

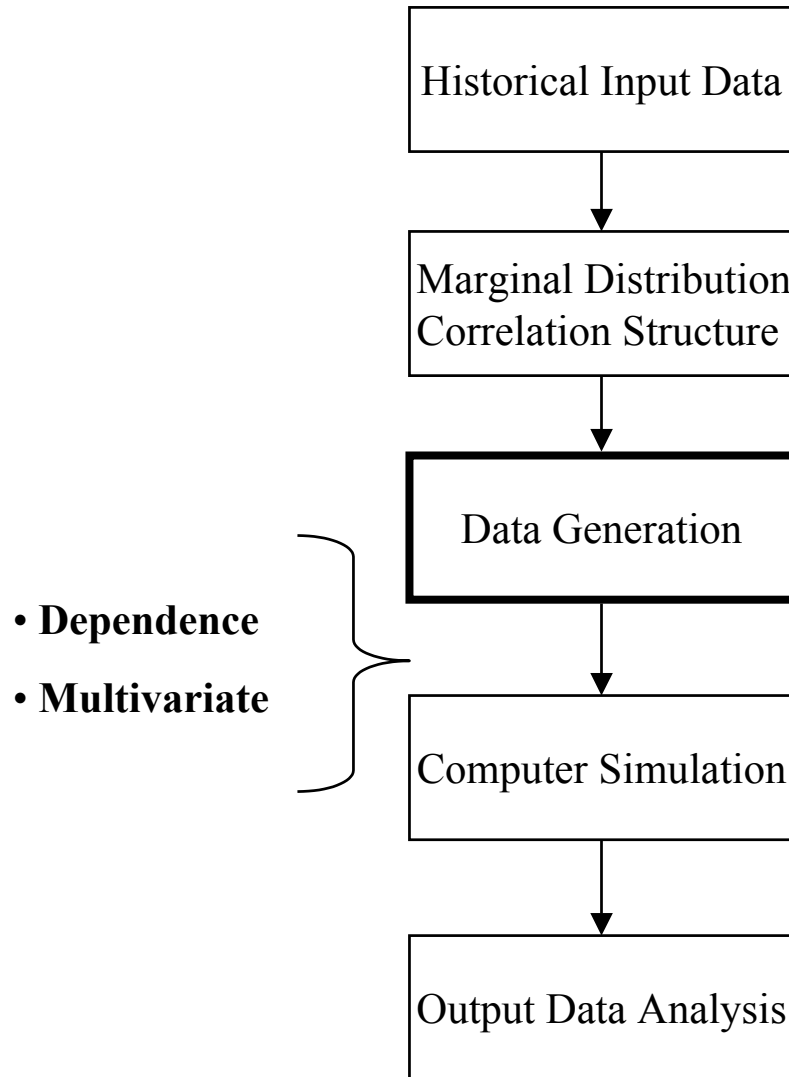
**Stochastic Input Modeling considering Dependencies:
Generating Multivariate Time Series
using a Vector Autoregressive Technique***

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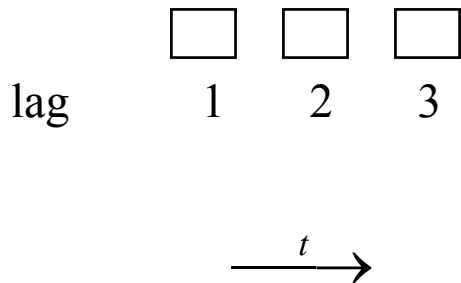
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Stochastic Input Modeling



Problem Setup & Contribution

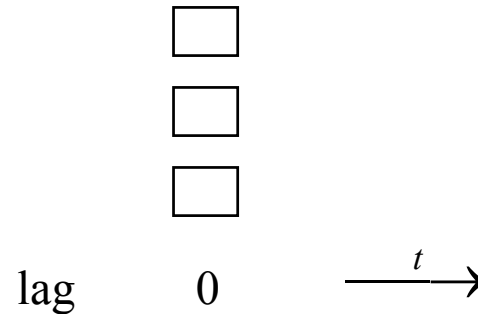
Univariate Time Series



Approaches:

- Exploit properties specific to the marginal distribution (Lewis et al. 1989)
- Construct a series of autocorrelated uniform random variables as the base process and apply the inverse transform method (Melamed 1991, Cario and Nelson 1996)

Random Vectors



Approach:

- Transform multivariate copula into vectors with arbitrary marginal distributions (Cario et al. 2000, Chen 2000, Ghosh and Henderson 2000)

Develop a comprehensive framework that pulls together the theory behind univariate time series and random vectors with dependent components and extend it to the multivariate time series

Modeling Approach

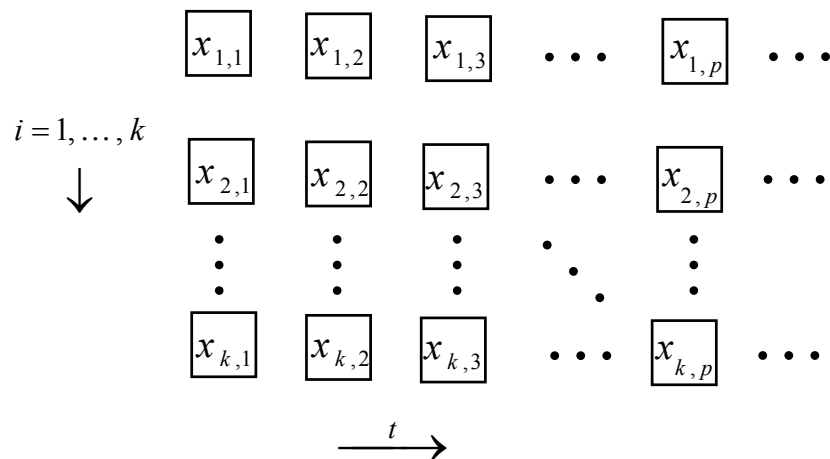
Customize a flexible model incorporating

- a vector auto-regressive process
- Johnson family of marginal distributions

Step 0

- For the input process X_t , obtain Johnson marginal distributions for each of the k series specified through their first four moments and the correlation matrices for lags ranging from 0 to p

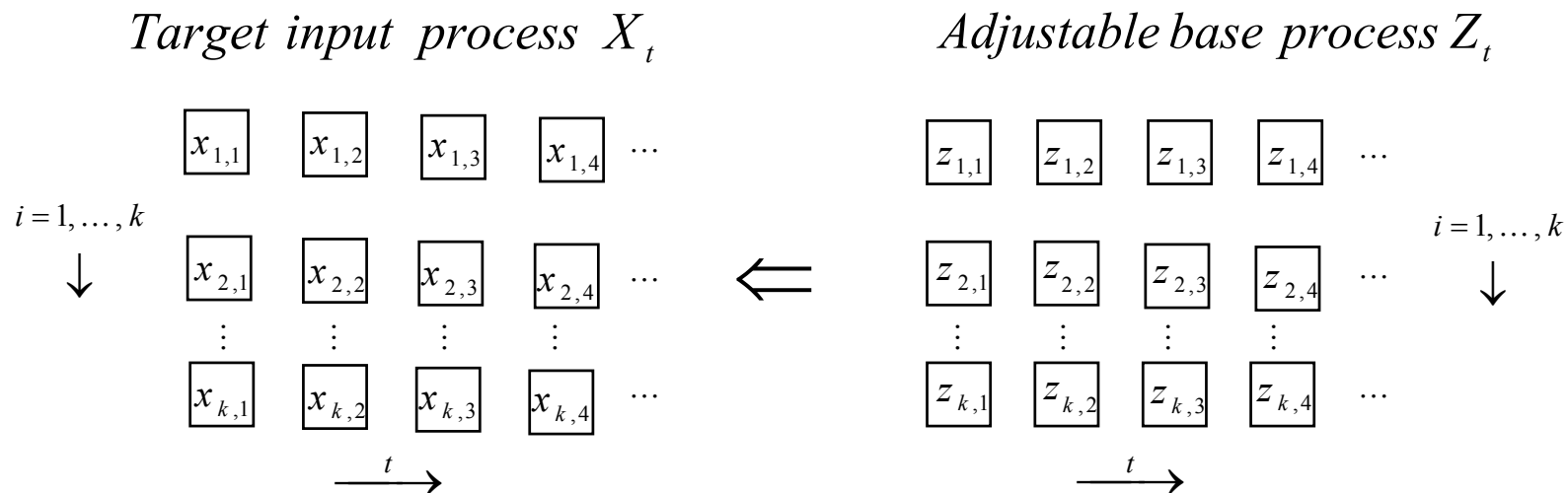
Input process X_t



Modeling Approach (cont'd)

Step 1

- Characterize the underlying vector autoregressive model, Z_t , by solving $pk^2 + k(k-1)/2$ correlation matching problems



- Determine system parameters using the correlation matrices $\Sigma_z^{(0)}, \dots, \Sigma_z^{(p)}$
- Generate standard normal random variates, $z_{i,t}$, $i = 1, \dots, k$, $t = 0, \dots, T$, using the system parameters

Step 2

- Transform $z_{i,t}$ into Johnson variates $x_{i,t}$ by $x_{i,t} = \varepsilon_i + \lambda_i g_i^{-1} \left(\frac{z_{i,t} - \gamma_i}{\delta_i} \right)$

Johnson Family of Distributions

Definition: For a Johnson variable X , $F_X(x) = \Phi \left\{ \gamma + \delta g \left(\frac{x - \varepsilon}{\lambda} \right) \right\}$

$$g(y) = \begin{cases} y & \text{for } S_N \text{ (normal) family,} \\ \log(y) & \text{for } S_L \text{ (lognormal) family,} \\ \sinh^{-1}(y) & \text{for } S_U \text{ (unbounded) family,} \\ \log\left(\frac{y}{1-y}\right) & \text{for } S_B \text{ (bounded) family.} \end{cases}$$

Φ : standard normal cumulative distribution function
 $(\gamma, \delta, \lambda, \varepsilon)$: Johnson parameters

Characteristics

- Match any feasible first four moments
- Provide unique distribution for each combination of skewness & kurtosis
- Generate random variates: $x = \varepsilon + \lambda g^{-1}\left(\frac{z - \gamma}{\delta}\right)$ where $z \approx N(0, 1)$

Vector Autoregressive Model

Definition

$$Z_t = \alpha_1 Z_{t-1} + \alpha_2 Z_{t-2} + \dots + \alpha_p Z_{t-p} + u_t, \quad t = 0, \pm 1, \pm 2, \dots$$

- Z_t is $(k \times 1)$, α_i is $(k \times k)$, u_t is $(k \times 1)$

Assumption

- Z_t is stationarity \Rightarrow stationarity of the input process X_t
- $u_t \approx N_k(0, \Sigma_u) \Rightarrow Z_t$ is a Gaussian process

System Parameter Estimation

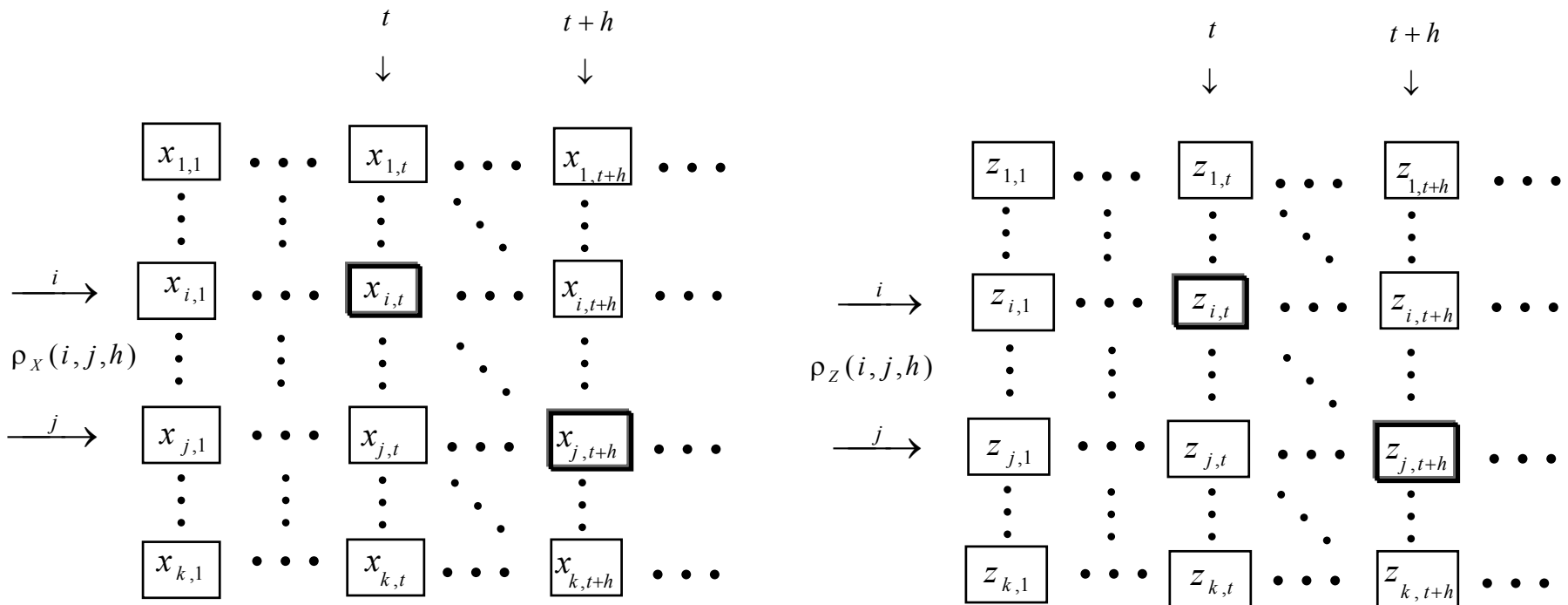
- Determine $\alpha_1, \dots, \alpha_p$ using multivariate Yule-Walker equations
- Solve $\Sigma_u = \Sigma_Z(0) - \alpha_1 \Sigma_Z'(1) - \dots - \alpha_p \Sigma_Z'(p)$

Correlation Matching Problem

Objective: Select the covariance structure for the vector autoregressive process that gives the desired covariance structure for the input process

Input process X_t

Base process Z_t



$$c_{ijh}[\rho_Z(i, j, h)] = \rho_X(i, j, h) \quad (\text{independent matching problems})$$

Correlation Matching Problem

$$\text{Corr} [X_{i,t}, X_{j,t+h}] = \rho_X(i, j, h) = \frac{E[X_{i,t} * X_{j,t+h}] - \mu_i \mu_j}{\sigma_i \sigma_j}$$

$$\begin{aligned} E [X_{i,t} X_{j,t+h}] &= E \left\{ F_{X_i}^{-1} [\Phi(Z_{i,t})] * F_{X_j}^{-1} [\Phi(Z_{j,t+h})] \right\} \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \underbrace{F_{X_i}^{-1} [\Phi(z_i)]}_{x_i = \varepsilon_i + \lambda_i g_i^{-1} \left(\frac{z_i - \gamma_i}{\delta_i} \right)} \underbrace{F_{X_j}^{-1} [\Phi(z_j)]}_{x_j = \varepsilon_j + \lambda_j g_j^{-1} \left(\frac{z_j - \gamma_j}{\delta_j} \right)} \overbrace{\mathfrak{G}(z_i, z_j, \rho_Z(i, j, h))}^{\rho_Z(i, j, h) \text{ appears only here}} dz_i dz_j \end{aligned}$$

Rewriting the correlation function as

$$c_{ijh} [\rho_Z(i, j, h)] = \frac{\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left[\varepsilon_i + \lambda_i f_i^{-1} \left(\frac{z_i - \gamma_i}{\delta_i} \right) \right] \left[\varepsilon_j + \lambda_j f_j^{-1} \left(\frac{z_j - \gamma_j}{\delta_j} \right) \right] \mathfrak{G}(z_i, z_j; \rho_Z(i, j, h)) dz_i dz_j - \mu_i \mu_j}{\sigma_i \sigma_j}$$

$$\Rightarrow \begin{matrix} i, j = 1, \dots, k \\ i \neq j \\ h = 0, \dots, p \end{matrix} c_{ijh} [\rho_Z(i, j, h)] = \rho_X(i, j, h) \Rightarrow pk^2 + k(k-1)/2 \text{ matching problems}$$

Implementation

Objective: Establish properties of the function $c_{ijh} [\rho_Z(i,j,h)]$ to perform an efficient numerical search to find $\rho_Z(i,j,h)$ within a predetermined precision

Useful Results

- $c_{ijh} [0] = 0$
- $\rho_Z(i,j,h) \geq (\leq) 0 \Rightarrow \rho_X(i,j,h) \geq (\leq) 0$
- Extremal values of $\rho_X(i,j,h)$ are attainable
- $c_{ijh} [\rho_Z(i,j,h)]$ is continuous and non-decreasing for $-1 \leq \rho_Z(i,j,h) \leq 1$

Any feasible bivariate correlation for F_{X_i} and F_{X_j} is attainable

Implementation (cont'd)

- The regula falsi (secant) solves each correlation matching problem very accurately and efficiently, particularly for the Johnson system:
 - Define $w[\rho_z] = c[\rho_z] - \rho_x$
 - If $\rho_x > 0 \Rightarrow y_0 = 0$ ($w(0) < 0$), $y_1 = 1$ ($w(1) > 0$)
Else $\rho_x < 0 \Rightarrow y_0 = 0$ ($w(0) < 0$), $y_1 = -1$ ($w(-1) < 0$)
 - Iterative procedure is given by
$$y_{m+1} = y_m - w(y_m) \frac{y_m - y_{m-1}}{w(y_m) - w(y_{m-1})}$$
 - Convergence is certain as $w(x_0)$ and $w(x_1)$ have opposite sign.
- Evaluate $w(y_m)$ through a numerical integration using a doubly-infinite hyper-cube transformation
- Estimate an accurate and reliable error for $w(y_m)$ using null rules

Hyper-cube Transformation

$$\psi(z_i^*) = \tan\left(\frac{\pi z_i^*}{2}\right), \quad -1 < z_i^* < 1, \quad \frac{\partial \psi(z_i^*)}{\partial z_i^*} = \frac{\pi}{2} \left(1 + \tan^2\left(\frac{\pi z_i^*}{2}\right)\right), \quad i=1, 2.$$

⇒

For $|\rho| < 1$:

$$\int_{-1}^1 \int_{-1}^1 \frac{w(\tan(\pi z_1^*/2), \tan(\pi z_2^*/2)) \cdot g(\tan(\pi z_1^*/2), \tan(\pi z_2^*/2), \rho)}{4/\pi^2 (1 + \tan^2(\pi z_1^*/2))^{-1} (1 + \tan^2(\pi z_2^*/2))^{-1}} dz_1 dz_2$$

where $w(z_1, z_2) = e^{-(z_1^2 + z_2^2)}$

For $\rho = 1$ and $\rho = -1$:

$$\left\{ \begin{array}{l} \int_{-1}^1 \int_{-1}^1 \frac{[\varepsilon_1 + \lambda_1 f_1^{-1}((\tan(\pi z_1^*/2) - \gamma_1)/\delta_1)] [\varepsilon_2 + \lambda_2 f_2^{-1}((\tan(\pi z_2^*/2) - \gamma_{21})/\delta_{21})]}{2\sqrt{2} \sqrt{\pi^{-1}} (1 + \tan^2(\pi z_1^*/2))^{-1} e^{\tan^2(\pi z_2^*/2)/2}} dz_1^* dz_2^* \\ \int_{-1}^1 \int_{-1}^1 \frac{[\varepsilon_1 + \lambda_1 f_1^{-1}((\tan(\pi z_1^*/2) \gamma_1)/\delta_1)] [\varepsilon_2 + \lambda_2 f_2^{-1}((-\tan(\pi z_2^*/2) \gamma_{21})/\delta_{21})]}{2\sqrt{2} \sqrt{\pi^{-1}} (1 + \tan^2(\pi z_1^*/2))^{-1} e^{\tan^2(\pi z_2^*/2)/2}} dz_1^* dz_2^* \end{array} \right.$$

Null Rules

Definition
$$N_d(\ell) = \sum_{i=1}^n u_i \ell(x_i), \exists u_i \neq 0, \sum_{i=1}^n u_i = 0.$$

Characteristics

- Maps to zero all polynomials of degree not more than d
- Equally strong with the cubature formula $\sum_{i=1}^n u_i^2 = \sum_{i=1}^n c_i^2$
- Decrease the probability of phase factor effects
- Heuristically distinguish possible types of behavior to extract an error estimate:

$$r^2 = \max \left\{ \frac{N_3(\ell)}{N_1(\ell)}, \frac{N_5(\ell)}{N_3(\ell)}, \frac{N_7(\ell)}{N_5(\ell)} \right\}$$

If $r < 1$, then strong asymptotic behavior

\Rightarrow

$$\Rightarrow \text{error estimate} \cong a r^{q-s} N_s(\ell),$$

Else if $r_{crit} \leq r \leq 1$, then weak asymptotic behavior

$$\Rightarrow \text{error estimate} \cong K r^2 N_s(\ell), \quad a = K r^{s-q+2}_{crit}.$$

Else if $r > 1$, then no asymptotic behavior

$$\Rightarrow \text{error estimate} \cong K N_s(\ell).$$

Example

Step 0

A bivariate random variable with marginals and correlation given as

- Unbounded distribution specified by $\mu = 0, \sigma = 1, \beta_1 = 0.5, \beta_2 = 5$

$$\Rightarrow \gamma_1 = 0, \delta_1 = 1.875, \lambda_1 = 1.875, \varepsilon_1 = 0$$

- Bounded distribution specified by $\mu = 0, \sigma = 1, \beta_1 = 1, \beta_2 = 2.8$

$$\Rightarrow \gamma_2 = 0.478, \delta_2 = 0.375, \lambda_2 = 4.123, \varepsilon_2 = -0.859$$

- Correlation matrices are given at lag 0 and lag 1 ($k=2, p=1$)

$$\sum_x(0) = \begin{pmatrix} 1 & 0.801 \\ 0.801 & 1 \end{pmatrix} \text{ and } \sum_x(1) = \begin{pmatrix} 0.829 & 0.309 \\ 0.516 & 0.713 \end{pmatrix}$$

Example (cont'd)

Step 1

- Characterize the underlying process by solving 5 correlation matching problems

$$\Sigma_Z(0) = \begin{pmatrix} 1 & 0.599 \\ 0.599 & 1 \end{pmatrix} \text{ and } \Sigma_Z(1) = \begin{pmatrix} 0.621 & 0.232 \\ 0.388 & 0.534 \end{pmatrix}$$

- Using the correlation structure, determine the system parameters for

$$Z_t = \alpha_1 Z_{t-1} + u_t, t=0, 1, \dots$$

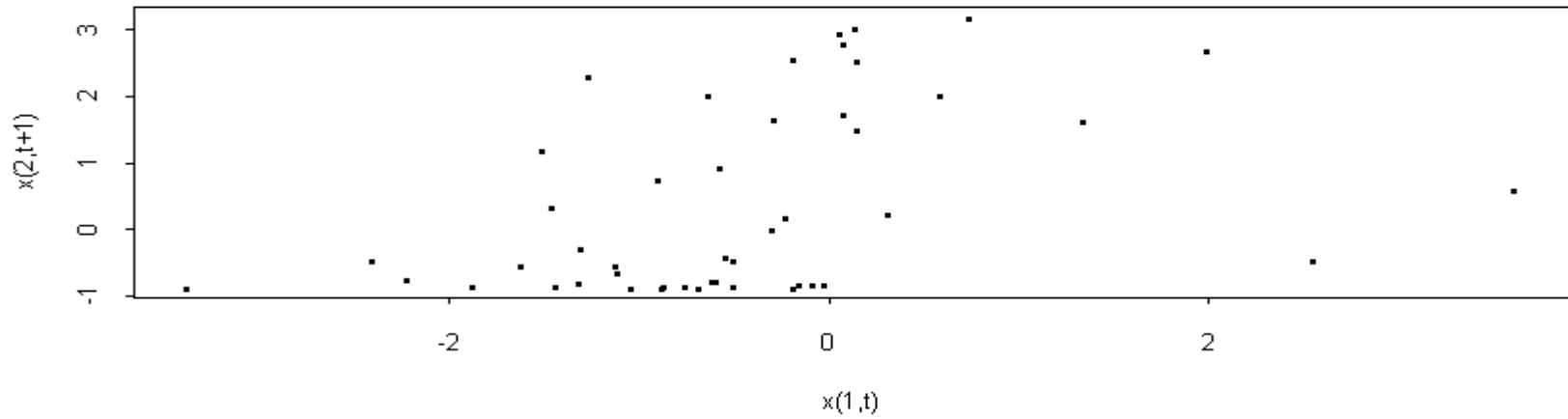
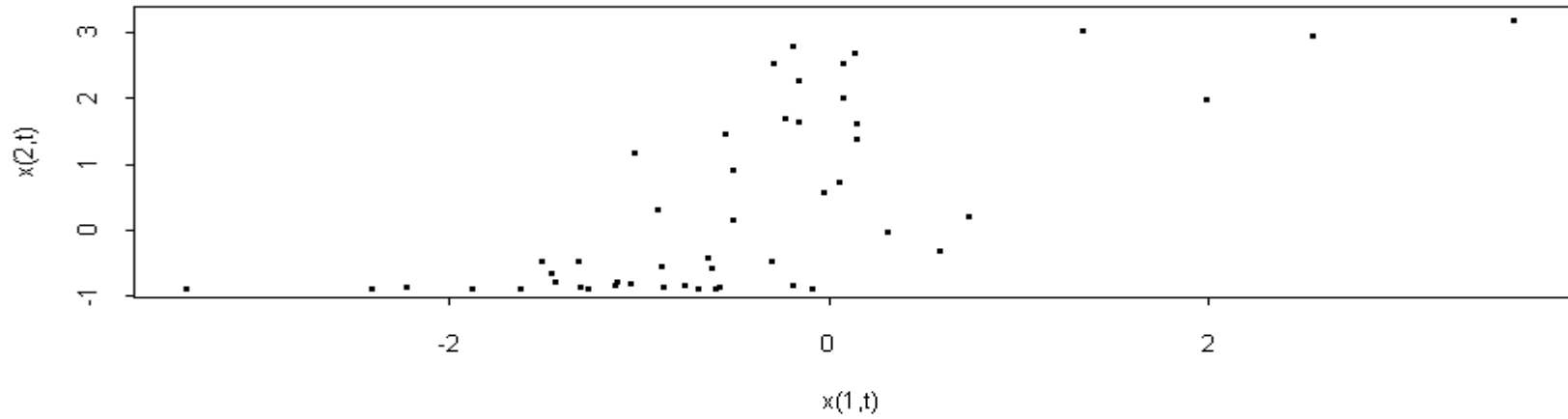
$$\alpha_1 = \begin{pmatrix} 0.752 & -0.218 \\ 0.105 & 0.471 \end{pmatrix} \quad \Sigma_u = \begin{pmatrix} 0.584 & 0.425 \\ 0.425 & 0.708 \end{pmatrix}$$

- Simulate the vector autoregressive model to obtain $z_{i,t}$

Step 2

- Using $z_{i,t}$, generate Johnson variates $x_{i,t} = \varepsilon_i + \lambda_i g_i^{-1} \left(\frac{z_{i,t} - \gamma_i}{\delta_i} \right)$

Example (cont'd)



Summary and Future Research

- Provided a general-purpose tool for modeling and generating dependent and multivariate input processes
- Given historical data
 - determine which Johnson system to use
 - estimate the parameters of the corresponding distribution
 - capture the dependence structure in the multivariate input data