1}{2}\delta)^2 - (\bar{z} - \frac{1}{2}\delta)^2 = (k - 1)^2 \bar{z}^2 + (N - 1)^2\frac{\delta^2}{4} + (N - 2)(N - 1)\bar{z}\delta + (N - 1)\bar{z}\delta > (N - 1)\bar{z}\delta$. The difference between the two is $\frac{(N - 1)^2\delta^2}{4} + (N - 2)(N - 1)\bar{z}\delta = (N - 1)\bar{z}\delta$. Therefore, $\frac{\partial^2 C_{isoN}}{\partial \delta^2} > 0$.
Denote $J = \sqrt{-2 \log \left( \frac{\sqrt{2\pi}\sigma}{\sqrt{Nc}} \right)} > 0$, then $\frac{\partial J}{\partial c} = \frac{1}{Jc}$, and from (29), $\Phi'(-J) = \frac{\sigma}{\sqrt{Nc}}$. Now sub the expression of $J$ into Equation (30), we have

$$C^*_\text{GB} = N\Phi [-J] c + (k - N)\bar{z}\sigma + \sqrt{N}\sigma J.$$ 

Take the first derivative of $c$,

$$\frac{\partial C^*_\text{GB}}{\partial c} = N\Phi [-J] + \frac{\sqrt{N}\sigma}{Jc} - \frac{N\Phi' [-J]}{J} = N\Phi [-J] = N\Phi \left[ -\sqrt{-2 \log \left( \frac{\sqrt{2\pi}\sigma}{\sqrt{Nc}} \right)} \right].$$ 

Therefore, $\frac{\partial C^*_\text{GB}}{\partial c} > 0$. Since $\frac{\partial J}{\partial c} = \frac{1}{Jc} > 0$, $\frac{\partial C^*_\text{GB}}{\partial c^2}$ decreases with $c$, i.e. $\frac{\partial^2 C^*_\text{GB}}{\partial c^2} < 0$.

**Step 3: Establish** $C^*_\text{GB} \left( c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} \right) > C_{\text{isoN}} \left( \delta = 0, c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} \right)$.

Plugging in $c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}$, 

$$C^*_\text{GB} \left( c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} \right) = \frac{\sqrt{2\pi}\sigma}{2} \sqrt{N} + (k - N)\bar{z}\sigma,$$

$$C_{\text{isoN}} \left( \delta = 0, c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} \right) = (N - 1)\Phi \left[ -\sqrt{N - 1}\bar{z} \right] \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} + \Phi \left[ (k - 1)\bar{z} \right] \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}.$$

Since $k > N$, $\Phi < 1$,

$$C_{\text{isoN}} \left( \delta = 0, c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} \right) < N\frac{\sqrt{2\pi}\sigma}{\sqrt{N}} = \frac{\sqrt{2\pi}\sigma}{2} \sqrt{N} < C^*_\text{GB} \left( c = \frac{\sqrt{2\pi}\sigma}{\sqrt{N}} \right).$$

**Step 4: Solve for cross point $\bar{c}$**.

Equating $C^*_\text{GB} = C_{\text{isoN}} (\delta = 0)$ and solve for $c$ gives $\bar{c}$.

Finally, with the results from Steps 1 - 4, we establish that $C^*_\text{GB} > C_{\text{isoN}}$, $\forall c \in \left[ \frac{\sqrt{2\pi}\sigma}{\sqrt{N}}, \bar{c} \right]$.

Q.E.D.

**B.8 Proof of Corollary 1**

In this proof, I first analyze conditions for the acquisition link to be ex post optimal. Then I examine whether the acquisition link forms at equilibrium, and then move to conditions for the positive subsidy to be optimal. Finally, I conclude that subsidized acquisition is cheaper than government bailout.

**Step 1: condition for the acquisition link to be ex post optimal.** Without acquisition link, total liquidation costs of group one and group two are respectively

$$C_{g1} = N\Phi \left[ \frac{-N + k - \bar{z}}{\sqrt{N}} \right] c, \quad C_{g2} = N\Phi \left[ \frac{-N - \hat{k} - \bar{z}}{\sqrt{N}} \right] c. \quad (48)$$

With the acquisition link, the total liquidation costs of the two groups become

$$C_{\text{total}} = \sum_{i=1}^{2N} \Pr (\bar{h}_i < 1) c = 2N\Phi \left[ \frac{-2N - \hat{k} + k - \bar{z} \sqrt{2N}}{\sqrt{2N}} \right] c. \quad (49)$$
The acquisition link generates positive surplus if and only if \( C_{g1} + C_{g2} > C_{\text{total}} \). Plugging in (48) and (49) and applying Lemma 3, we get

\[
N \Phi \left[ \frac{-N + k}{\sqrt{N}} \right] c + N \Phi \left[ \frac{-N - \hat{k}}{\sqrt{N}} \right] c > 2N \Phi \left[ \frac{-2N - \hat{k} + k}{\sqrt{2N}} \right] c \iff \hat{k} > k - 2N.
\]

**Step 2: condition for the acquisition link to be formed ex post at equilibrium.** I show that as long as this acquisition is socially optimal, \( C_{g1} + C_{g2} > C_{\text{total}} \), the acquisition link will form ex post at equilibrium. Since prices are already set between other banks, with only bilateral prices \( \left( p_{N+1}^N, p_{N+1}^N \right) \) to be contracted. Hence whether the acquisition link can form at equilibrium is equivalent to whether the bilateral surplus between \( N \) and \( N + 1 \) is positive. The value of firm \( N \) without the ex post acquisition link is\(^{49}\)

\[
\bar{V}_N = 1 + \left( 1 - \frac{N - 1}{2} \delta - \frac{k}{N} \right) \bar{z} \sigma - \Phi \left[ \frac{k - N}{\sqrt{N}} \bar{z} \right] c - \bar{V}_{N+1} \sigma - \Phi \left[ \frac{-N - 1}{\sqrt{N}} \bar{z} \right] c + \Phi \left( -\frac{1}{\sqrt{N}} \bar{z} \right) c.
\]

Notice that when \( k = 0 \), \( \bar{V}_N = V^a_N \), which matches the outside option of firm \( N \). The value of firm \( N + 1 \) without the ex post acquisition link is

\[
\bar{V}_{N+1} = 1 + \frac{\hat{k} + N}{N} \bar{z} \sigma - \Phi \left[ \frac{-N - \hat{k}}{\sqrt{N}} \bar{z} \right] c.
\]

The bilateral surplus is

\[
\Phi \left[ \frac{-N - \hat{k}}{\sqrt{N}} \bar{z} \right] c + \Phi \left[ \frac{k - N}{\sqrt{N}} \bar{z} \right] c > 2 \Phi \left[ \frac{-2N - \hat{k} + k}{\sqrt{2N}} \bar{z} \right] c \iff \hat{k} > k - 2N
\]

which recovers precisely the condition for positive total acquisition surplus. This shows that if and only if \( \hat{k} > k - 2N \), the acquisition link is efficient and forms in equilibrium after \( \theta \) realizes.

**Step 3: the positive acquisition subsidy is optimal if the liquidation cost is large enough.** When \( \hat{k} \leq k - 2N \), I next show that the positive acquisition subsidy is optimal if the liquidation cost is large enough. Let the positive government subsidy be \( B_A \sigma \) given to the acquire \( N + 1 \). The total cost with subsidized acquisition becomes

\[
C_{\text{sub}A} = \sum_{i=1}^{2N} \Pr \left( \hat{h} < 1 \right) c + B_A \sigma = 2N \Phi \left[ \frac{k - \hat{k} - 2N}{\sqrt{2N}} \bar{z} - B_A \right] c + B_A \sigma.
\]

\( B_A \) satisfies the first order condition

\[
\Phi' \left[ \frac{k - \hat{k} - 2N}{\sqrt{2N}} \bar{z} - B_A \right] = \frac{\sigma}{\sqrt{2Nc}}.
\]

Solving for \( B_A \) gives (31), and we require that \( c > \frac{\sqrt{\pi} \sigma}{\sqrt{N}} \) and \( \hat{k} \leq k - 2N \).

**Step 4: subsidized acquisition is preferred to government bailout.** I show that the subsidized acquisition is less costly thus preferred to government bailout. Based on Proposition

\[\text{Step 4: subsidized acquisition is preferred to government bailout.}\]
of group one and group two are respectively the liquid \( N_1 \) and \( N_2 \), subsidized acquisition is the only feasible option. For \( c > \frac{\sqrt{\pi} \sigma}{\sqrt{N}} \), costs with government bailout for the two groups are

\[
C_{GB}^* = N\Phi \left[ -2\sqrt{2} \log \left( \frac{\sqrt{2\pi} \sigma}{\sqrt{N}} \right) \right] + (k - N)\hat{z} + \sqrt{N}\sigma \left[ -2\sqrt{2} \log \left( \frac{\sqrt{2\pi} \sigma}{\sqrt{N}} \right) \right] + N\Phi \left[ -N - \hat{k} \right] \hat{z}.
\]

Costs with subsidized acquisition is

\[
C_{subA}^* = \left( k - \hat{k} - 2N \right) \hat{z} + \sqrt{2}N \left[ -2\sqrt{2} \log \left( \frac{\sqrt{2\pi} \sigma}{\sqrt{N}} \right) \right] + 2N\Phi \left[ -2\sqrt{2} \log \left( \frac{\sqrt{2\pi} \sigma}{\sqrt{N}} \right) \right] c.
\]

Denote \( J = \sqrt{2}N \hat{z} > 0 \), \( H = \frac{\hat{k} + N}{\sqrt{N}} \hat{z} > 0 \), then

\[
C_{GB}^* = \sqrt{N}\sigma J + N\Phi [-J] c + \sqrt{N}\sigma H + N\Phi [-H] c + \left( k - \hat{k} - 2N \right) \hat{z}.
\]

From (29), \( \Phi'(-J) = \frac{\sigma}{\sqrt{N}} \). Hence, function \( f(x) = \sqrt{N}\sigma x + N\Phi [-x] c \), satisfies \( f'(J) = 0 \), \( f''(x) > 0, \forall x > 0 \). This implies \( C_{GB}^* > 2\sqrt{N\sigma J} + 2N\Phi [-J] c + \left( k - \hat{k} - 2N \right) \hat{z} > \sqrt{2N\sigma J} + 2N\Phi [-J] c + \left( k - \hat{k} - 2N \right) \hat{z} \).

In a similar approach, denote \( G = \sqrt{2}N \hat{z} > 0 \), then \( C_{subA}^* = 2\sqrt{2NG\sigma} + 2N\Phi [-G] c + \left( k - \hat{k} - 2N \right) \hat{z} \), and from (50), \( \Phi' [-G] = \frac{\sigma}{\sqrt{2NG\sigma}} \). Function \( f(x) = \sqrt{2N\sigma x} + 2N\Phi [-x] c, \quad x > 0 \), achieves global \( x > 0 \) minimum at \( x = G \). This implies \( C_{GB}^* > C_{subA}^* \) Q.E.D.

**B.9 Proof of Proposition 9**

I first show that condition (32) implies positive social surplus from the acquisition link between the liquid \( N_1 + 1 \) and the distressed firm \( N_1 \). Without acquisition link, total liquidation costs of group one and group two are respectively

\[
C_{g1} = N_1\Phi \left[ \sqrt{N_1} \left( \frac{k}{N_1} - 1 \right) \hat{z} \right] c, \quad C_{g2} = N_2\Phi \left[ \sqrt{N_2} \left( -1 - \frac{k}{N_2} \right) \hat{z} \right] c
\]

With the acquisition link, the total liquidation costs of the two groups become

\[
C_{total} = \sum_{i=1}^{N_1+N_2} \Pr \left( h_i < 1 \right) c = (N_1 + N_2)\Phi \left[ \sqrt{N_1} + N_2 \left( \frac{k - \hat{k}}{N_1 + N_2} - 1 \right) \hat{z} \right] c.
\]

The acquisition link generates positive surplus if and only if \( C_{g1} + C_{g2} > C_{total} \), i.e.

\[
\frac{N_1}{N_1 + N_2} \Phi \left[ \frac{k - N_1}{\sqrt{N_1}} \hat{z} \right] + \frac{N_2}{N_1 + N_2} \Phi \left[ \frac{-\hat{k} - N_2}{\sqrt{N_2}} \hat{z} \right] > \Phi \left[ \frac{k - \hat{k} - (N_1 + N_2)}{\sqrt{N_1 + N_2}} \hat{z} \right].
\]

Under (32), \( \hat{k} > \max \left( \frac{N_1 + N_2 - \sqrt{N_1}}{N_1 + N_2 - \sqrt{N_2}} (k - N_1) - N_2, \quad k - N_1 - N_2 \right) \). It follows

\[
\left( N_2 + \hat{k} \right) \left( \sqrt{N_1} + N_2 - \sqrt{N_2} \right) (k - N_1) > \left( k - N_1 \right) \left( \sqrt{N_1} + N_2 - \sqrt{N_1} \right) \iff \left( -N_1 \sqrt{N_1} + N_2 k \right) > \sqrt{N_1 + N_2} k - \sqrt{N_1 + N_2} \hat{k} - \sqrt{N_1 + N_2} (N_1 + N_2) \iff \frac{N_1 \sqrt{N_1} (1 - \frac{\hat{k}}{N_1}) \hat{z} + N_2 \sqrt{N_2} (1 - \frac{k}{N_2}) \hat{z}}{N_1 + N_2} > \frac{k - \hat{k} - (N_1 + N_2)}{\sqrt{N_1 + N_2}} \hat{z}.
\]

(52)
Given $\Phi(.)$ is convex when $\frac{k-\hat{k}-(N_1+N_2)}{\sqrt{N_1+N_2}} \bar{z} < 0$, by definition

$$\frac{N_1}{N_1+N_2} \Phi \left[ \frac{k-N_1}{\sqrt{N_1}} \bar{z} \right] + \frac{N_2}{N_1+N_2} \Phi \left[ \frac{-\hat{k} - N_2}{\sqrt{N_2}} \bar{z} \right] \geq \Phi \left[ \frac{N_1\sqrt{N_1} \left( -1 + \frac{k}{N_1} \right) \bar{z}}{N_1 + N_2} \right] + \frac{N_2\sqrt{N_2} \left( -1 - \frac{k}{N_2} \right) \bar{z}}{N_1 + N_2}.$$  

Combining with Equation (52), we establish (51).

Next I show that under (32), the bilateral acquisition surplus is positive when $N_2 \geq N_1$. Since prices are already set between other banks, there are only bilateral prices $\left( p_{N_1}^{N_1+1} , p_{N_1+1}^{N_1} \right)$ to be contracted. Hence whether the acquisition link can form at equilibrium is equivalent to whether the bilateral surplus between $N_1$ and $N_1 + 1$ is positive. The value of firm $N_1$ without acquisition is

$$\hat{V}_{N_1} = 1 + \left( 1 - \frac{N_1 - 1}{2} \delta - \frac{k}{N_1} \right) \bar{\sigma} - \Phi \left[ \frac{k - N_1}{\sqrt{N_1}} \right] c - \Phi \left( \frac{N_1 - 1}{2} \delta - \bar{z} \right) c + \Phi \left( -\sqrt{N_1} \bar{z} \right) c.$$

Notice that when $k = 0$, $\hat{V}_{N_1} = V_{N_1}^0$, which matches the outside option of firm $N_1$. The value of firm $N_1 + 1$ without the acquisition link is

$$\hat{V}_{N_1+1} = 1 + \left( 1 + \frac{\hat{k}}{N_2} \right) \bar{\sigma} - \Phi \left[ \frac{-N_2 - \hat{k}}{\sqrt{N_2}} \right] c.$$

With the acquisition link, the value of firm $N_1$, and firm $N_1 + 1$ are respectively

$$\hat{V}_{N_1}^A = 1 + \left( 1 - \frac{N_1 - 1}{2} \delta - \frac{k - \hat{k}}{N_1 + N_2} \right) \bar{\sigma} - \Phi \left[ \frac{k - \hat{k} - (N_1 + N_2)}{\sqrt{N_1 + N_2}} \bar{z} \right] c - \Phi \left( -\bar{z} + \frac{N_1 - 1}{2} \delta \right) c + \Phi \left( -\sqrt{N_1} \bar{z} \right) c,$$

$$\hat{V}_{N_1+1}^A = 1 + \left( 1 - \frac{k - \hat{k}}{N_1 + N_2} \right) \bar{\sigma} - \Phi \left[ \frac{k - \hat{k} - (N_1 + N_2)}{\sqrt{N_1 + N_2}} \bar{z} \right] c.$$

The bilateral surplus minus the total surplus is

$$\frac{\hat{V}_{N_1}^A + \hat{V}_{N_1+1}^A - \hat{V}_{N_1} - \hat{V}_{N_1+1}}{2} = \frac{C_{g1} + C_{g2} - C_{total}}{N_1 + N_2}$$

$$= \frac{1}{2} \left( \frac{k - \hat{k}}{N_1} - 2 \frac{k - \hat{k}}{N_1 + N_2} \right) \bar{\sigma} + \frac{N_2 - N_1}{2(N_1 + N_2)} \left( \Phi \left[ \frac{k - N_1}{\sqrt{N_1}} \bar{z} \right] - \Phi \left[ \frac{-\hat{k} - N_2}{\sqrt{N_2}} \bar{z} \right] \right) c$$

$$= \frac{(N_2 - N_1)}{2N_1N_2(N_1 + N_2)} \bar{\sigma} + \frac{N_2 - N_1}{2(N_1 + N_2)} \left( \Phi \left[ \frac{k - N_1}{\sqrt{N_1}} \bar{z} \right] - \Phi \left[ \frac{-\hat{k} - N_2}{\sqrt{N_2}} \bar{z} \right] \right) c$$

$$= \frac{N_2 - N_1}{2(N_1 + N_2)} \left[ \frac{N_2k + N_1\hat{k}}{N_1N_2} \bar{\sigma} + \left( \Phi \left[ \frac{k - N_1}{\sqrt{N_1}} \bar{z} \right] - \Phi \left[ \frac{-\hat{k} - N_2}{\sqrt{N_2}} \bar{z} \right] \right) c \right].$$

which is non-negative when $N_2 \geq N_1$. In other words, when $N_2 \geq N_1$, and $C_{g1} + C_{g2} - C_{total} > 0$,

$$\frac{\hat{V}_{N_1}^A + \hat{V}_{N_1+1}^A - \hat{V}_{N_1} - \hat{V}_{N_1+1}}{2} \geq \frac{C_{g1} + C_{g2} - C_{total}}{N_1 + N_2} > 0.$$

If $N_1 > N_2$, the average bilateral surplus is smaller than the average social surplus. Under condition (33), the bilateral surplus is negative. Q.E.D.
C Additional Empirical Results

In this Appendix, I provide additional empirical results to supplement the findings in Section 6.

Table A.I presents supplementary univariate correlations to Table 3. I adopt alternative indicators for economic activity and systemic risk, including the *Recession Probability* from Chauvet and Piger (2008), the subcomponents of the Chicago Fed National Activity Index (CFNAI) on personal consumption and housing (C&H) and employment, unemployment, and hours (EU&H). Finally, following Giglio, Kelly, and Pruitt (2013), I take the systemic risk measures relating to liquidity and credit conditions in the financial market: the *Default Spread* (difference between 3-Month BAA bond yields and the Treasury) and the *Term Spread* (difference between 10-Year and 3-Month Treasury).

**Table A.I. Summary Statistics and Univariate Correlations**

<table>
<thead>
<tr>
<th></th>
<th>Mean</th>
<th>StDev</th>
<th>Sacf</th>
<th>Correlations w/ log Z-score</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mean</td>
<td>Disp</td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>A. Economic activity and systemic risk</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Recession Probability</td>
<td>0.08</td>
<td>0.23</td>
<td>0.83</td>
<td>0.00</td>
</tr>
<tr>
<td>CFNAI: Personal Consumption and Housing</td>
<td>-0.03</td>
<td>0.13</td>
<td>0.93</td>
<td>-0.03</td>
</tr>
<tr>
<td>CFNAI: Employment, Unemployment, and Hours</td>
<td>-0.06</td>
<td>0.31</td>
<td>0.86</td>
<td>-0.03</td>
</tr>
<tr>
<td>Default Spread</td>
<td>4.19</td>
<td>1.54</td>
<td>0.93</td>
<td>-0.14</td>
</tr>
<tr>
<td>Term Spread</td>
<td>1.87</td>
<td>1.11</td>
<td>0.91</td>
<td>-0.25**</td>
</tr>
<tr>
<td><strong>B. Lending and interbank lending</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Financial Business Leverage</td>
<td>29.40</td>
<td>5.59</td>
<td>0.94</td>
<td>-0.16</td>
</tr>
<tr>
<td>Security Broker-Dealers Leverage</td>
<td>41.11</td>
<td>17.94</td>
<td>0.73</td>
<td>0.51***</td>
</tr>
<tr>
<td>Δ% Non-financial Corporate Liability</td>
<td>0.01</td>
<td>0.01</td>
<td>0.46</td>
<td>0.14</td>
</tr>
<tr>
<td>All Comm. Bank Credit over Assets</td>
<td>0.81</td>
<td>0.03</td>
<td>0.94</td>
<td>-0.13</td>
</tr>
<tr>
<td>Small Comm. Interbank Loan over Assets</td>
<td>0.02</td>
<td>0.01</td>
<td>0.87</td>
<td>0.10</td>
</tr>
</tbody>
</table>

Notes: This table supplements to Table 3 and reports the summary statistics for alternative measures of economic activity and systemic risk, lending and interbank lending, as well as their univariate correlation coefficients with the mean and dispersion of financials’ log Z-scores. Group A series are taken from FRED. Group B series are constructed from the Fed’s Z.1 and H.8 release. Data availability on bank holding companies restricts the analysis to 1986-2013. Sacf is the first-order sample autocorrelation coefficient. The last two columns report the correlation coefficients between the cross-sectional mean and dispersion of log Z-score and each series in groups A-B together with the significance levels. *, **, *** denote statistical significance at the 5%, 1%, and 0.1% level.

The alternative indicators for lending and interbank lending include the leverage of both financial business and the security broker-dealers discussed in Adrian, Etula, and Muir (2014), the growth rate of non-financial corporate liability, the credit and loans of all commercial banks over assets, and the interbank loans over assets of small and medium-sized commercial banks. The correlation coefficients show a clear pattern: the aggregate indicators correlate significantly with dispersion, whereas only the leverage of security broker-dealers comoves strongly with the mean.
Table A.II presents supplementary predictive regression results to Table 4. Using the same method as in Table 4, I run predictive regressions to forecast the alternative measures. The estimates strongly echo the findings from the correlation analysis. Both the significance level of the regression coefficients and the differences in $R^2$'s with and without dispersion in the regressors suggest the robustness of the predictive power of dispersion series.

The economic magnitude of the predictive power is also sizable. Take the forecasting of Recession Probability for instance, holding the controls fixed, a one-standard-deviation increase in the Dispersion ($=0.22$) predicts a 0.095 ($= 0.22 \times 0.43$) increase in the Recession Probability in the next quarter, whereas a one-standard-deviation decrease in the Mean ($=0.03$) predicts a 0.046 ($= 0.03 \times 1.54$) raise in the future Recession Probability.

**Table A.II. Predictive Regressions using Distress Dispersion**

<table>
<thead>
<tr>
<th>Forecasting</th>
<th>A. Recession Probability</th>
<th>B. CFNAI: CH</th>
<th>C. Default Spread</th>
<th>D. Term Spread</th>
<th>I. Bk Credit over Assets</th>
<th>K. Smal Bk Interbank L over Assets</th>
</tr>
</thead>
<tbody>
<tr>
<td>Dispersion</td>
<td>0.43** 0.79** 1.07* 1.29*</td>
<td>-0.38*** -0.75*** -1.10*** -1.43***</td>
<td>0.04*** 0.08*** 0.13*** 0.18***</td>
<td>0.01* 0.02* 0.04* 0.06</td>
<td>-0.09*** -0.18*** -0.28*** -0.37***</td>
<td>-0.01 -0.02 -0.04* -0.05**</td>
</tr>
<tr>
<td>Mean</td>
<td>-1.54** -3.28** -4.83** -5.12*</td>
<td>0.27 0.70 1.23 1.76</td>
<td>0.00 -0.02 -0.08 -0.12</td>
<td>-0.02 -0.05 -0.13 -0.22</td>
<td>-0.02 0.07 0.18 0.27</td>
<td>-0.06 0.10 0.11 0.09</td>
</tr>
<tr>
<td>$R^2$</td>
<td>40.98 43.91 45.95 42.37</td>
<td>70.18 72.88 74.29 75.30</td>
<td>83.97 79.92 74.97 69.61</td>
<td>87.19 80.07 72.93 66.32</td>
<td>80.41 82.41 83.95 85.16</td>
<td>48.13 54.09 56.03 56.46</td>
</tr>
<tr>
<td>$R^2$ w/o disp</td>
<td>33.95 37.23 39.95 36.96</td>
<td>53.71 56.23 57.85 59.39</td>
<td>72.23 65.71 58.38 49.95</td>
<td>86.49 78.48 70.32 62.16</td>
<td>64.37 65.53 65.98 65.79</td>
<td>45.00 49.57 49.55 47.33</td>
</tr>
</tbody>
</table>

**Notes:** This table summarizes the ability of distress dispersion to forecast future economic activity, systemic risk, failure rates, distressed acquisition rates, and bank lending behavior. In A-K, quarterly time series are regressed on the cross-sectional dispersion and mean of log Z-score controlling for the term spread, the leverage of financial business and security broker-dealers, and the growth rate of real non-financial corporate liability. Forecasting horizons range from one to four quarters and the data cover the years 1986-2013. The table reports the predictive regression coefficients on the dispersion and mean of log Z-score, the $R^2$, as well as the $R^2$ when the regressions are run without the dispersion series. *, **, *** denote statistical significance (based on Newey-West standard errors) at the 5%, 1%, and 0.1% level.